

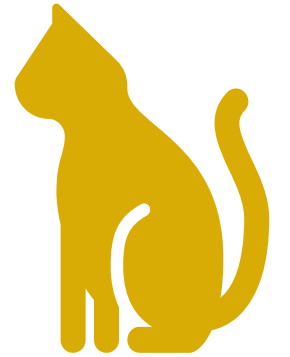


Advances in Approximate Inference

Yingzhen Li, Cheng Zhang

Microsoft Research Cambridge

What is the Number?



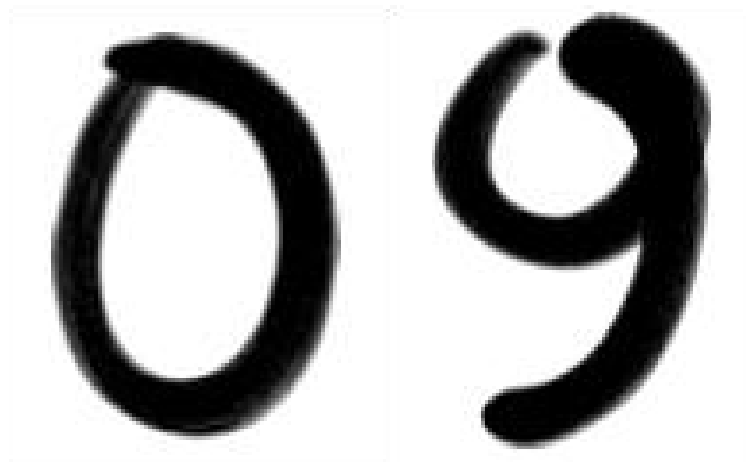
07? 09?

67? 69?

...



What is the Number?



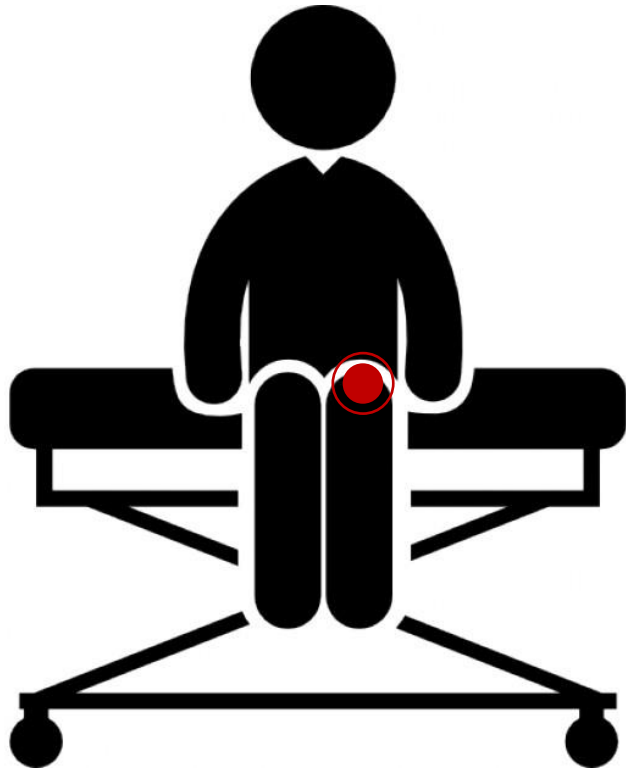
What is the Diagnosis?



Injury?
Osteoarthritis?
Neuropathic pain?
... ..



What is the Diagnosis?



Neuropathic pain
(might have spine injury)



Uncertainty is Important



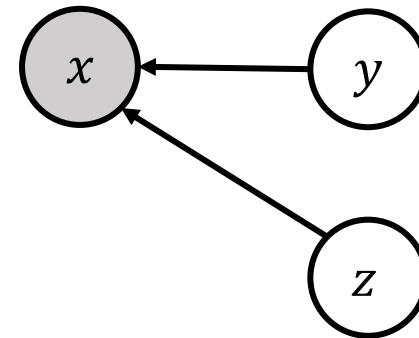


Bayesian ML / Probability Theory



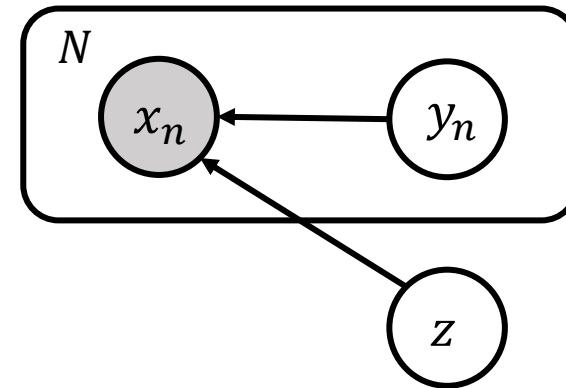
Decision making under uncertainty

Graphical Representation



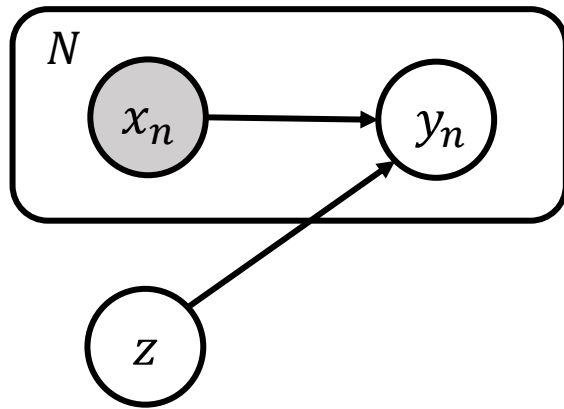
$$p(x, y, z) = p(y)p(z)p(x|y, z)$$

Graphical Representation

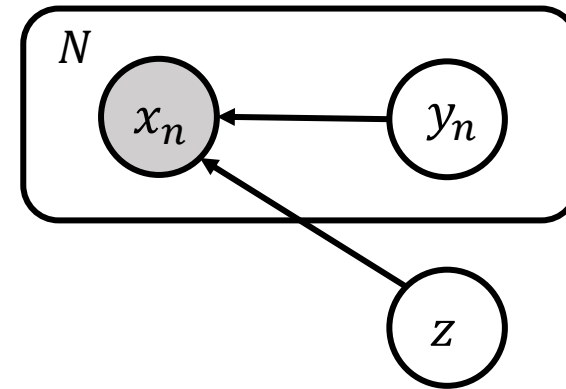


$$p(\mathbf{x}, \mathbf{y}, z) = p(z) \prod_n^N p(y_n) P(y_n | y_n, z)$$

Graphical Representation

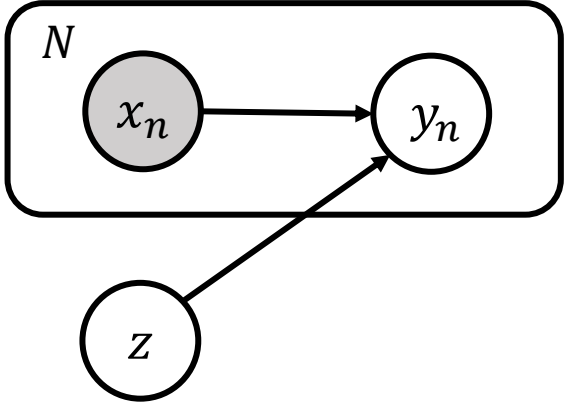


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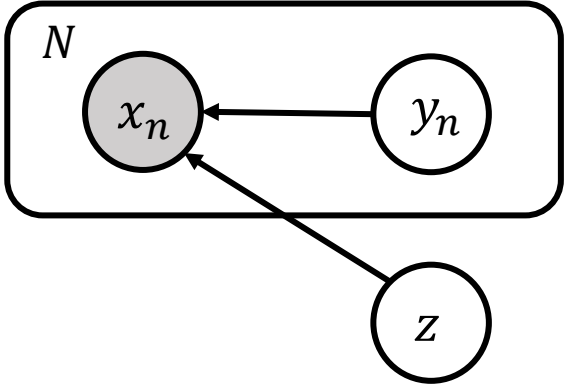


$$p(\mathbf{x}, \mathbf{y}, z) = p(z) \prod_n^N p(y_n) P(x_n | y_n, z)$$

Discriminative Model vs Generative Model



$$p(\mathbf{x}, \mathbf{y}, z) = p(z) \prod_n p(x_n) P(y_n | x_n, z)$$

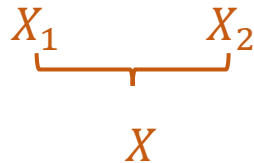


$$p(\mathbf{x}, \mathbf{y}, z) = p(z) \prod_n p(y_n) P(x_n | y_n, z)$$

Discriminative Model Example

- Bayesian Logistic Regression

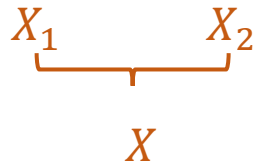
Name	A-level math score	# parents in STEM	Study STEM?
Alice	89	0	0 (No)
Bob	95	1	1 (Yes)
Ty	82	1	0 (No)
Emma	98	2	1 (Yes)
Anna	92	0	0 (No)
Mo	88	1	0 (No)
Li	95	0	1 (Yes)



Discriminative Model Example:

- Bayesian Logistic Regression

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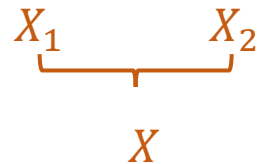


$$p(y = 1) = \frac{1}{1 + e^{w_0 + w_1 x_1 + w_2 x_2}}$$

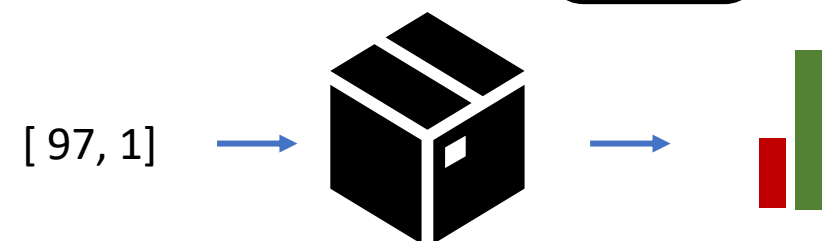
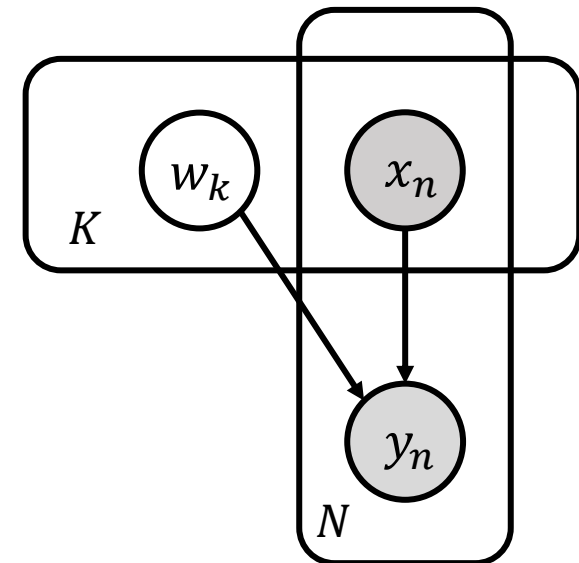
Discriminative Model Example

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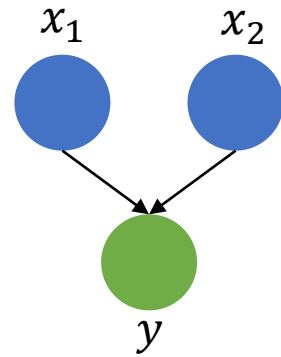


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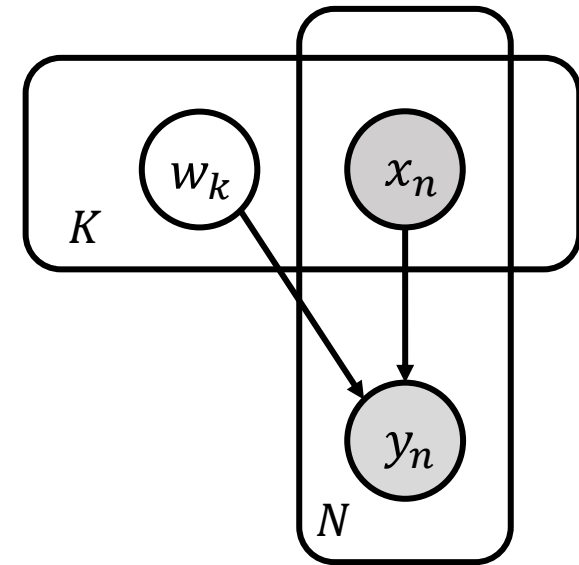


Discriminative Model Example

Computation Graph

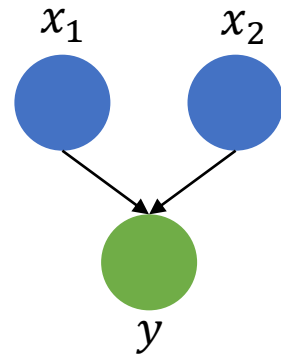


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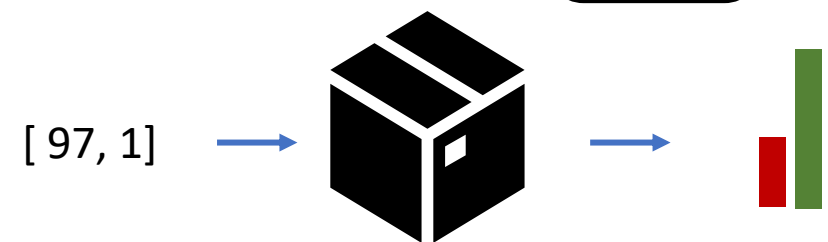
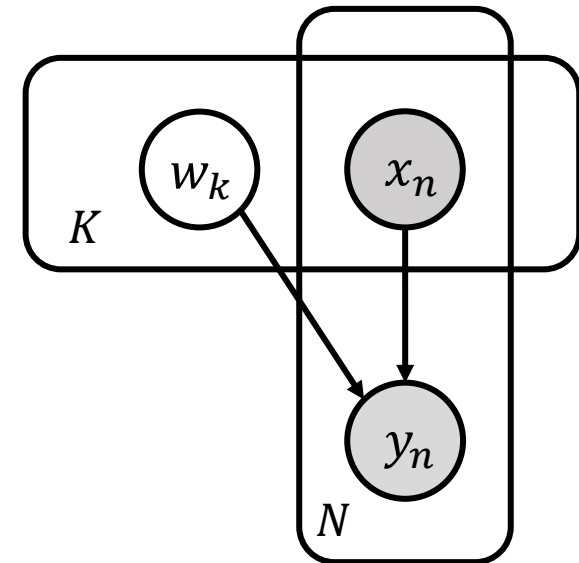
Discriminative Model Example

Computation Graph



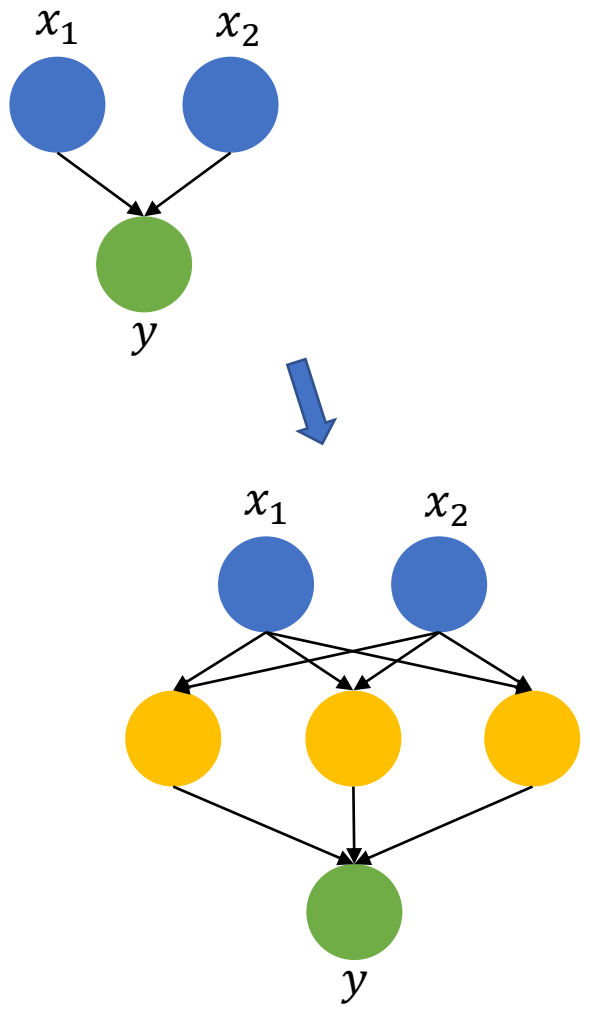
$$p(y = 1) = \frac{1}{1 + e^{w_0 + w_1 x_1 + w_2 x_2}}$$

$\mathbf{w}^T X$
↓
 $\mathbf{w}^T \Phi(X)$



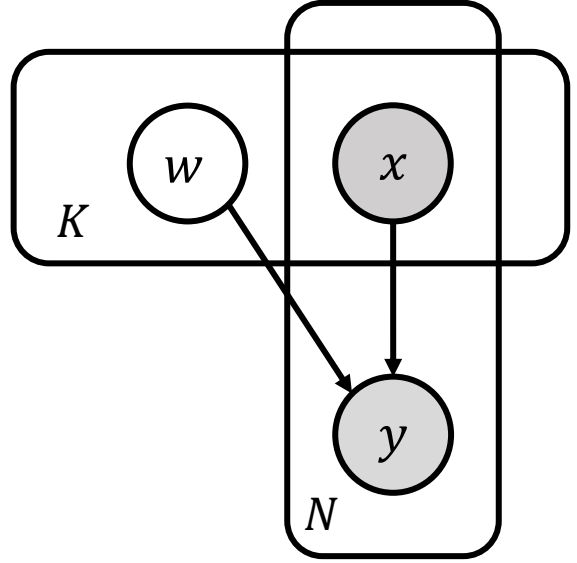
Discriminative Model Example:

Computation Graph



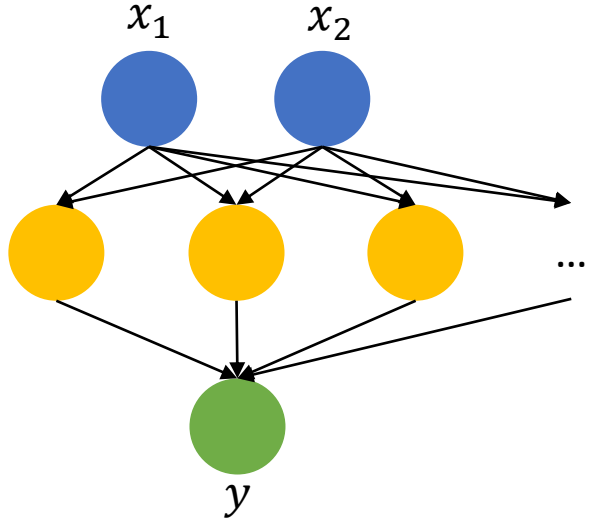
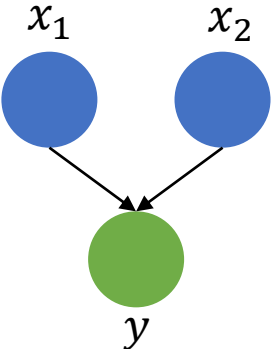
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$w^T X$
↓
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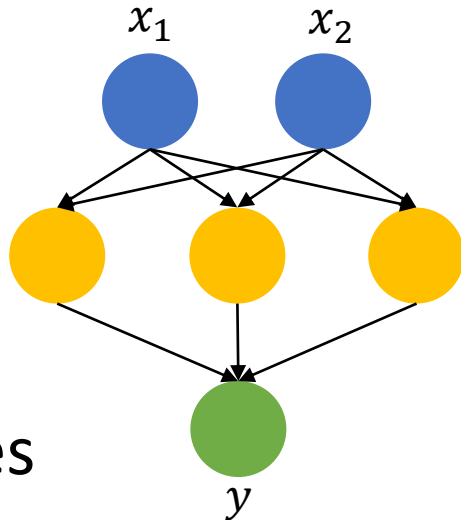


Discriminative Model Example

Computation Graph



Gaussian Processes

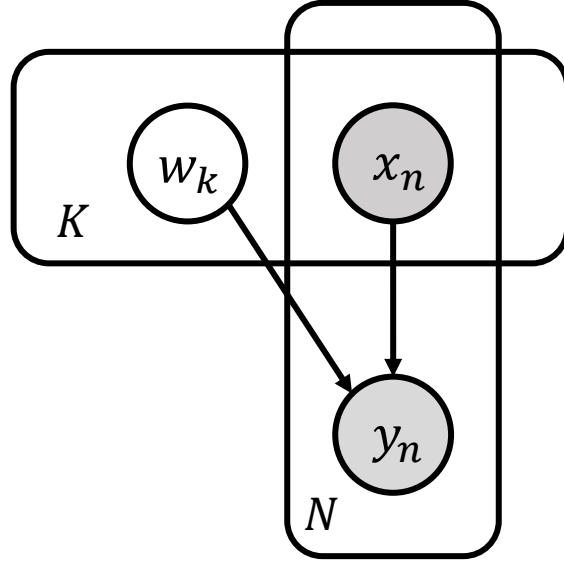


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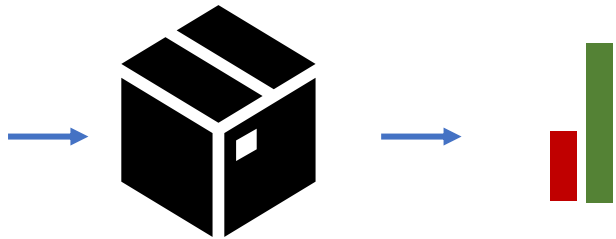
$$W^T X$$

↓

$$W^T \Phi(X)$$

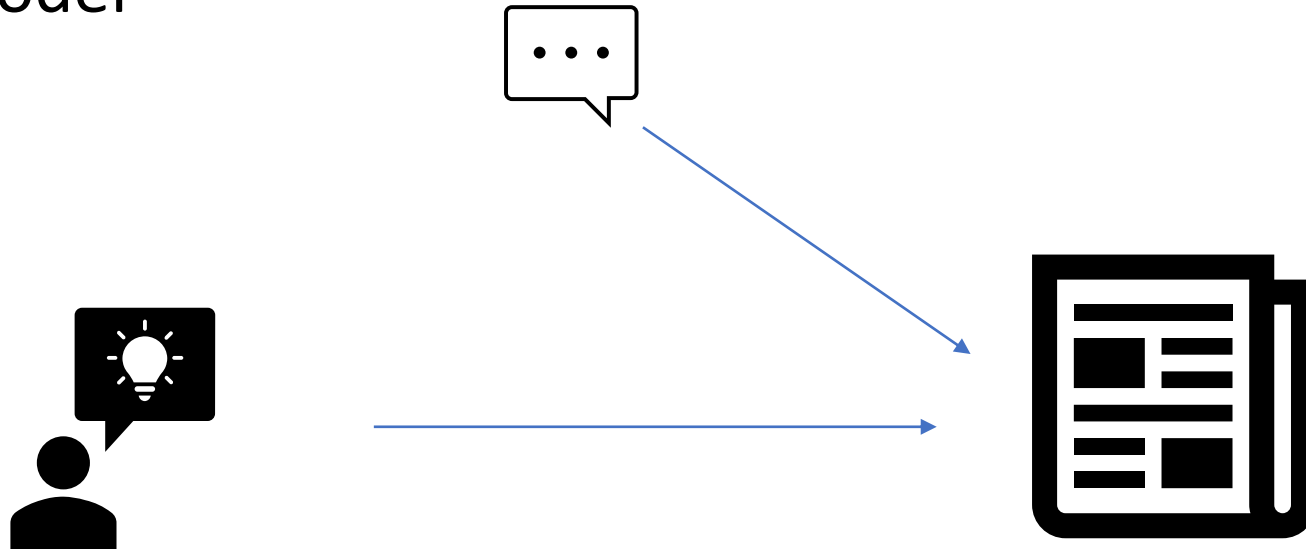


[97, 1]



Generative Model Example

- Topic Model

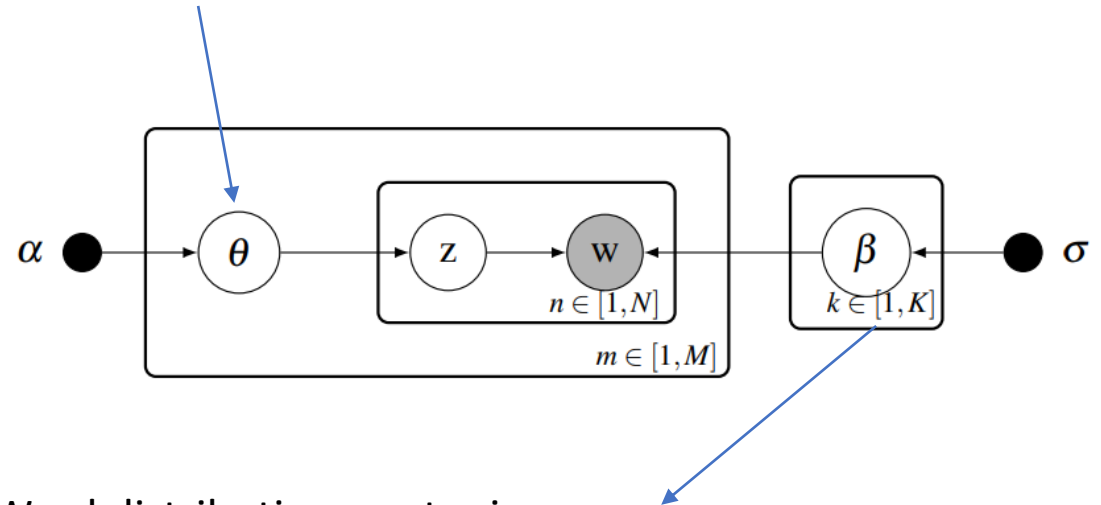


Generative Model Example:

- Latent Dirichlet Allocation

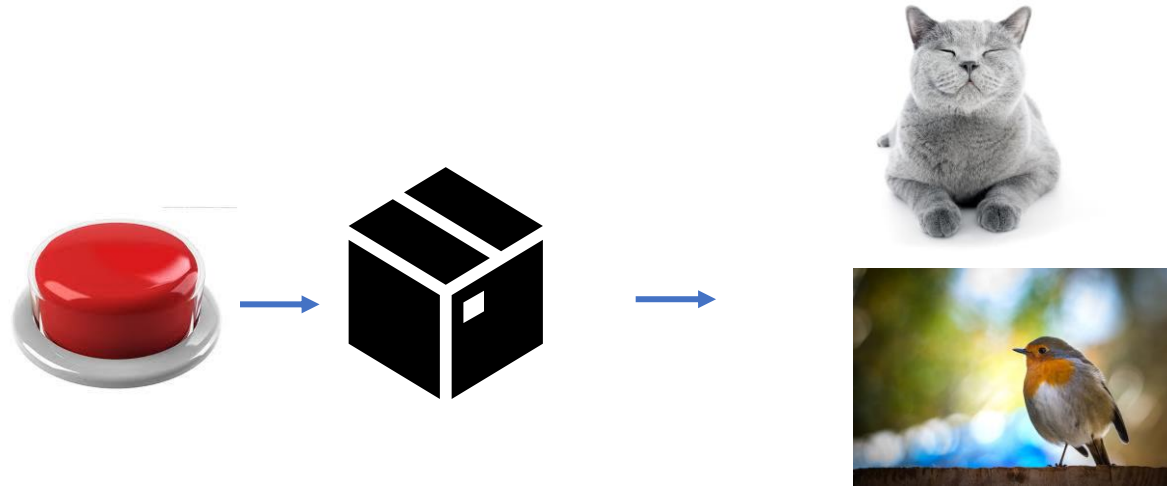
ID	topic	neural	distribution	...
1	15	2	19	...
2	1	13	21	...
3	0	16	1	...

Topic distribution per document:
e.g. 30% “topic model”, 40% “natural language processing”,
30% “interpretability”



Word distribution per topic:
e.g. under “topic model”: “Dirichlet” 2%, “topic” 4%,
“Categorical” 1.5%,

Generative Model Example



Topics

Documents

User embedding

Movie rating

Underlying health conditions

Symptoms

How to infer the unknowns?



The Central Computation for Inference

- Inference: infer the **unknowns**
 - Unobserved/latent variables in the model
 - Quantities depending on the latent variables in the model

The Central Computation for Inference

- Inference: infer the **unknowns**
 - Unobserved/latent variables in the model
 - Quantities depending on the latent variables in the model

$$\int F(\theta) \pi(\theta) d\theta$$

Diagram illustrating the components of the integral expression $\int F(\theta) \pi(\theta) d\theta$:

- $F(\theta)$ is labeled as the **integrand function** (orange text).
- $\pi(\theta)$ is labeled as the **probability measure** (blue text).
- $d\theta$ is labeled as the **prob. density** (green text).
- θ is labeled as the **Random variable (unobserved)** (blue text).

(For discrete probability measures, integration becomes discrete sum.)

Bayesian Inference

$$\pi(\theta) = p(\theta|data)$$

$$P(\theta | data) = \frac{P(\theta)P(data | \theta)}{P(data)}$$

- $P(\theta)$: prior distribution
- $P(data | \theta)$: likelihood of θ given $data$
- $P(\theta | data)$: posterior distribution of θ given $data$
- $P(data)$: marginal likelihood/model evidence

$$P(data) = \int P(\theta)P(data | \theta)$$



Image courtesy of Sebastian Nowozin

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Computation Challenge

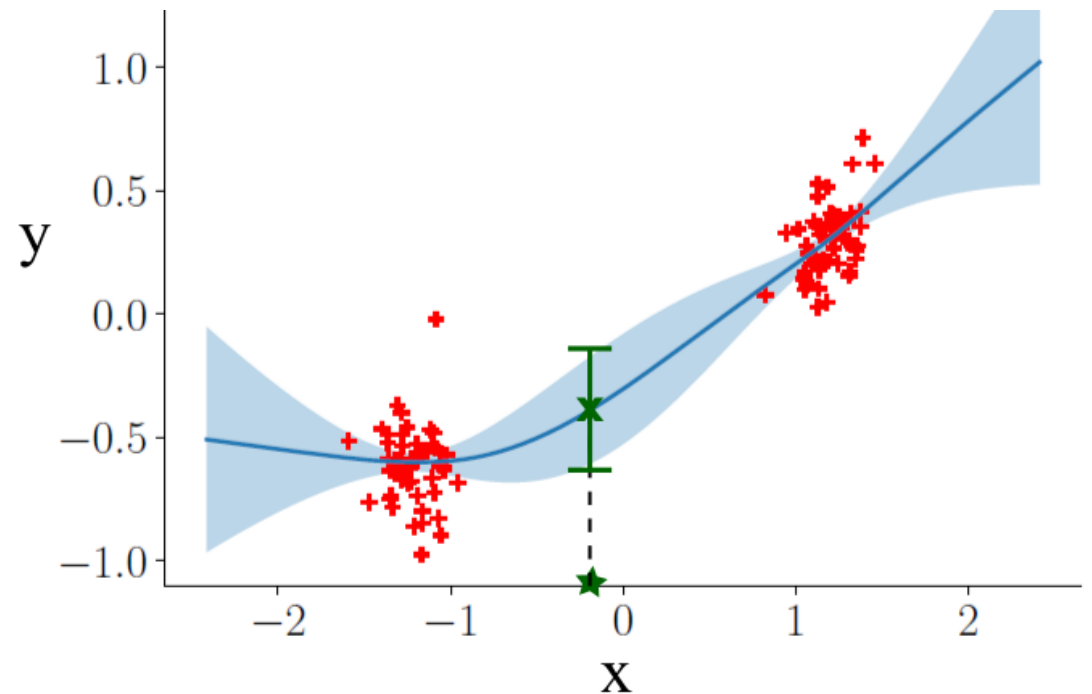
- The central equation for inference:

$$\int F(\theta) \pi(\theta) d\theta$$

“What is the prediction distribution of the **test output** given a **test input**?”

$$F(\theta) = p(y|x, \theta), \pi(\theta) = p(\theta | D),$$

$D =$ observed datapoints



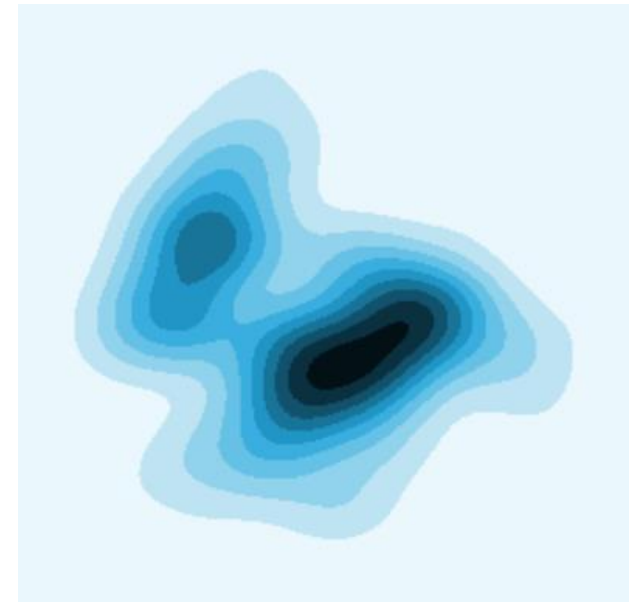
Computation Challenge

- The central equation for inference:

$$\int F(\theta) \pi(\theta) d\theta$$

“What is the mean of this distribution?”

$F(\theta) = \theta$, $\pi(\theta)$ can be complicated and high dimensional



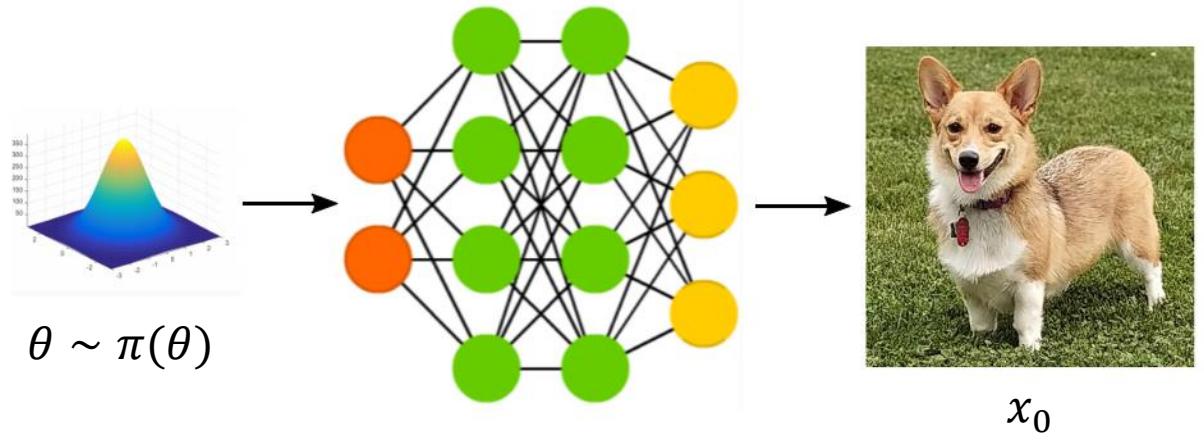
Computation Challenge

- The central equation for inference:

$$\int F(\theta) \pi(\theta) d\theta$$

“What is the probability of generating this image?”

$$F(\theta) = \delta(NN(\theta) = x_0), \pi(\theta) = N(0, I)$$



Computation Challenge

- The central equation for inference:

$$\int F(\theta) \pi(\theta) d\theta$$

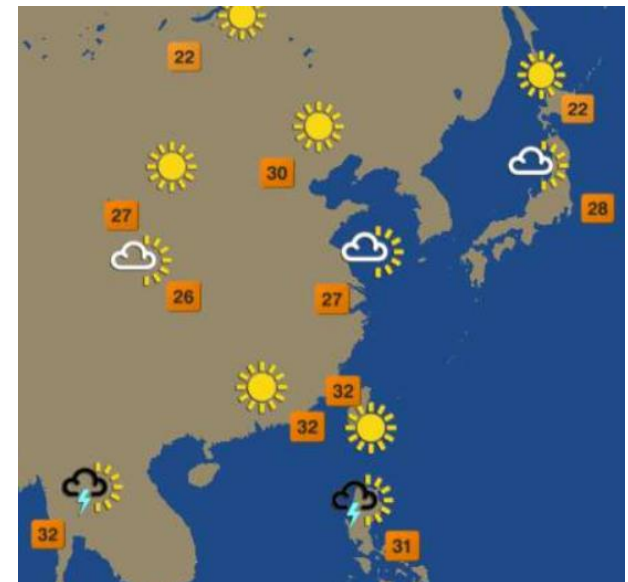
“What is the weather forecast for tomorrow?”

Answering this in a Bayesian way:

θ : forecasting simulator settings

D : historical weather record

$F(\theta) = \text{Simulator}(\theta), \pi(\theta) = p(\theta | D)$



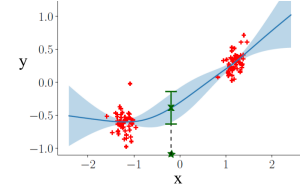
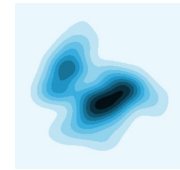
Nature laughs at the
difficulties of integration.

--Pierre-Simon Laplace



Integration in Bayesian Computation

The probability distribution $\pi(\theta)$ is intractable



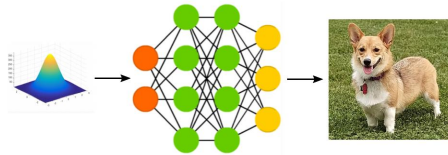
Approximate Inference
(in a strict sense)

$$\int F(\theta) \pi(\theta) d\theta$$

This tutorial

Integration in Bayesian Computation

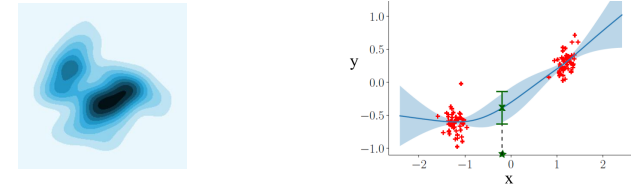
The integrand $F(\theta)$ is intractable



Implicit Models

Bayesian Optimisation, Probabilistic Numerics

The probability distribution $\pi(\theta)$ is intractable



Approximate Inference

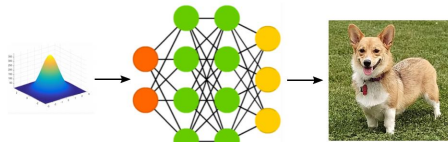
(in a strict sense)

$$\int F(\theta) \pi(\theta) d\theta$$

This tutorial

Integration in Bayesian Computation

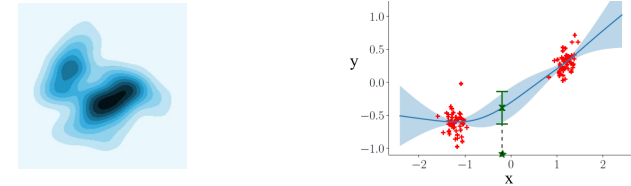
The integrand $F(\theta)$ is intractable



Implicit Models

Bayesian Optimization, Probabilistic Numerics

The probability distribution $\pi(\theta)$ is intractable



Approximate Inference

(in a strict sense)

$$\int F(\theta) \pi(\theta) d\theta$$

This tutorial

Both $F(\theta)$ and $\pi(\theta)$ are intractable



Approximate Bayesian Computation

Approximate Inference

- Central task: approximate $\pi(\theta)$

$$\overset{\text{😊}}{q}(\theta) \approx \overset{\text{😈}}{\pi}(\theta)$$

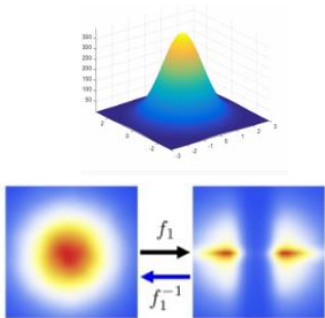
(Assumed $\int F(\theta)q(\theta)d\theta$ can be computed or approximated efficiently.)

Approximate Inference

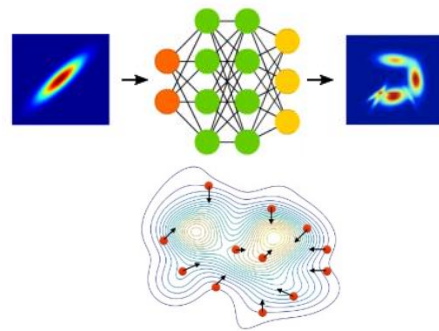
- Central task: approximate $\pi(\theta)$

$$\underbrace{q(\theta)}_{\text{😊}} \approx \pi(\theta)_{\text{😈}}$$

Approximate distribution design



Explicit distributions



Implicit distributions

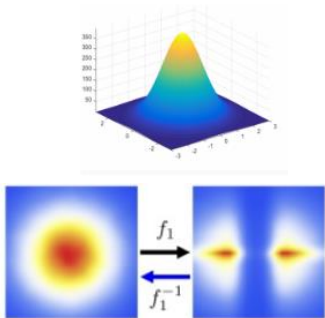
Approximate Inference

- Central task: approximate $\pi(\theta)$

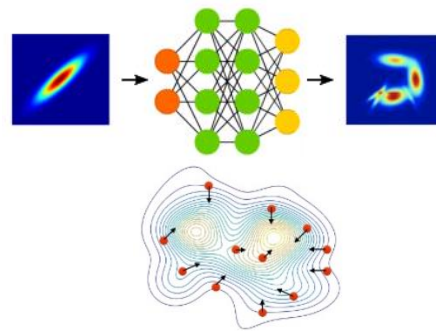
$$\underbrace{q(\theta)}_{\text{😊}} \approx \underbrace{\pi(\theta)}_{\text{😈}}$$

Approximate distribution design

Algorithm for fitting $q(\theta)$ to $\pi(\theta)$



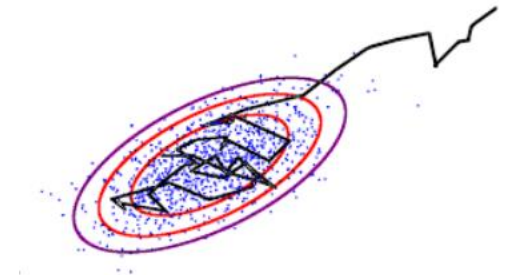
Explicit distributions



Implicit distributions

$$\min \text{Loss}(q(\theta), \pi(\theta))$$

Optimisation-based approaches



Sampling-based approaches

Tutorial Outline



Basics

Probabilistic modelling
Approximate inference
Variational inference



Advances

Scalable variational inference
Monte Carlo techniques
Amortized inference
 q distribution design
Optimization objective design



Applications

Bayesian neural networks
Partially observed VAEs
Future challenges

Bayesian Inference

$$P(\theta | D) = \frac{P(\theta)P(D | \theta)}{P(D)}$$

- $P(\theta)$: prior
- $P(D | \theta)$: likelihood
- $P(\theta | D)$: posterior
- $P(D)$: marginal



Image courtesy of Sebastian Nowozin

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Variational Inference (VI)

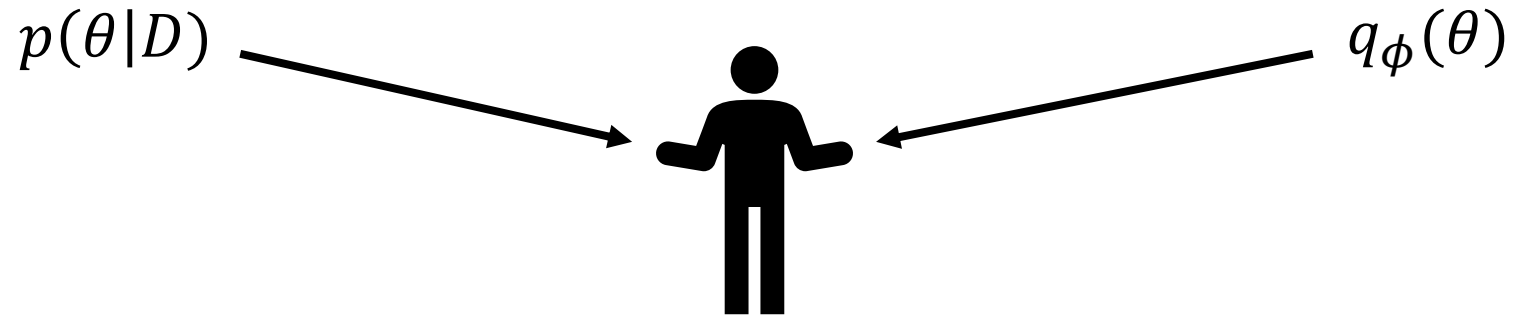
The posterior

$$p(\theta|D) = p(D|\theta)p(\theta)/p(D)$$

The variational distribution

$$q_{\phi}(\theta)$$

Inference as Optimization



Kullback-Leibler (KL) divergence

Kullback-Leibler Divergence

$$KL[q(\theta)||p(\theta)] = -\int q(\theta) \log \frac{p(\theta)}{q(\theta)} d\theta = E_{q(\theta)} \left[\log \frac{p(\theta)}{q(\theta)} \right]$$

- When $p = q$, KL is 0
- Otherwise, $KL > 0$
- It measures how similar are these two distributions

Let's Derive the Objective of VI

- Minimize $KL[q(\theta)||p(\theta|D)]$

$$KL[q(\theta)||p(\theta|D)] = -E_{q(\theta)} \left[\log \frac{p(\theta|D)}{q(\theta)} \right]$$

Let's Derive the Objective of VI

- Minimize $KL[q(\theta)||p(\theta|D)]$

$$KL[q(\theta)||p(\theta|D)] = -E_{q(\theta)} \left[\log \frac{p(\theta|D)}{q(\theta)} \right]$$

$$= -E_{q(\theta)} \left[\log \frac{p(\theta,D)}{p(D)q(\theta)} \right] = -E_{q(\theta)} \left[\log \frac{p(\theta,D)}{q(\theta)} - \log p(D) \right]$$

Let's Derive the Objective of VI

- Minimize $KL[q(\theta)||p(\theta|D)]$

$$KL[q(\theta)||p(\theta|D)] = -E_{q(\theta)} \left[\log \frac{p(\theta|D)}{q(\theta)} \right]$$

$$= -E_{q(\theta)} \left[\log \frac{p(\theta,D)}{p(D)q(\theta)} \right] = -E_{q(\theta)} \left[\log \frac{p(\theta,D)}{q(\theta)} - \log p(D) \right]$$

$$= \boxed{\log p(D)} - E_{q(\theta)} \left[\log \frac{p(\theta,D)}{q(\theta)} \right]$$


Model Evidence

Let's Derive the Objective of VI

Minimize $KL[q(\theta)||p(\theta|D)]$

$$KL[q(\theta)||p(\theta|D)] = \log p(D) - E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right]$$

Maximize $E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right]$



Let's Derive the Objective of VI

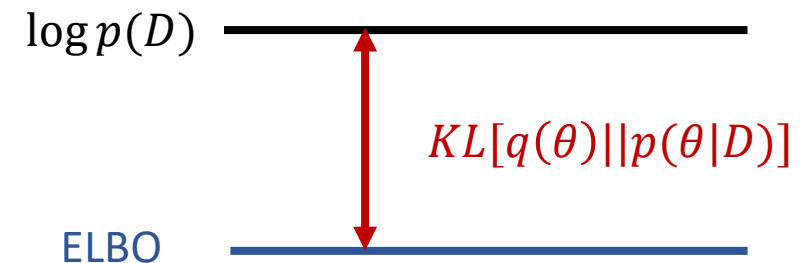
Minimize $KL[q(\theta)||p(\theta|D)]$

$$KL[q(\theta)||p(\theta|D)] = \boxed{\log p(D)} - E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right]$$

Model Evidence

Maximize $L = E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right]$

Evidence Lower Bound (ELBO)



“Model Evidence = ELBO + KL”

Alternative Derivation

Let's start with the model evidence

$$\log p(D)$$

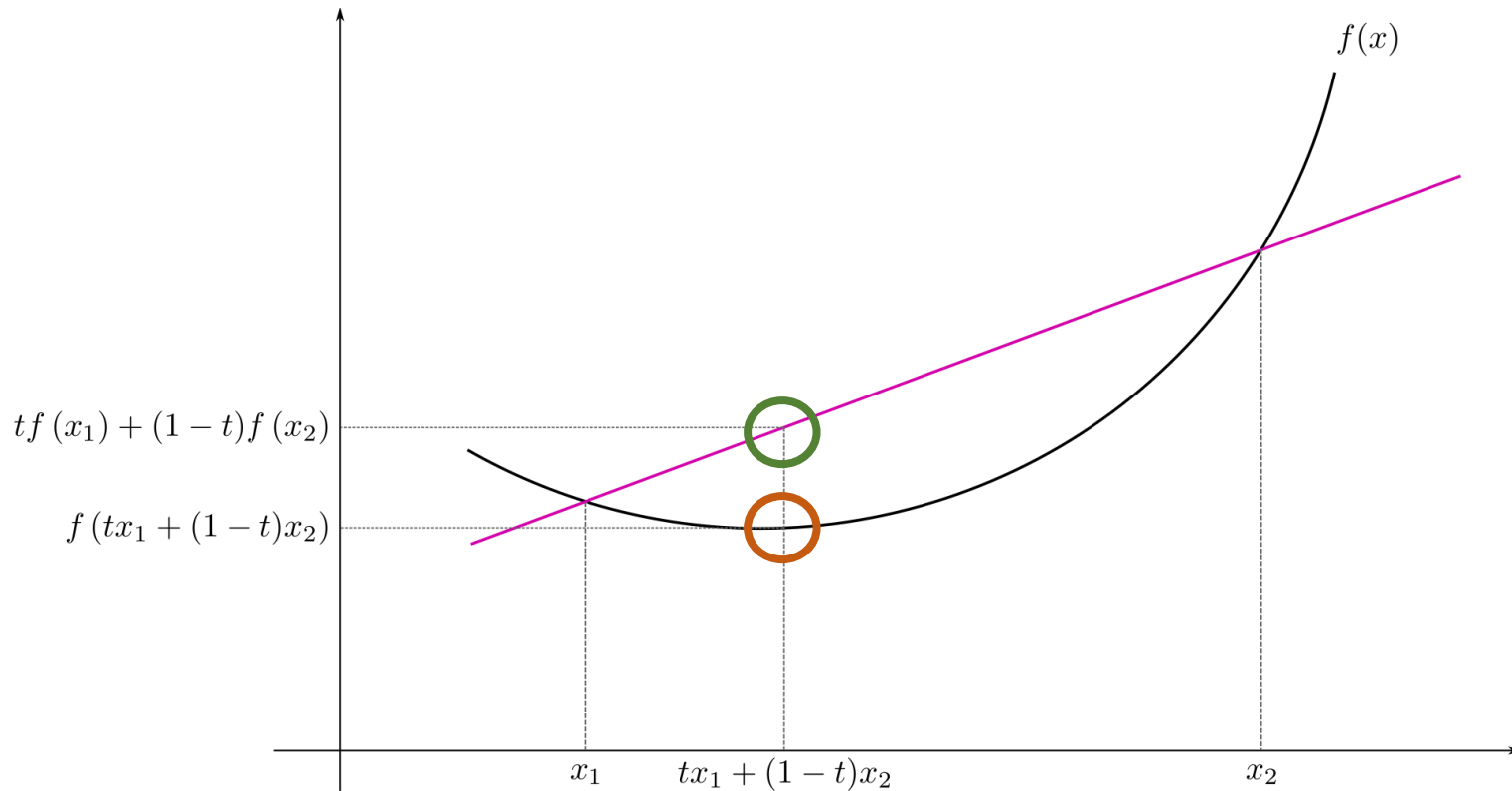
Alternative Derivation

$$\log p(D) = \log \int p(\theta, D) d\theta$$

Alternative Derivation

$$\begin{aligned}\log p(D) &= \log \int p(\theta, D) d\theta \\ &= \log \int \frac{p(\theta, D) q(\theta)}{q(\theta)} d\theta \\ &= \log E_{q(\theta)} \left[\frac{p(\theta, D)}{q(\theta)} \right]\end{aligned}$$

Jensen's Inequality



If f is a convex function, then

$$f(E[X]) \leq E[f(X)]$$


If f is a concave function, then

$$f(E[X]) \geq E[f(X)]$$

Alternative Derivation

$$\begin{aligned}\log p(D) &= \log \int p(\theta, D) d\theta \\ &= \log \int \frac{p(\theta, D) q(\theta)}{q(\theta)} d\theta\end{aligned}$$

$$= \log E_{q(\theta)} \left[\frac{p(\theta, D)}{q(\theta)} \right]$$

Jensen's inequality  $\geq E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right]$

Log is a concave function, then
 $f(E[X]) \geq E[f(X)]$

Alternative Derivation

Model Evidence

$$\log p(D) = \log \int p(\theta, D) d\theta$$

$$= \log \int \frac{p(\theta, D) q(\theta)}{q(\theta)} d\theta$$

$$= \log E_{q(\theta)} \left[\frac{p(\theta, D)}{q(\theta)} \right]$$

$$\geq E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right]$$

Evidence Lower Bound (ELBO)

Variational Inference (VI)

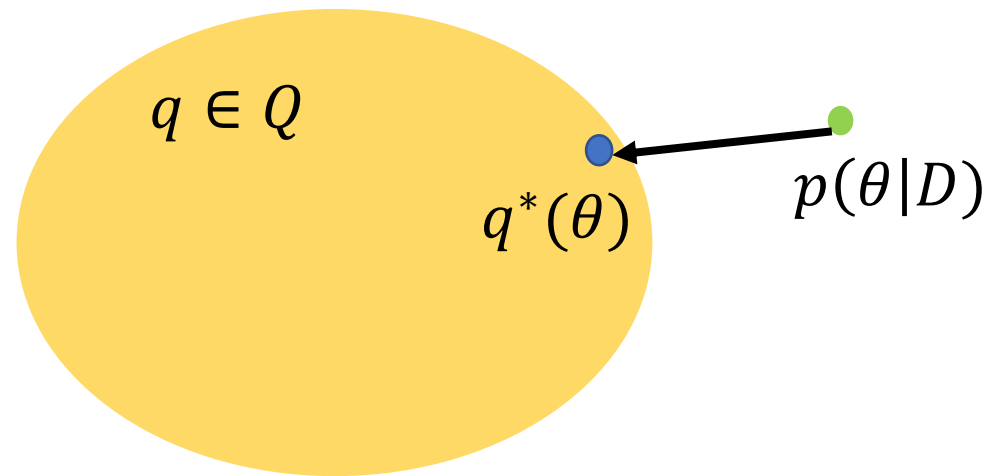
The posterior

$$p(\theta|D) = p(D|\theta)p(\theta)/p(D)$$

The variational distribution

$$q_\phi(\theta)$$

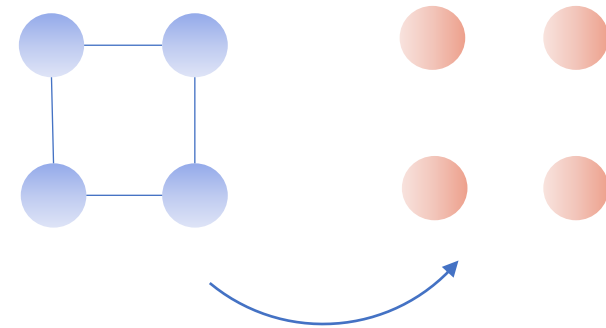
$$L = E_{q_\phi(\theta)} \left[\log \frac{p(D, \theta)}{q_\phi(\theta)} \right] = \log p(D) - KL[q_\phi(\theta) || p(\theta)]$$



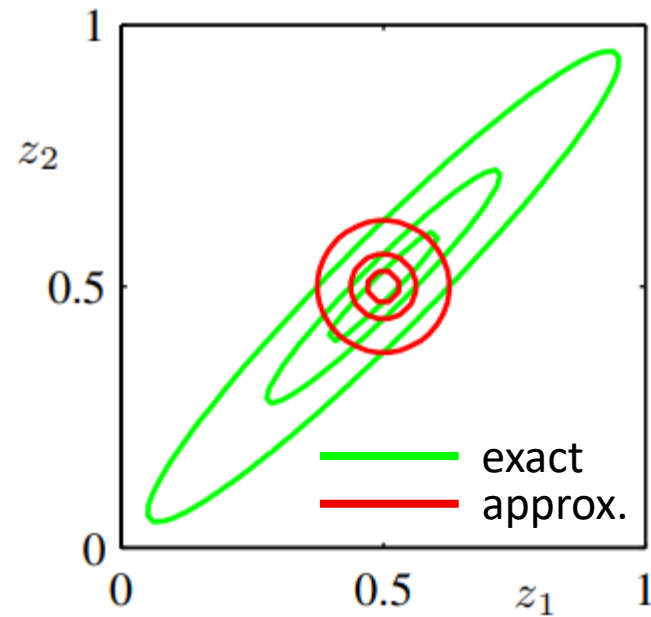
Mean-field Variational Inference

- A type of choices of the variational distribution
- The name origins in the mean field theory of physics
- The variational distribution factorizes

$$q_{\phi}(\boldsymbol{\theta}) = \prod_{i=1}^K q_{\phi_i}(\theta_i)$$



A Gaussian Example



$$p(\mathbf{z}) = N(\mathbf{z}|\mu, \Lambda^{-1})$$

$$q(\mathbf{z}) = q(z_1)q(z_2)$$

Mean-field Variational Inference

ELBO Fully Factorized Variational Distribution

$$L = E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right] \longleftarrow q(\boldsymbol{\theta}) = \prod_{i=1}^K q_{\phi_i}(\theta_i)$$

Mean-field Variational Inference

ELBO

Fully Factorized Variational Distribution

$$L = E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right] \longleftarrow q(\boldsymbol{\theta}) = \prod_{i=1}^K q_{\phi_i}(\theta_i)$$

$$L = \int q(\theta_j) E_{q(\theta_{-j})} [\log p(\theta_j, D | \theta_{-j})] d\theta_j - \int q(\theta_j) \log p(\theta_j) d\theta_j + c_j$$

Mean-field Variational Inference

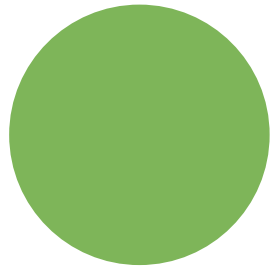
ELBO

Fully Factorized Variational Distribution

$$L = E_{q(\theta)} \left[\log \frac{p(\theta, D)}{q(\theta)} \right] \longleftarrow q(\boldsymbol{\theta}) = \prod_{i=1}^K q_{\phi_i}(\theta_i)$$

$$L = \int q(\theta_j) E_{q(\theta_{-j})} [\log p(\theta_j, D | \theta_{-j})] d\theta_j - \int q(\theta_j) \log p(\theta_j) d\theta_j + c_j$$

$$q^*(\theta_j) \propto \exp(E_{q(\theta_{-j})} [\log p(\theta_j, D | \theta_{-j})])$$

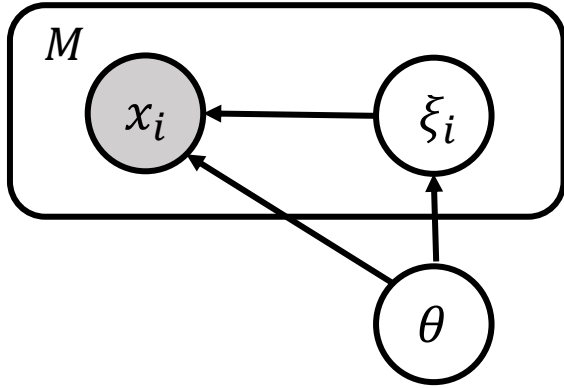


The slide features decorative curved lines in the top corners, transitioning from light green to light blue. The main content is centered on the left side.

Part II: Advances

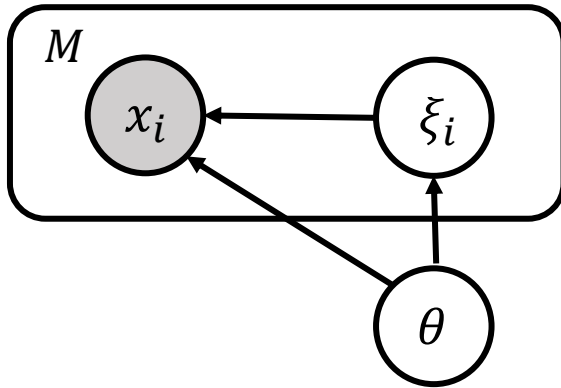
- **Scalable variational inference**
- **Monte Carlo methods**
- **Amortized inference**
- Approximate distribution design
- Optimization objective design

Stochastic Variational Inference



$$p(\theta, \xi, \mathbf{x}) = p(\theta) \prod_{i=1}^M p(\xi_i | \theta) p(x_i | \xi_i, \theta)$$

Stochastic Variational Inference

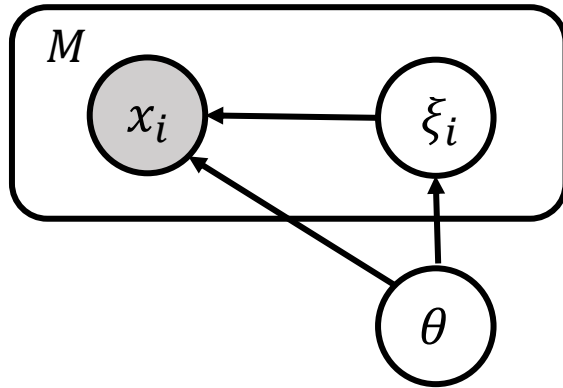


$$p(\theta, \xi, \mathbf{x}) = p(\theta) \prod_{i=1}^M p(\xi_i | \theta) p(x_i | \xi_i, \theta)$$

$$\begin{aligned} L &= E_q \left[\log \frac{p(\theta, \xi, \mathbf{x})}{q(\theta, \xi)} \right] \\ &= E_q \left[\log \frac{p(\theta) \prod_{i=1}^M p(\xi_i | \theta) p(x_i | \xi_i, \theta)}{q(\theta) \prod_{i=1}^M q(\xi_i)} \right] \\ &= E_q \left[\log \frac{p(\theta)}{q(\theta)} \right] + \sum_{i=1}^M E_q \left[\log \frac{p(\xi_i | \theta) p(x_i | \xi_i)}{q(\xi_i)} \right] \end{aligned}$$

- $O(M)$ time to compute in each update iteration
- M can be extremely large
- Even one iteration might not be affordable

Stochastic Variational Inference



$$p(\theta, \xi, \mathbf{x}) = p(\theta) \prod_{i=1}^M p(\xi_i | \theta) p(x_i | \xi_i, \theta)$$

$$L = E_q \left[\log \frac{p(\theta)}{q(\theta)} \right] + \sum_{i=1}^M E_q \left[\log \frac{p(\xi_i | \theta) p(x_i | \xi_i)}{q(\xi_i)} \right]$$

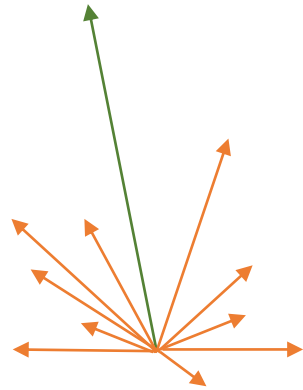


stochastic approximation
with $S \ll M$

$$\hat{L} = E_q \left[\log \frac{p(\theta)}{q(\theta)} \right] + \frac{M}{S} \sum_{i=1}^S E_q \left[\log \frac{p(\xi_i | \theta) p(x_i | \xi_i)}{q(\xi_i)} \right]$$

Computational complexity: $O(M) \rightarrow O(S)$

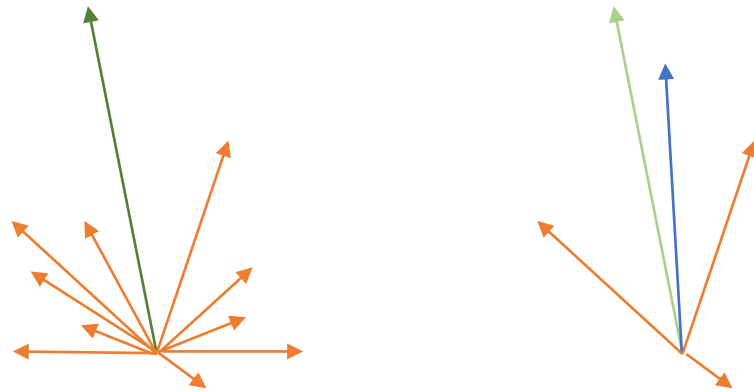
How Stochastic Gradient Works



—→ $\nabla F(x_i)$ gradient of each single data point x_i

—→ $E_x[\nabla F(x)]$ batch gradient considering all data points

How Stochastic Gradient Works

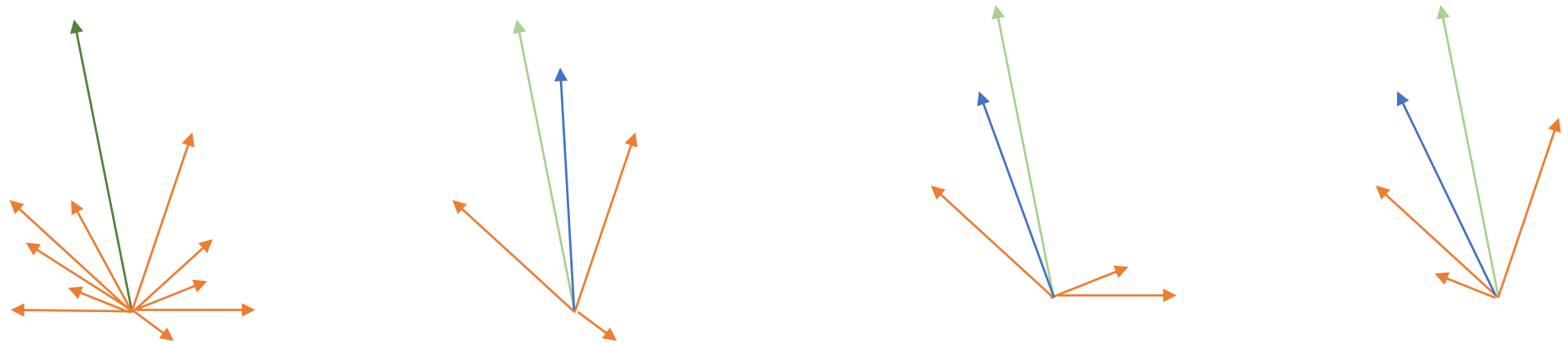


→ $\nabla F(x_i)$ gradient of each single data point x_i

→ $E_x[\nabla F(x)]$ batch gradient considering all $M = 10$ data points

→ $\frac{M}{S} \sum_{s=1}^S \nabla F(x_s)$ mini-batch gradient/stochastic gradient estimated using $S=3$ data points

How Stochastic Gradient Works



→ $\nabla F(x_i)$ gradient of each single data point x_i

→ $E_x[\nabla F(x)]$ batch gradient considering all $M = 10$ data points

→ $\frac{M}{S} \sum_{s=1}^S \nabla F(x_s)$ mini-batch gradient/stochastic gradient estimated using $S=3$ data points

Stochastic Variational Inference

$$L = E_q \left[\log \frac{p(\theta)}{q(\theta)} \right] + \sum_{i=1}^M E_q \left[\log \frac{p(\xi_i|\theta)p(x_i|\xi_i)}{q(\xi_i)} \right] \xrightarrow{\text{gradient}} \nabla L = \nabla E_q \left[\log \frac{p(\theta)}{q(\theta)} \right] + \sum_{i=1}^M E_q \left[\nabla \log \frac{p(\xi_i|\theta)p(x_i|\xi_i)}{q(\xi_i)} \right]$$

↓ Stochastic approximation with $S \ll M$ ↓

$$\hat{L} = E_q \left[\log \frac{p(\theta)}{q(\theta)} \right] + \frac{M}{S} \sum_{i=1}^S E_q \left[\log \frac{p(\xi_i|\theta)p(x_i|\xi_i)}{q(\xi_i)} \right] \xrightarrow{\text{gradient}} \nabla \hat{L} = \nabla E_q \left[\log \frac{p(\theta)}{q(\theta)} \right] + \frac{M}{S} \sum_{i=1}^S E_q \left[\nabla \log \frac{p(\xi_i|\theta)p(x_i|\xi_i)}{q(\xi_i)} \right]$$

Stochastic Gradient

Nature laughs at the
difficulties of integration.

--Pierre-Simon Laplace



Monte Carlo Approximation

- To approximate: $E_{p(x)}[f(x)]$
- MC Approximation:
 1. Sample $x_1, x_2, \dots, x_K \sim p(x)$
 2. Evaluate $f(x_i)$ for each sample
 3. Compute $E[f(x)] \approx \frac{1}{K} \sum_{i=1}^K f(x_i)$
Unbiased Monte Carlo estimate

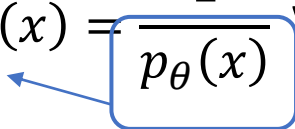
Log-derivative Trick

$$\nabla_{\theta} \log p_{\theta}(x)$$

Log-derivative Trick

$$\nabla_{\theta} \log p_{\theta}(x) = \frac{1}{p_{\theta}(x)} \nabla p_{\theta}(X)$$

Log-derivative Trick

$$\nabla_{\theta} \log p_{\theta}(x) = \frac{1}{p_{\theta}(x)} \nabla p_{\theta}(x)$$


$$\nabla_{\theta} p_{\theta}(x) = p_{\theta}(x) \nabla_{\theta} \log p_{\theta}(x)$$

REINFORCE Gradients

$$\text{ELBO } L = E_{q_{\phi}(\theta)} \left[\log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right]$$

Gradient of the ELBO

$$\nabla_{\phi} L = \nabla_{\phi} E_{q_{\phi}(\theta)} \left[\log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right] = \int \nabla_{\phi} \left\{ q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right\} d\theta$$

Glynn (1990). Likelihood ratio gradient estimation for stochastic systems. *Communications of the ACM*, 33(10), 75–84.

Williams (1992). Simple statistical gradient-following algorithms for connectionist reinforcement learning. *Machine Learning*, 8(3-4), 229–256.

Fu (2006). Gradient estimation. *Handbooks in Operations Research and Management Science*, 13, 575–616.

REINFORCE Gradients

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Glynn (1990). Likelihood ratio gradient estimation for stochastic systems. *Communications of the ACM*, 33(10), 75–84.

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REINFORCE Gradients

Log-derivative trick

$$\nabla q_{\phi}(\theta) = q_{\phi}(\theta) \nabla_{\phi} \log q_{\phi}(\theta)$$

$$\text{ELBO } L = E_{q_{\phi}(\theta)} \left[\log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right]$$

Gradient of the ELBO

$$\begin{aligned} \nabla_{\phi} L &= \nabla_{\phi} E_{q_{\phi}(\theta)} \left[\log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right] = \int \nabla_{\phi} \left\{ q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right\} d\theta \\ &= \int \nabla_{\phi} q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} d\theta + \int q_{\phi}(\theta) \nabla_{\phi} \log \frac{p(\theta, D)}{q_{\phi}(\theta)} d\theta \\ &= \int \nabla_{\phi} q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} d\theta - \int \nabla_{\phi} q_{\phi}(\theta) d\theta \end{aligned}$$

Glynn (1990). Likelihood ratio gradient estimation for stochastic systems. *Communications of the ACM*, 33(10), 75–84.

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REINFORCE Gradients

$$\text{ELBO } L = E_{q_\phi(\theta)} \left[\log \frac{p(\theta, D)}{q_\phi(\theta)} \right]$$

Gradient of the ELBO

$$\nabla_\phi L = \nabla_\phi E_{q_\phi(\theta)} \left[\log \frac{p(\theta, D)}{q_\phi(\theta)} \right] = \int \nabla_\phi \left\{ q_\phi(\theta) \log \frac{p(\theta, D)}{q_\phi(\theta)} \right\} d\theta$$

$$= \int \nabla_\phi q_\phi(\theta) \log \frac{p(\theta, D)}{q_\phi(\theta)} d\theta + \int q_\phi(\theta) \nabla_\phi \log \frac{p(\theta, D)}{q_\phi(\theta)} d\theta$$

$$= \int q_\phi(\theta) \nabla_\phi \log q_\phi(\theta) \log \frac{p(\theta, D)}{q_\phi(\theta)} d\theta - \int \nabla_\phi q_\phi(\theta) d\theta$$

$$= \nabla_\phi \int q_\phi(\theta) d\theta = \nabla_\phi 1 = 0$$

$$\begin{aligned} & \nabla_\phi \log \frac{p(\theta, D)}{q_\phi(\theta)} \\ &= \frac{q_\phi(\theta)}{p(\theta, D)} \left(-\frac{p(\theta, D)}{q_\phi(\theta)^2} \right) \nabla q_\phi(\theta) \\ &= -\frac{\nabla_\phi q_\phi(\theta)}{q_\phi(\theta)} \end{aligned}$$

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REINFORCE Gradients

$$\text{ELBO } L = E_{q_{\phi}(\theta)} \left[\log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right]$$

Gradient of the ELBO

$$\begin{aligned} \nabla_{\phi} L &= \nabla_{\phi} E_{q_{\phi}(\theta)} \left[\log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right] = \int \nabla_{\phi} \left\{ q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right\} d\theta \\ &= \int \nabla_{\phi} q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} d\theta + \int q_{\phi}(\theta) \nabla_{\phi} \log \frac{p(\theta, D)}{q_{\phi}(\theta)} d\theta \\ &= \int q_{\phi}(\theta) \nabla_{\phi} \log q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} d\theta - \underbrace{\int \nabla_{\phi} q_{\phi}(\theta) d\theta}_{= \nabla_{\phi} \int q_{\phi}(\theta) d\theta = \nabla_{\phi} 1 = 0} \\ &= E_{q_{\phi}(\theta)} \left[\underbrace{\nabla_{\phi} \log q_{\phi}(\theta)}_{\text{Score function}} \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right] \end{aligned}$$

Glynn (1990). Likelihood ratio gradient estimation for stochastic systems. *Communications of the ACM*, 33(10), 75–84.

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BBVI

$$\text{ELBO } L = E_{q_{\phi}(\theta)} \left[\log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right] f(\theta)$$

Gradient of the ELBO

$$\nabla_{\phi} L = E_{q_{\phi}(\theta)} \left[\underbrace{\nabla_{\phi} \log q_{\phi}(\theta)}_{\text{Score function}} \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right] f(\theta)$$

- To approximate: $E[f(x)]$

- MC Approximation:

1. Sample $x_1, x_2, \dots, x_K \sim p(x)$
2. Evaluate $f(x_i)$ for each sample
3. Compute $E[f(x)] \approx \frac{1}{K} \sum_{i=1}^K f(x_i)$



Ranganath et al. Black box variational inference. AISTATS 2014

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BBVI

Gradient of the ELBO

$$\nabla_{\phi} L = E_{q_{\phi}(\theta)} \left[\nabla_{\phi} \log q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right]$$

- To approximate: $E[f(x)]$
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 2. Evaluate $f(x_i)$ for each sample
 3. Compute $E[f(x)] \approx \frac{1}{K} \sum_{i=1}^K f(x_i)$



1. Sample $\theta_1, \theta_2, \dots, \theta_K \sim q_{\phi}(\theta)$

BBVI

Gradient of the ELBO

$$\nabla_{\phi} L = E_{q_{\phi}(\theta)} \left[\nabla_{\phi} \log q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right]$$

- To approximate: $E[f(x)]$
- MC Approximation:
 1. Sample $x_1, x_2, \dots, x_K \sim p(x)$
 2. Evaluate $f(x_i)$ for each sample
 3. Compute $E[f(x)] \approx \frac{1}{K} \sum_{i=1}^K f(x_i)$



1. Sample $\theta_1, \theta_2, \dots, \theta_K \sim q_{\phi}(\theta)$
2. Evaluate $\nabla_{\phi} \log q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)}$ for each sample

BBVI

Gradient of the ELBO

$$\nabla_{\phi} L = E_{q_{\phi}(\theta)} \left[\nabla_{\phi} \log q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)} \right]$$

- To approximate: $E[f(x)]$

- MC Approximation:

1. Sample $x_1, x_2, \dots, x_K \sim p(x)$
2. Evaluate $f(x_i)$ for each sample
3. Compute $E[f(x)] \approx \frac{1}{K} \sum_{i=1}^K f(x_i)$



1. Sample $\theta_1, \theta_2, \dots, \theta_K \sim q_{\phi}(\theta)$
2. Evaluate $\nabla_{\phi} \log q_{\phi}(\theta) \log \frac{p(\theta, D)}{q_{\phi}(\theta)}$ for each sample
3. The approximated gradient is:

$$\nabla_{\phi} \hat{L} = \frac{1}{K} \sum_{i=1}^K \nabla_{\phi} \log q_{\phi}(\theta_i) \log \frac{p(\theta_i, D)}{q_{\phi}(\theta_i)}$$

Black-box Variational Inference (BBVI)



Go beyond conjugate exponential family

A.2. Update of ρ

The partial derivative is given below:

$$\begin{aligned} \text{Let } \bar{\delta} &= \sum_{j=1}^M \prod_{i=1}^T \left(\sum_{k=1}^K \rho_{jk} \zeta_{jk} \exp\left(\frac{1}{2} \mu_{jk}\right) \right) \\ \frac{\partial \mathcal{L}}{\partial \rho_{jk}} &= \sum_{i=1}^T \zeta_{jk} \mathbb{E}[\log p(w_{jk} | \theta_i)] + \mathbb{E}[\log h_i] - 1 - \log \rho_{jk} \\ &\quad + \mu_{jk} \left(\frac{1}{2} \sum_{i=1}^T \zeta_{jk} \right) \\ &\quad - \bar{\delta}^{-1} \left(\sum_{i=1}^T \left(\prod_{m=1}^T \left(\sum_{k=1}^K \rho_{jm} \zeta_{jm} \exp\left(\frac{1}{2} \mu_{jm}\right) \right) \right) \right. \\ &\quad \left. \frac{\zeta_{jk} \exp\left(\frac{1}{2} \mu_{jk}\right)}{\sum_{i=1}^T \left(\sum_{k=1}^K \rho_{jk} \zeta_{jk} \exp\left(\frac{1}{2} \mu_{jk}\right) \right)} \right). \end{aligned} \quad (18)$$

A.3. Update of ζ

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial \zeta_{jk}} &= \sum_{i=1}^T \left(\zeta_{jk} \left(\sum_{m=1}^T \rho_{jm} \sum_{k=1}^K (\Psi(\lambda_{jk}) - \Psi(\sum_{l=1}^K \lambda_{jl})) |w_{jm} = i \right) \right. \\ &\quad \left. + \zeta_{jk} (\Psi(a_{1j}) - \Psi(a_{1j} + a_{2j})) + \sum_{i=1}^T (\Psi(a_{2i}) - \Psi(a_{1i} + a_{2i})) \right) \\ &\quad - \zeta_{jk} \log \zeta_{jk} - \sum_{i=1}^T \left(\mu_{jk}^i \left(\frac{1}{2} \sum_{i=1}^T \rho_{jk} \zeta_{jk} \right) \right. \\ &\quad \left. - \log \left(\sum_{i=1}^T \prod_{m=1}^T \left(\sum_{k=1}^K \rho_{jm} \zeta_{jm} \exp\left(\frac{1}{2} \mu_{jm}\right) \right) \right) \right). \end{aligned} \quad (19)$$

where Ψ is the digamma function. For $i \in \{1, \dots, T\}$, we write:

$$h_i = \sum_{j=1}^M \left(\prod_{m=1}^T \left(\sum_{k=1}^K \rho_{jm} \zeta_{jm} \exp\left(\frac{1}{2} \mu_{jm}\right) \right) \right) \left(\sum_{k=1}^K \rho_{jk} \exp\left(\frac{1}{2} \mu_{jk}\right) \right),$$

which yields:

$$\sum_{i=1}^T \prod_{m=1}^T \left(\sum_{k=1}^K \rho_{jm} \zeta_{jm} \exp\left(\frac{1}{2} \mu_{jm}\right) \right) = \sum_{i=1}^T h_i \zeta_{jk}^{-1}.$$

$\mathcal{L}_{\zeta_{jk}}$ can now be rewritten as:

$$\begin{aligned} \mathcal{L}_{\zeta_{jk}} &= \sum_{i=1}^T \left(\zeta_{jk} \left(\sum_{m=1}^T \rho_{jm} \sum_{k=1}^K (\Psi(\lambda_{jk}) - \Psi(\sum_{l=1}^K \lambda_{jl})) |w_{jm} = i \right) \right. \\ &\quad \left. + \zeta_{jk} (\Psi(a_{1j}) - \Psi(a_{1j} + a_{2j})) + \sum_{i=1}^T (\Psi(a_{2i}) - \Psi(a_{1i} + a_{2i})) \right) \\ &\quad - \zeta_{jk} \log \zeta_{jk} - \sum_{i=1}^T \left(\mu_{jk}^i \left(\frac{1}{2} \sum_{i=1}^T \rho_{jk} \zeta_{jk} \right) - \log(h_i \zeta_{jk}^{-1}) \right). \end{aligned} \quad (20)$$

We follow the approach of [14] to derive the fixed point update. Suppose we have a previous value ζ_{jk}^{old} . Consider

the inequality $\log(x) \leq x^{-1} + \log(x) - 1$, where equality holds if and only if $x = e$. Thus, set $x = h_i \zeta_{jk}^{old}$ and $x = h_i \zeta_{jk}^{new}$. The new bound becomes:

$$\begin{aligned} \mathcal{L}_{\zeta_{jk}} &\geq \sum_{i=1}^T \sum_{j=1}^M \left(\zeta_{jk} \left(\sum_{m=1}^T \rho_{jm} \sum_{k=1}^K (\Psi(\lambda_{jk}) - \Psi(\sum_{l=1}^K \lambda_{jl})) |w_{jm} = i \right) \right. \\ &\quad \left. + \zeta_{jk} (\Psi(a_{1j}) - \Psi(a_{1j} + a_{2j})) \right. \\ &\quad \left. + \sum_{i=1}^T (\Psi(a_{2i}) - \Psi(a_{1i} + a_{2i})) - \zeta_{jk} \log \zeta_{jk} \right) \\ &\quad + \sum_{i=1}^T \left(\mu_{jk}^i \left(\frac{1}{2} \sum_{i=1}^T \rho_{jk} \zeta_{jk} \right) - (h_i \zeta_{jk}^{old})^{-1} h_i \zeta_{jk} \right. \\ &\quad \left. - \log(h_i \zeta_{jk}^{old}) + 1 \right) = \mathcal{L}_{\zeta_{jk}}^*. \end{aligned} \quad (21)$$

We compute the derivative for the new bound:

$$\begin{aligned} \frac{\partial \mathcal{L}_{\zeta_{jk}}^*}{\partial \zeta_{jk}} &= \sum_{i=1}^T \sum_{j=1}^M \left(\Psi(\lambda_{jk}) - \Psi(\sum_{l=1}^K \lambda_{jl}) |w_{jm} = i \right) \\ &\quad + (\Psi(a_{1j}) - \Psi(a_{1j} + a_{2j})) + \sum_{i=1}^T (\Psi(a_{2i}) - \Psi(a_{1i} + a_{2i})) \\ &\quad - 1 - \log \zeta_{jk} + \frac{1}{2} \mu_{jk}^i \rho_{jk} - (h_i \zeta_{jk}^{old})^{-1} h_i. \end{aligned} \quad (22)$$

Finally, we set the derivative to zero to get the fixed point update:

$$\begin{aligned} \zeta_{jk} &= \exp \left(\sum_{i=1}^T \sum_{j=1}^M \left(\Psi(\lambda_{jk}) - \Psi(\sum_{l=1}^K \lambda_{jl}) |w_{jm} = i \right) \right. \\ &\quad \left. + (\Psi(a_{1j}) - \Psi(a_{1j} + a_{2j})) + \sum_{i=1}^T (\Psi(a_{2i}) - \Psi(a_{1i} + a_{2i})) \right. \\ &\quad \left. - 1 + \frac{1}{2} \mu_{jk}^i \rho_{jk} - (h_i \zeta_{jk}^{old})^{-1} h_i \right) \\ &= \exp \left(\sum_{i=1}^T \sum_{j=1}^M \mathbb{E}[\log p(w_{jk} | \theta_i)] + \mathbb{E}[\log \pi_j] \right. \\ &\quad \left. + \frac{1}{2} \mu_{jk}^i \rho_{jk} - (h_i \zeta_{jk}^{old})^{-1} h_i \right). \end{aligned} \quad (23)$$

A.4. Update of μ

Let $\bar{\delta} = \sum_{j=1}^M \prod_{i=1}^T \left(\sum_{k=1}^K \rho_{jk} \zeta_{jk} \exp\left(\frac{1}{2} \mu_{jk}\right) \right)$.

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial \mu_{jk}} &= \sum_{i=1}^T \left(\mu_{jk}^i \left(\frac{1}{2} \sum_{i=1}^T \rho_{jk} \zeta_{jk} \right) \right. \\ &\quad \left. - \bar{\delta}^{-1} \prod_{m=1}^T \left(\sum_{k=1}^K \rho_{jm} \zeta_{jm} \exp\left(\frac{1}{2} \mu_{jm}\right) \right) \right. \\ &\quad \left. \cdot \frac{\sum_{i=1}^T \left(\sum_{m=1}^T \rho_{jm} \zeta_{jm} \right) \cdot \frac{1}{2} \exp\left(\frac{1}{2} \mu_{jk}\right)}{\sum_{i=1}^T \left(\sum_{k=1}^K \rho_{jk} \zeta_{jk} \exp\left(\frac{1}{2} \mu_{jk}\right) \right)} \right) \end{aligned} \quad (24)$$

To incorporate the constraint that $\sum_{k=1}^K \zeta_{jk} = 1$, we use here instead of \dots , since the normalizing factor is dropped in the above result.

Reparameterization Trick

Express $q_\phi(\theta)$ as $\epsilon \sim r(\epsilon)$, $\theta = g(\epsilon, \phi)$

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$$\text{ELBO } L = E_{q_\phi(\theta)} \left[\log \frac{p(\theta, D)}{q_\phi(\theta)} \right] = E_{r(\epsilon)} \left[\log \frac{p(g(\epsilon, \phi), D)}{q_\phi(g(\epsilon, \phi))} \right]$$

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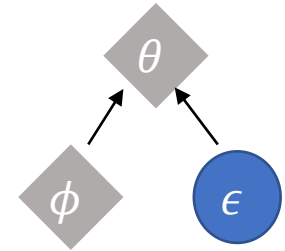
$$\nabla_\phi \hat{L} = \frac{1}{K} \sum_{k=1}^K \nabla_\phi \log \frac{p(g(\epsilon_k, \phi), D)}{q_\phi(g(\epsilon_k, \phi))}, \epsilon_k \sim r(\epsilon)$$



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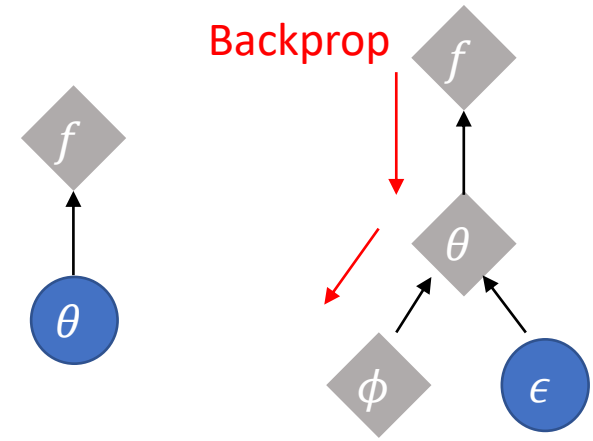
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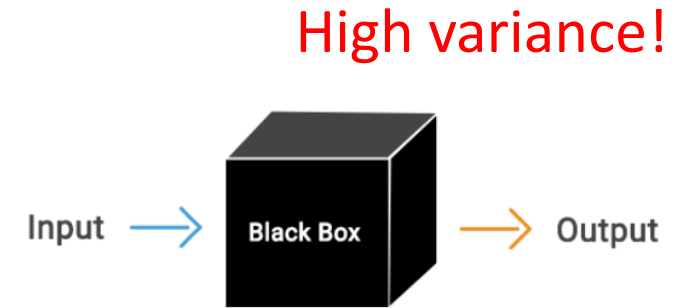
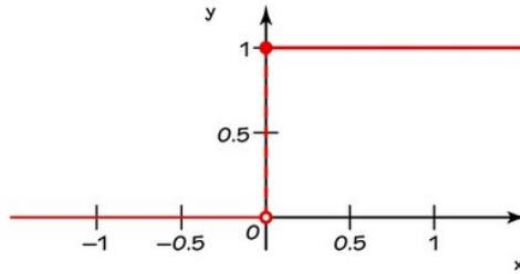
$$\text{Gradient } \nabla_\phi L = \nabla_\phi E_{r(\epsilon)} \left[\log \frac{p(g(\epsilon, \phi), D)}{q_\phi(g(\epsilon, \phi))} \right] = E_{r(\epsilon)} \left[\nabla_\phi \log \frac{p(g(\epsilon, \phi), D)}{q_\phi(g(\epsilon, \phi))} \right]$$

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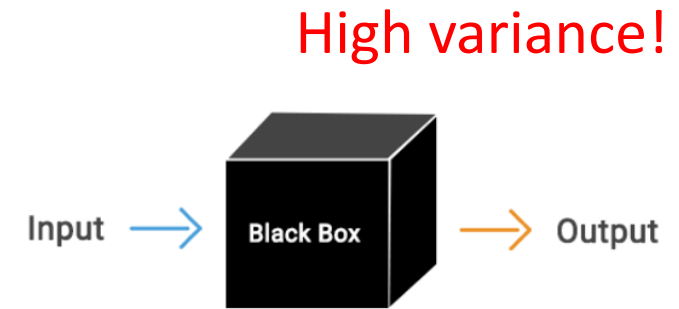
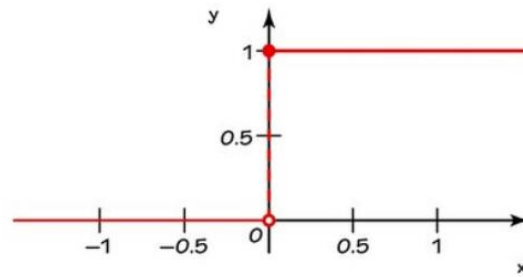
Variance Reduction Techniques in MCVI

- When non-differentiable, falls back to REINFORCE gradient



Variance Reduction Techniques in MCVI

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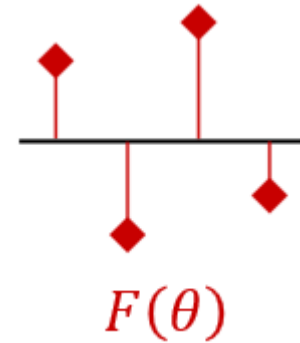


- Solutions to high variance REINFORCE gradients:
 - Low variance unbiased estimators with control variates
 - Biased estimators to enable reparam. trick (potentially low variance)

Variance Reduction Techniques in MCVI

- Control variate method:
 - Assume we want to estimate with MC simulation

$$E_{q(\theta)}[F(\theta)] \approx \frac{1}{K} \sum_k^K F(\theta_k), \quad \theta_k \sim q(\theta)$$

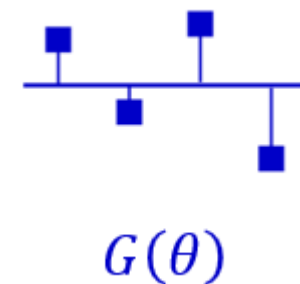
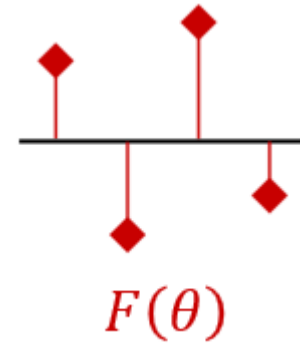


Variance Reduction Techniques in MCVI

- Control variate method:
 - Assume we want to estimate with MC simulation

$$E_{q(\theta)}[F(\theta)] \approx \frac{1}{K} \sum_k^K F(\theta_k), \quad \theta_k \sim q(\theta)$$

- Control variate: define a control function $G(\theta)$ satisfying:
 - $V_{q(\theta)}[G(\theta)] < \infty$
 - Known or fast computable $E_{q(\theta)}[G(\theta)]$



Variance Reduction Techniques in MCVI

- Control variate method:
 - Then define the new MC estimator

$$E_{q(\theta)}[F(\theta)] \approx \frac{1}{K} \sum_k^K \hat{F}(\theta_k), \quad \theta_k \sim q(\theta),$$

$$\hat{F}(\theta) = F(\theta) - G(\theta) + E_{q(\theta)}[G(\theta)]$$

$$V_{q(\theta)}[\hat{F}(\theta)] = V_{q(\theta)}[F(\theta)] + V_{q(\theta)}[G(\theta)] - 2 \text{Cov}_{q(\theta)}[F(\theta), G(\theta)]$$

< 0 if F and G are strongly and positively correlated

Variance Reduction Techniques in MCVI

- Application to REINFORCE gradient:

- $F(\theta) = \log \frac{p(D, \theta)}{q_\phi(\theta)} \nabla_\phi \log q_\phi(\theta)$
 $\quad \quad \quad \underline{\quad \quad \quad} := f(\theta)$

- Define $G(\theta) = g(\theta) \nabla_\phi \log q_\phi(\theta)$

- Variance reduced gradient: $\hat{F}(\theta) = \underbrace{(f(\theta) - g(\theta)) \nabla_\phi \log q_\phi(\theta)}_{:= \Delta(\theta)} + E_{q_\phi(\theta)}[G(\theta)]$

Variance Reduction Techniques in MCVI

- Application to REINFORCE gradient:

- $F(\theta) = \log \frac{p(D, \theta)}{q_\phi(\theta)} \nabla_\phi \log q_\phi(\theta)$
 $\log \frac{p(D, \theta)}{q_\phi(\theta)}$ $:= f(\theta)$

- Define $G(\theta) = g(\theta) \nabla_\phi \log q_\phi(\theta)$

- Variance reduced gradient: $\hat{F}(\theta) = \underbrace{(f(\theta) - g(\theta))}_{:= \Delta(\theta)} \nabla_\phi \log q_\phi(\theta) + E_{q_\phi(\theta)}[G(\theta)]$

- “Baseline” approach:

$$g(\theta) = b$$

$$\Rightarrow E_{q(\theta)}[G(\theta)] = b E_{q(\theta)}[\nabla_\phi \log q(\theta)] = b \nabla_\phi \int q_\phi(\theta) d\theta = b \nabla_\phi 1 = 0$$

(log-derivative trick)

$$\Rightarrow \hat{F}(\theta) = \Delta(\theta) \nabla_\phi \log q_\phi(\theta)$$

Variance Reduction Techniques in MCVI

- Application to REINFORCE gradient:

- $$F(\theta) = \log \frac{p(D, \theta)}{q_\phi(\theta)} \nabla_\phi \log q_\phi(\theta)$$

$\underline{\hspace{10em}} := f(\theta)$

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 $\hspace{15em} := \Delta(\theta)$

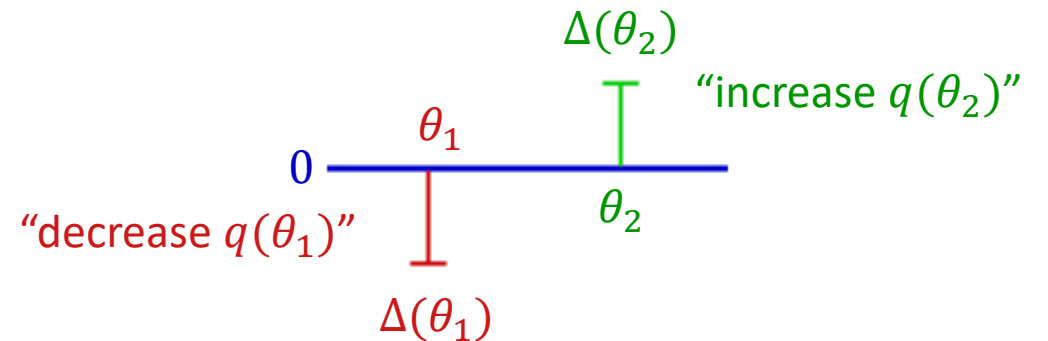
- “Baseline” approach:

$$g(\theta) = b$$

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log-derivative trick

$$\Rightarrow \hat{F}(\theta) = \Delta(\theta) \nabla_\phi \log q_\phi(\theta)$$



b fitted by minimising either $V_{q(\theta)}[\hat{F}(\theta)]$ or $E_{q(\theta)}[\Delta(\theta)^2]$

Variance Reduction Techniques in MCVI

- Application to REINFORCE gradient:

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- “Taylor expansion” approach (e.g. 1st order):

$$g(\theta) = f(\theta_0) + \nabla_{\theta_0} f(\theta_0)(\theta - \theta_0)$$

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$:= \Delta(\theta)$

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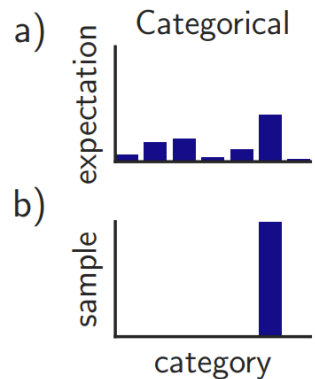
$$\Rightarrow E_{q(\theta)}[G(\theta)] = \underbrace{(f(\theta_0) - \nabla_{\theta_0} f(\theta_0) \theta_0)}_{= 0 \text{ (log-derivative trick)}} E_{q(\theta)}[\nabla_\phi \log q(\theta)]$$

$$\begin{aligned} &+ \nabla_{\theta_0} f(\theta_0) E_{q(\theta)}[\theta \nabla_\phi \log q(\theta)] \\ &\quad = \nabla_{\theta_0} f(\theta_0) \nabla_\phi E_{q(\theta)}[\theta] \text{ (log-derivative trick)} \end{aligned}$$

$$\Rightarrow \hat{F}(\theta) = \Delta(\theta) \nabla_\phi \log q_\phi(\theta) + \underbrace{\nabla_{\theta_0} f(\theta_0) \nabla_\phi E_{q(\theta)}[\theta]}_{:= \nabla_\phi f(\theta_0) \text{ if } \theta_0 = E_{q(\theta)}[\theta]}$$

Variance Reduction Techniques in MCVI

- Gumbel-Softmax trick
 - Biased gradient estimator
 - Empirically found to have smaller variance

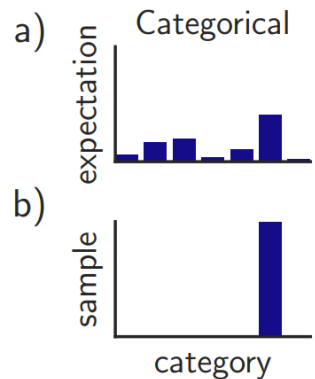


Categorical distribution:

$$p(y = k) = \pi_k, \sum_k \pi_k = 1$$

Variance Reduction Techniques in MCVI

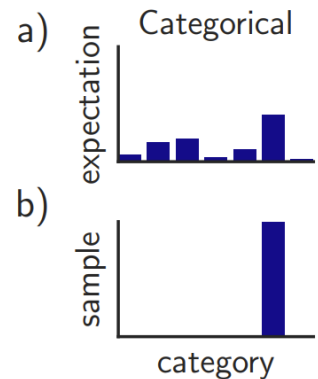
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Gumbel trick to sample y :
 $y = \arg \max [g_k + \log \pi_k],$
 $g_k \sim \text{Gumbel}(0, 1)$

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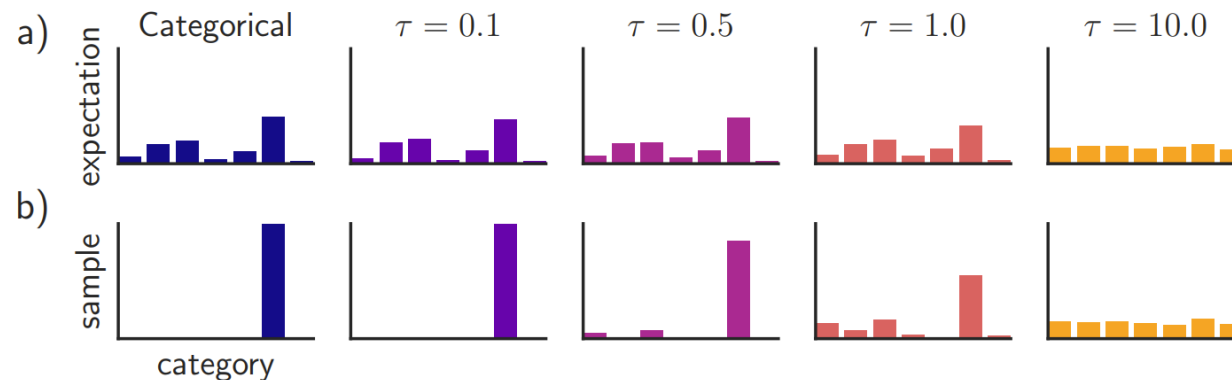
Gumbel trick to sample y :
 $y = \arg \max [g_k + \log \pi_k],$
 $g_k \sim \text{Gumbel}(0, 1)$

replace with softmax



Variance Reduction Techniques in MCVI

- Gumbel-Softmax trick
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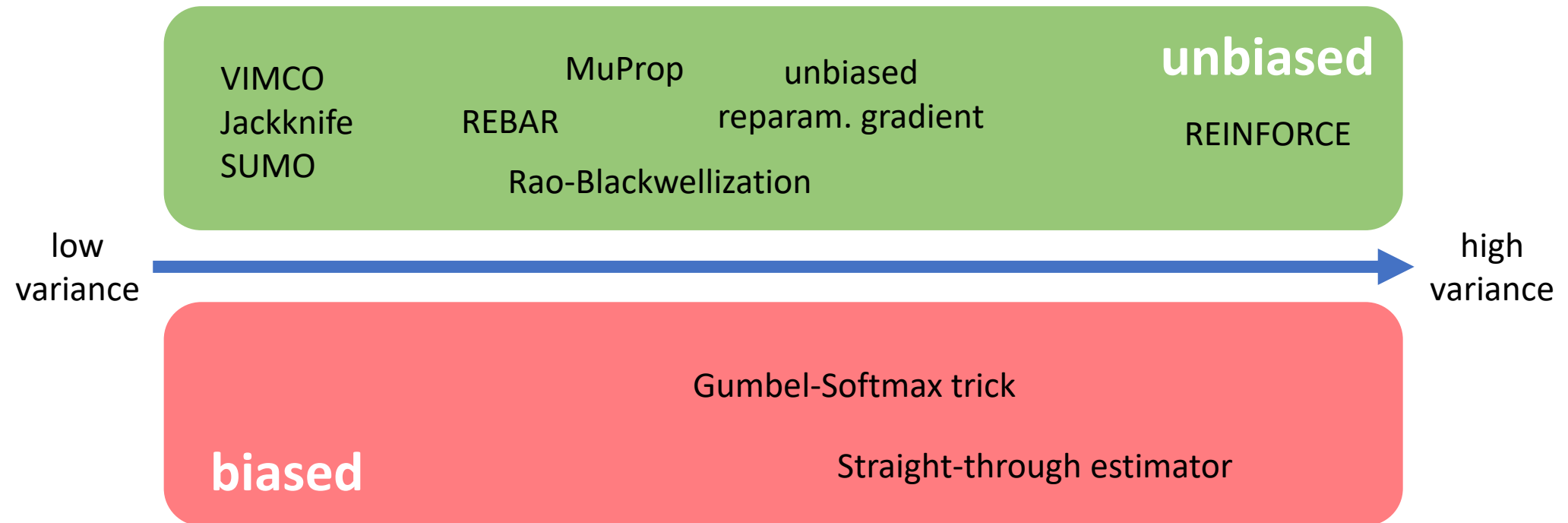


Concrete distribution/Gumbel-Softmax trick:

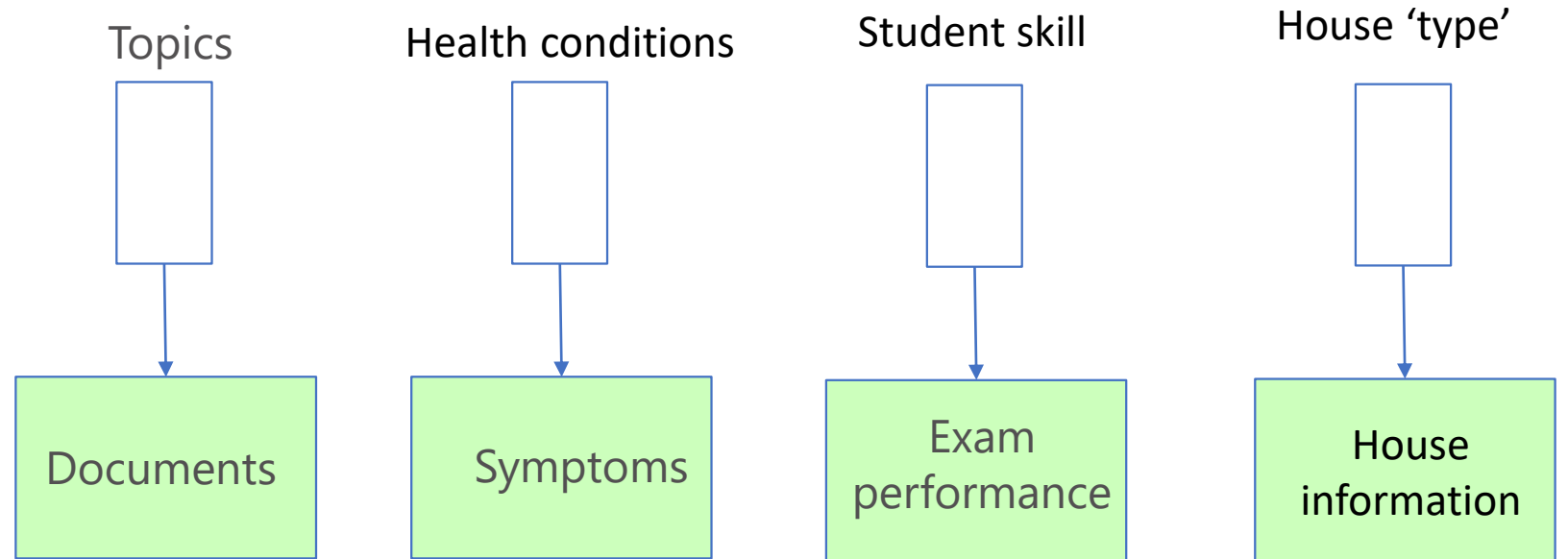
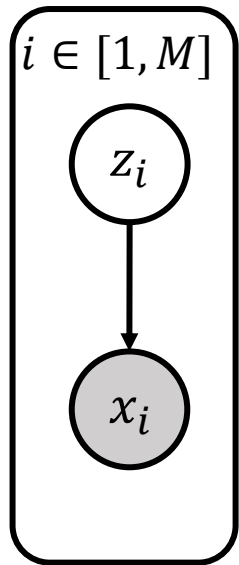
sample the “soft vector” (instead of one-hot encoding of y)

$$[y_1, \dots, y_K] = \textit{softmax}\left(\left[\frac{(g_1 + \log \pi_1)}{\tau}, \dots, \frac{(g_K + \log \pi_K)}{\tau}\right]\right)$$

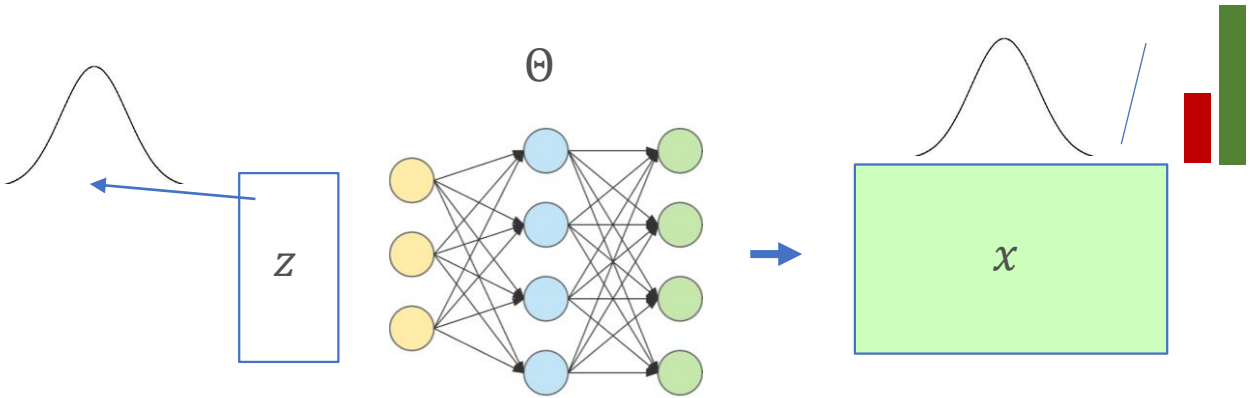
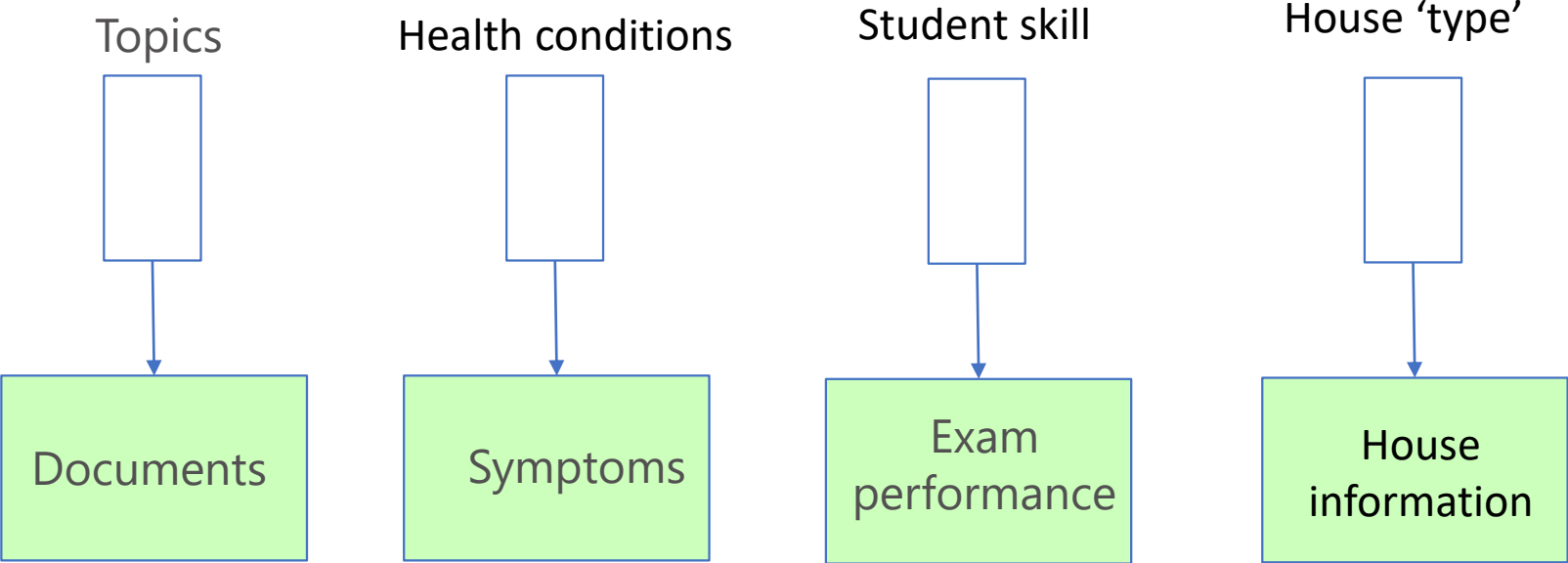
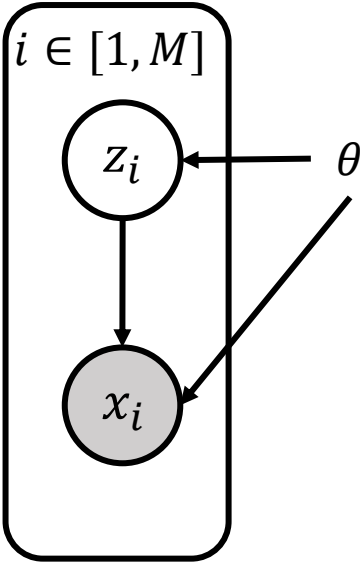
Variance Reduction Techniques in MCVI



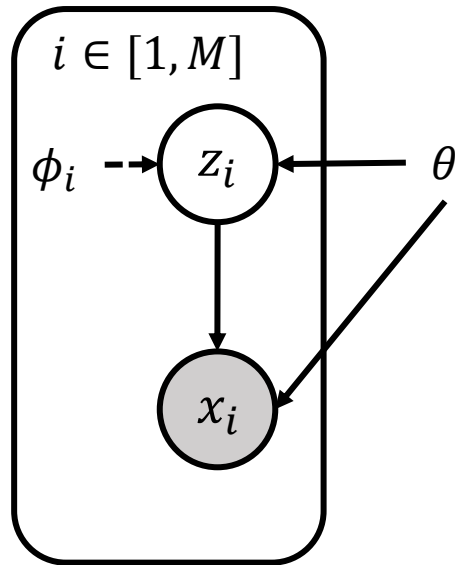
Latent Variable Model



Deep Latent Variable Model

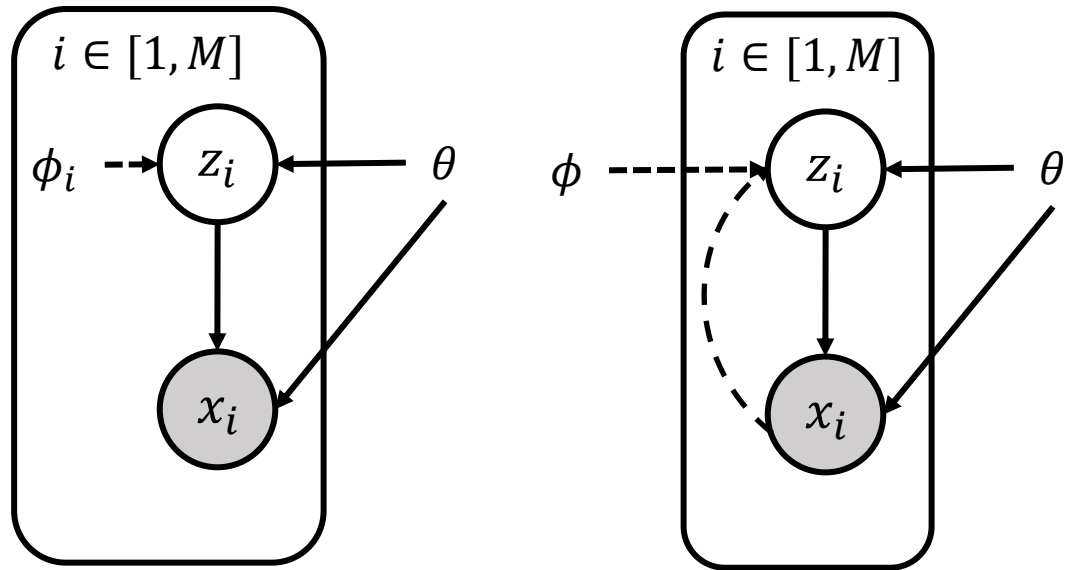


Amortized Inference



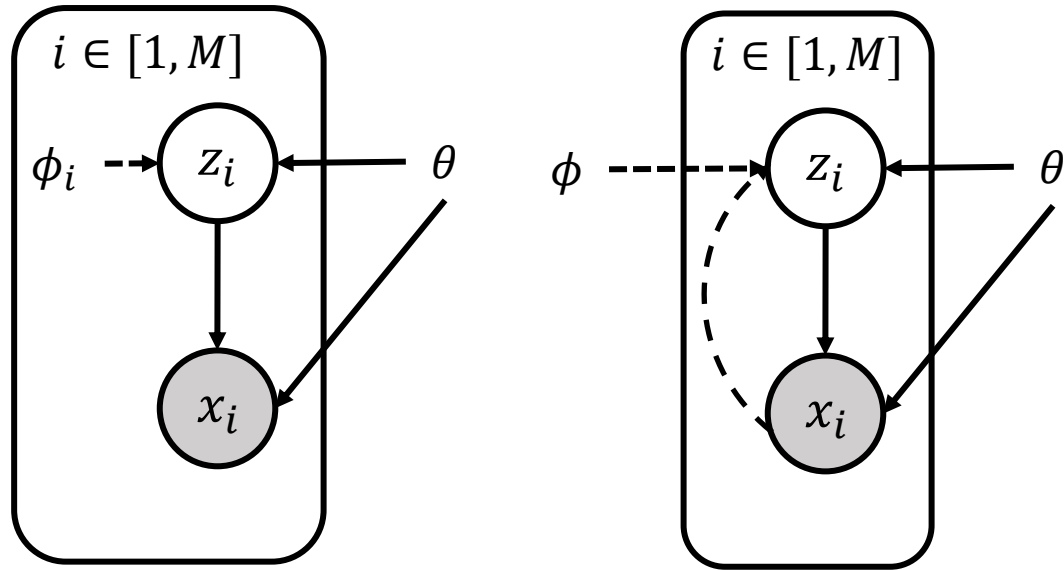
ϕ parameter for variational distribution
 θ decoder parameter

Amortized Inference



ϕ parameter for variational distribution
 θ decoder parameter

Amortized Inference

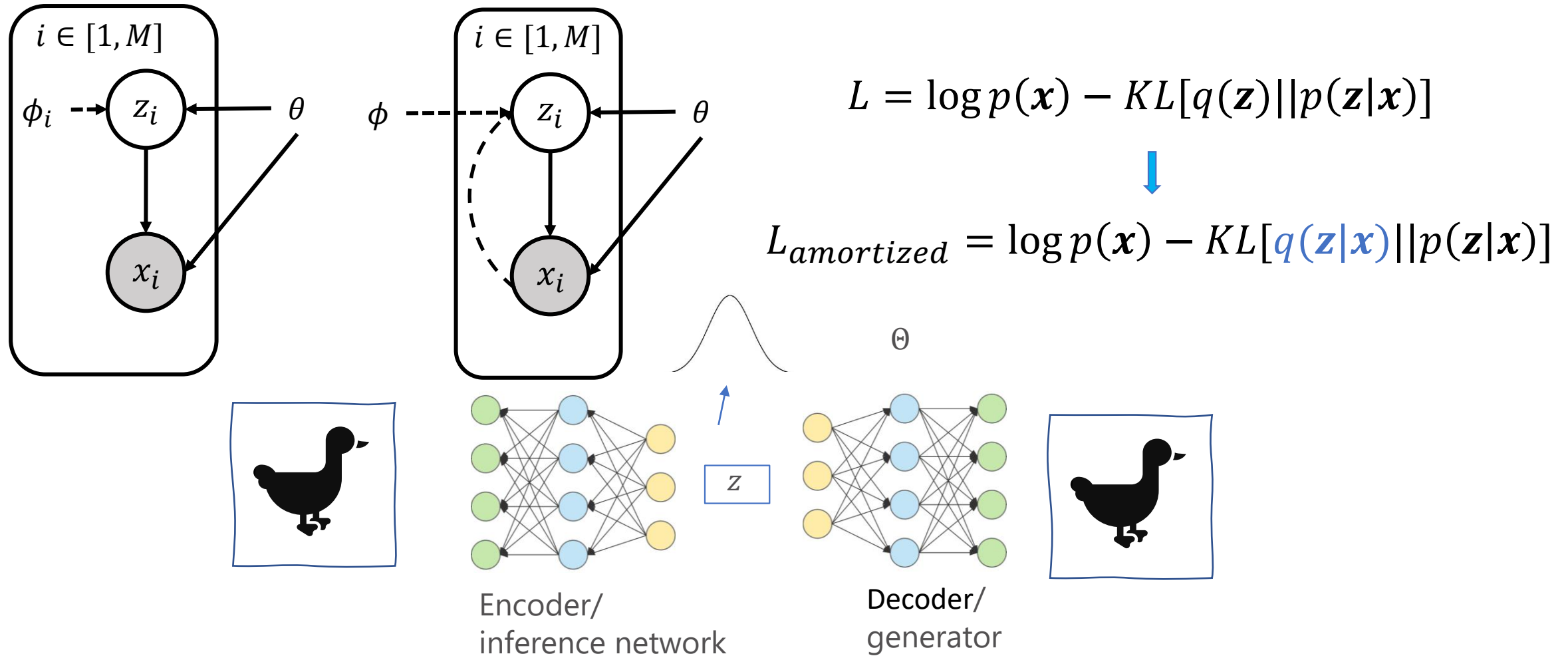


$$L = \log p(\mathbf{x}) - KL[q(\mathbf{z}) || p(\mathbf{z}|\mathbf{x})]$$

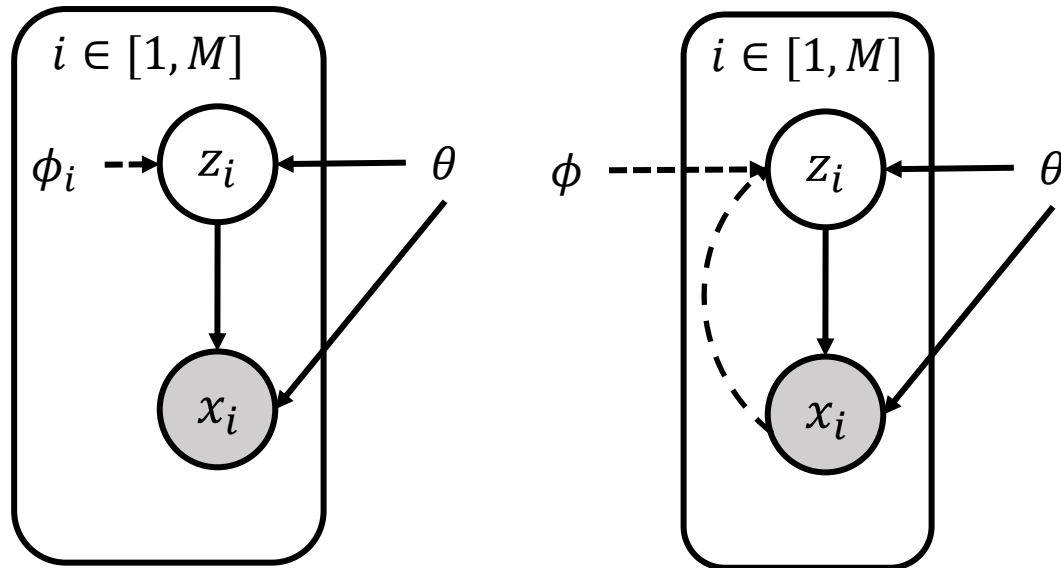


$$L_{amortized} = \log p(\mathbf{x}) - KL[q(\mathbf{z}|\mathbf{x}) || p(\mathbf{z}|\mathbf{x})]$$

Variational Auto-Encoders (VAE)



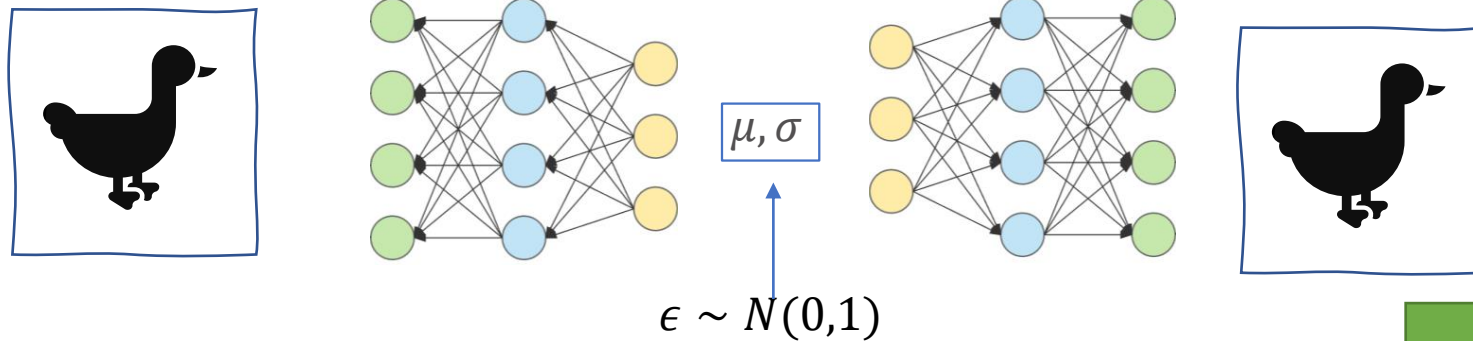
Variational Auto-Encoders (VAE)



$$L = \log p(\mathbf{x}) - KL[q(\mathbf{z})||p(\mathbf{z}|\mathbf{x})]$$



$$L_{amortized} = \log p(\mathbf{x}) - KL[q(\mathbf{z}|\mathbf{x})||p(\mathbf{z}|\mathbf{x})]$$



$$z \sim N(\mu, \sigma^2)$$

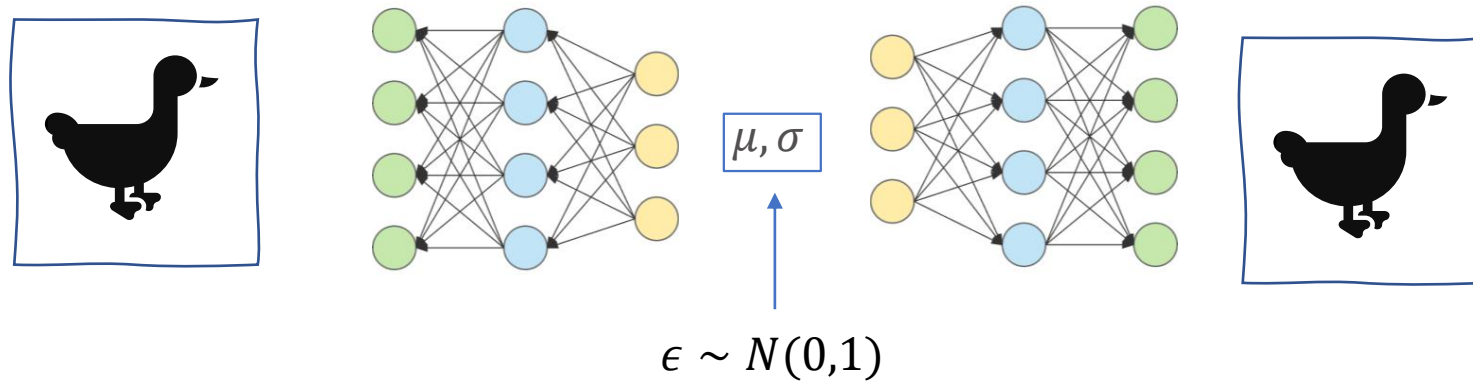
$$\epsilon \sim N(0, 1), \theta = \mu + \sigma\epsilon$$

Kingma and Welling. Auto-encoding variational bayes. ICLR 2014.

Rezende et al. Stochastic backpropagation and approximate inference in deep generative models. ICML 2014.



Variational Auto-Encoders (VAE)

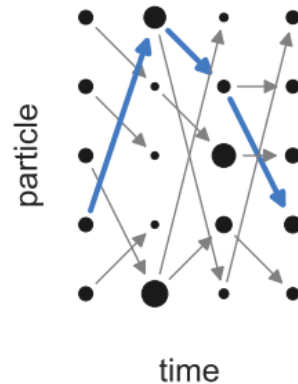
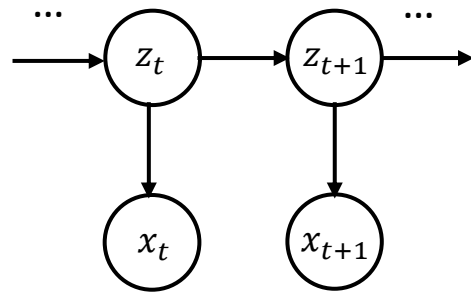


$$\begin{aligned} L_{amortized} &= \log p(\mathbf{x}) - KL[q(\mathbf{z}|\mathbf{x})||p(\mathbf{z}|\mathbf{x})] \\ &= E_{\mathbf{z} \sim q(\mathbf{z}|\mathbf{x})}[\log p_{\theta}(\mathbf{x}|\mathbf{z})] - KL[q(\mathbf{z}|\mathbf{x})||p(\mathbf{z}|\mathbf{x})] \end{aligned}$$

How to apply amortization to other inference methods?

Amortized Inference: Further Examples

- Amortized SMC



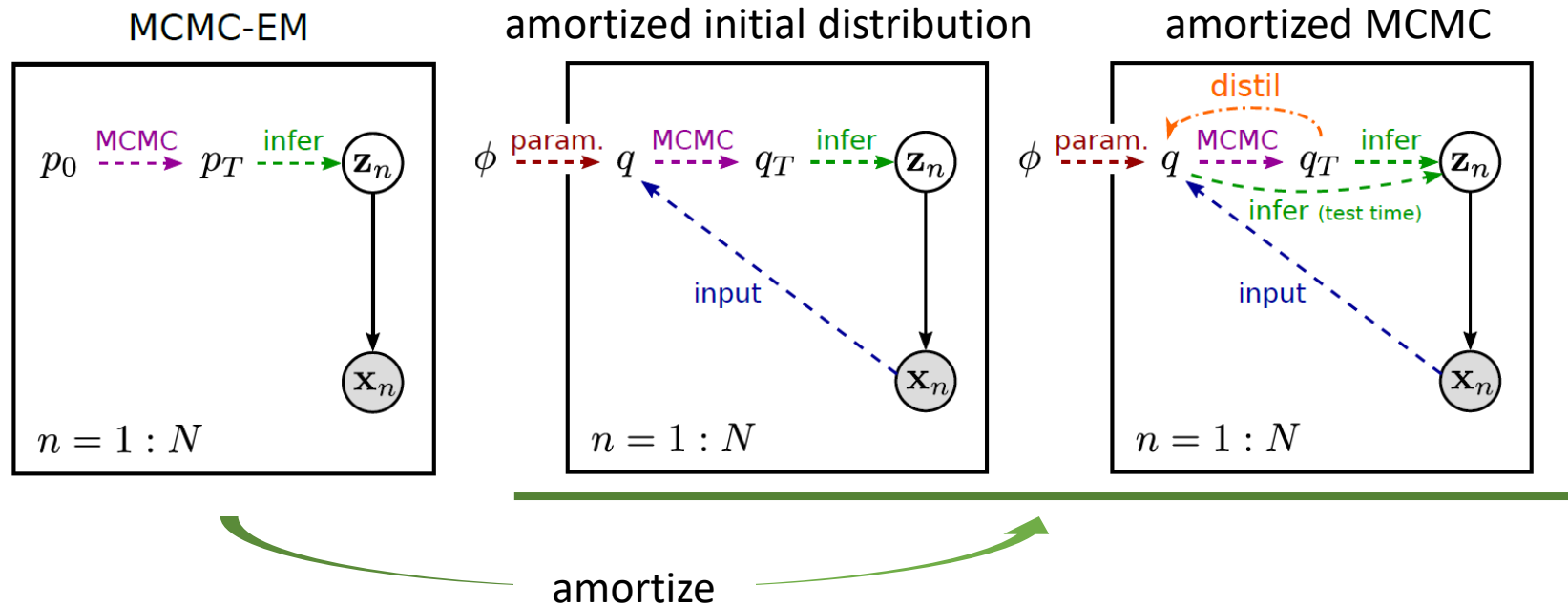
amortize

Find the optimal proposal distribution for each sequence $\{x_{1:T}\}$ & each time step

Explicitly parameterise & optimise $(x_{1:T}, z_{1:t}) \rightarrow$ proposal dist. for z_t

Amortized Inference: Further Examples

- Amortized MCMC

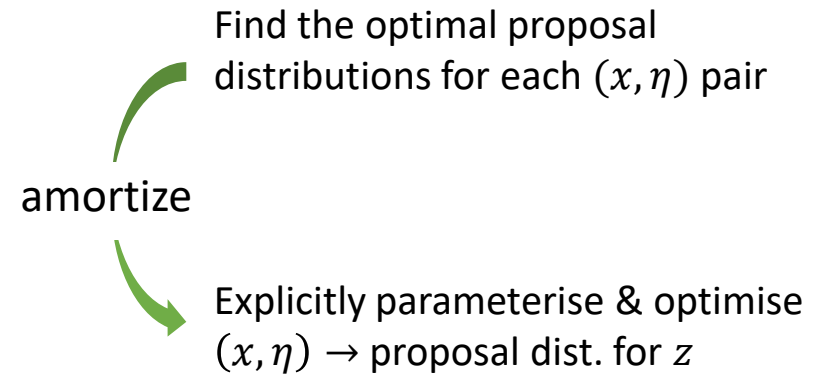


Amortized Inference: Further Examples

- Amortized Monte Carlo integration

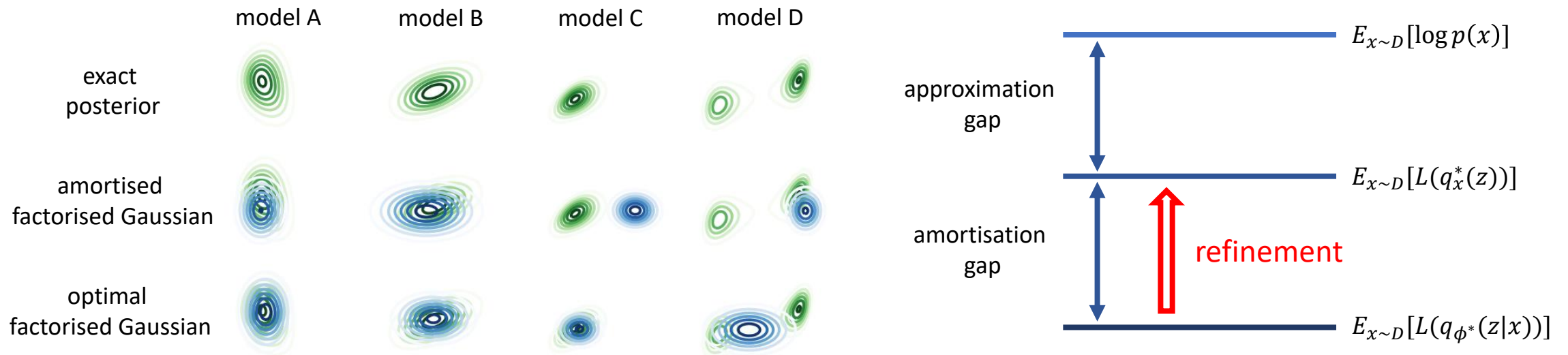
Goal: estimate with
importance sampling

$$E_{p(z|x)}[F_{\eta}(z)]$$



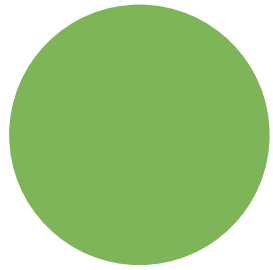
Amortized Inference: Limitations

- Amortised approximate posteriors in practice are sub-optimal



- The “**refinement**” idea:

- Initialise $q(z|x) = N(z; \mu, \sigma^2)$ with the amortised solution $\mu \leftarrow \mu_{\phi}(x), \sigma \leftarrow \sigma_{\phi}(x)$
- Then run T more VI gradient steps to update μ, σ

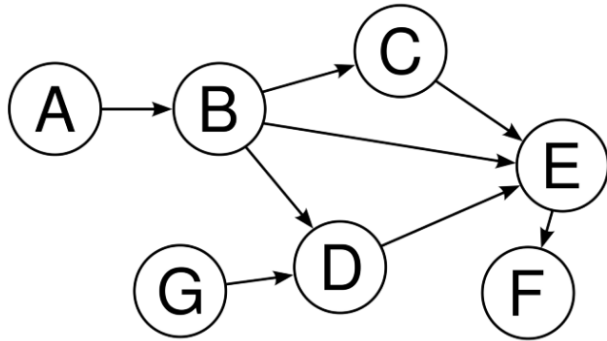


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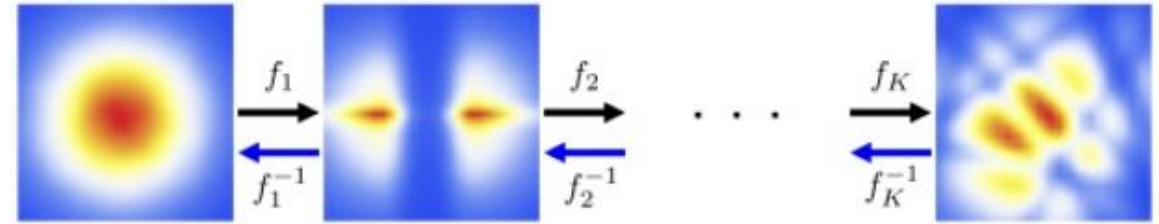
Part II: Advances

- Scalable variational inference
- Monte Carlo methods
- Amortized inference
- **Approximate distribution design**
- **Optimization objective design**

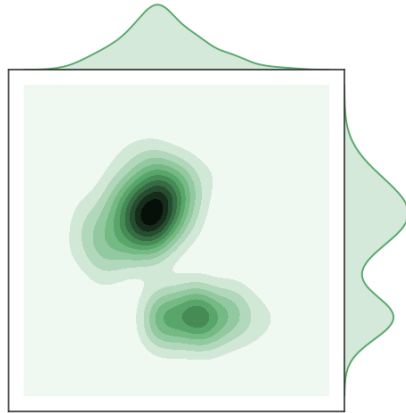
Designing q Distributions



Structured approximations



Normalizing flows



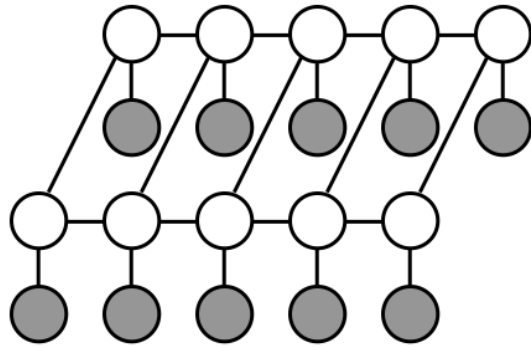
Auxiliary variables & mixture distributions



Implicit approximate posteriors

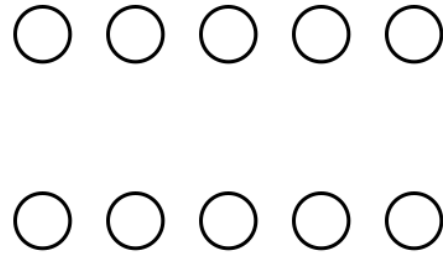
Structured Approximations

- introduce dependencies between random variables for q :



Hidden Markov Model

Exact posterior $p(z | x)$
 $z_i \not\perp z_j | x$

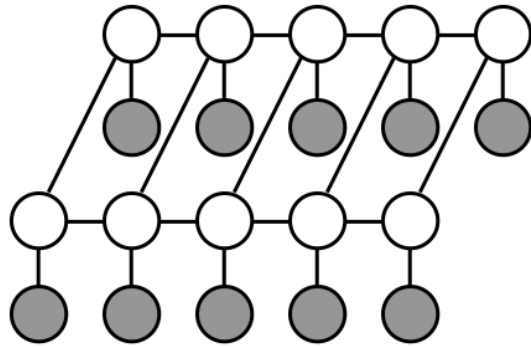


Mean-field approximation

$$q(z) = \prod_i q(z_i)$$

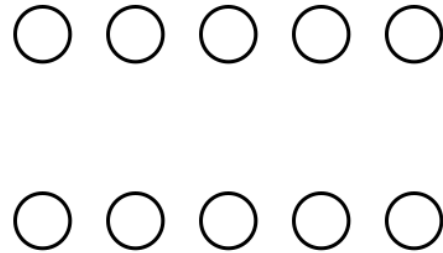
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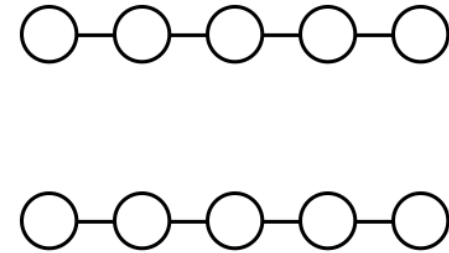
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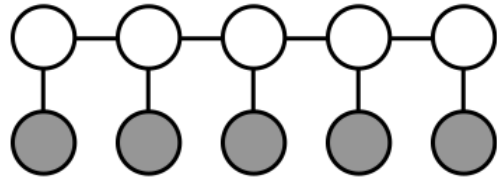
Structured approximation

$$q(z) = \prod_s q(z_s)$$
$$q(z_s) = q(\{z_i\}_{i \in s})$$

Main design question:
the grouping and
conditional dependency
structure

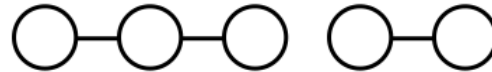
Structured Approximations

- Auto-regressive distributions (as a specific dependency structure)



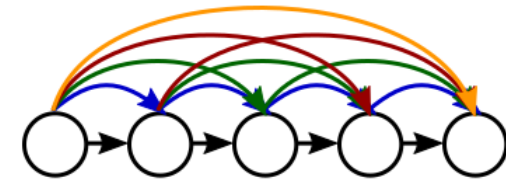
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$q(z) = \prod_s q(z_s)$
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Auto-regressive approximation

$q(z) = \prod_i q(z_i | z_{<i})$
 $q(z_1 | z_{<1}) = q(z_1)$

Main design question:
the ordering of the
latent variables

Normalizing Flows

- Change-of-variable formula:
 - x is a random variable with probability density function (PDF) $p_X(x)$
 - $y = f(x)$ is an invertible mapping

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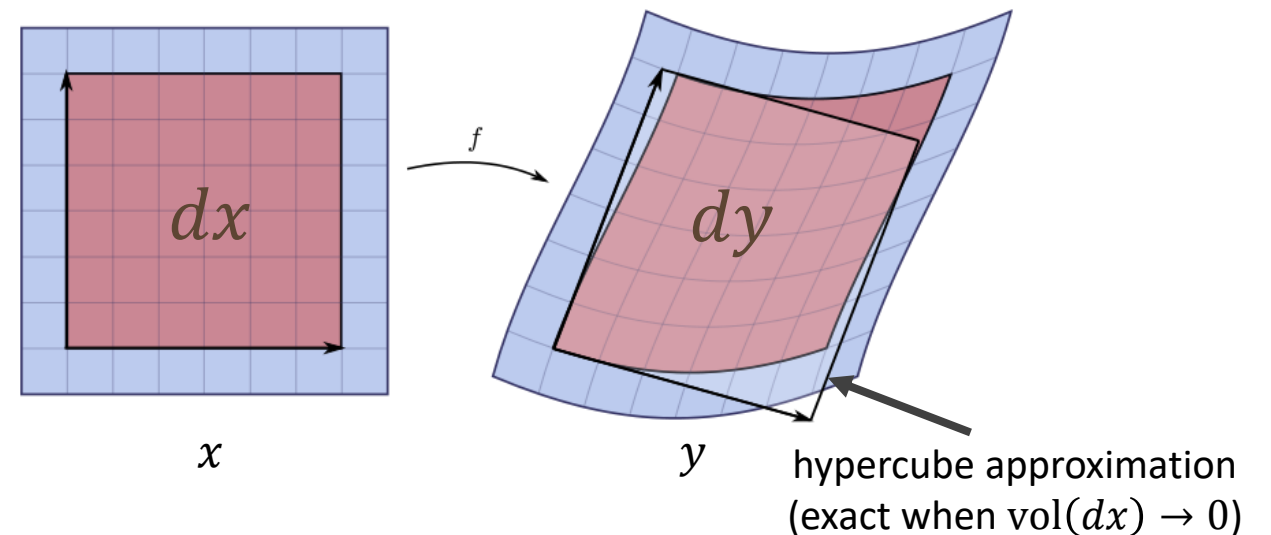
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prob. mass of region around y

prob. mass of region around x

$$p_Y(y) = p_X(x) \left| \det \left(\frac{dx}{dy} \right) \right|$$
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Normalizing Flows

- Variational inference with Normalizing flow
 - Assume $q_0(z_0) = N(z_0; 0, I)$
 - Define $z = f_\phi(z_0)$ where $f_\phi(\cdot)$ is an invertible mapping parameterized by ϕ

$$q(z) = q_0(z_0) \left| \det \left(\frac{dz}{dz_0} \right) \right|^{-1} \quad \text{with } z_0 = f_\phi^{-1}(z)$$

(change of variable: $q(z)dz = q_0(z_0)dz_0$)

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reparam. trick:

$$z \sim q(z) \Leftrightarrow z_0 \sim q_0(z_0), z = f_\phi(z_0)$$

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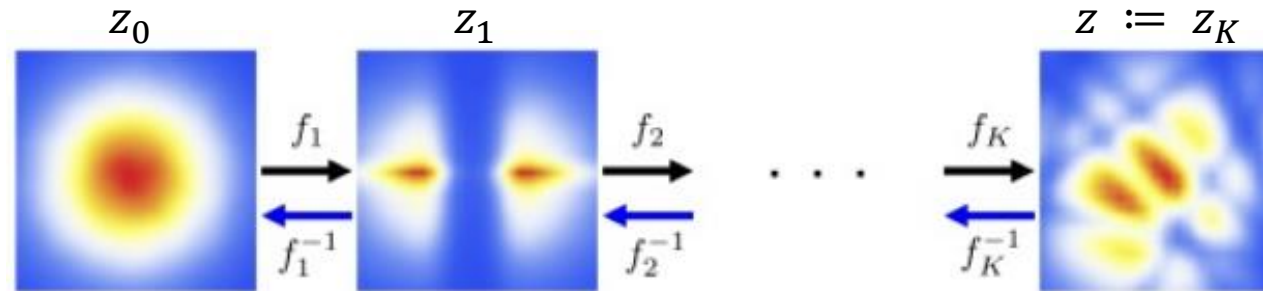
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- Computing ELBO requires $\log \left| \det \left(\frac{df_\phi}{dz_0} \right) \right|$

reparam. trick:
 $z \sim q(z) \Leftrightarrow z_0 \sim q_0(z_0), z = f_\phi(z_0)$

Normalizing Flows

- Variational inference with Normalizing flow
 - Idea: define f_ϕ such that $\log \left| \det \left(\frac{df_\phi}{dz_0} \right) \right|$ is easy to compute!
 - Chain simple invertible mappings together to make a flexible mapping



$$f_\phi = f_K \circ f_{K-1} \circ \dots \circ f_1, f_k(\cdot) := f_{\phi_k}(\cdot), \phi = \{\phi_k\}_{k=1}^K$$

- For each simple mapping, hopefully the Jacobian log-determinant is easy to compute

$$\Rightarrow \log \left| \det \left(\frac{df_\phi}{dz_0} \right) \right| = \sum_{k=1}^K \log \left| \det \left(\frac{dz_k}{dz_{k-1}} \right) \right|$$

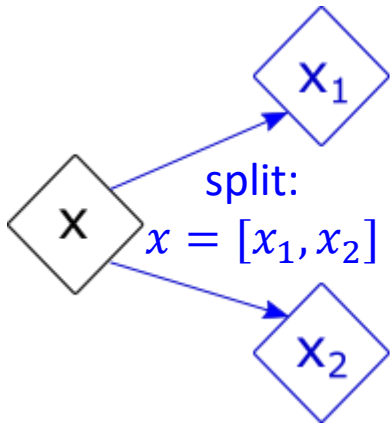
Normalizing Flows

- Goal: construct f_k to enable fast compute of $\log \left| \det \left(\frac{dz_k}{dz_{k-1}} \right) \right|$
 - Example (RealNVP): $y := f_{\phi_k}(x)$ computed as follows



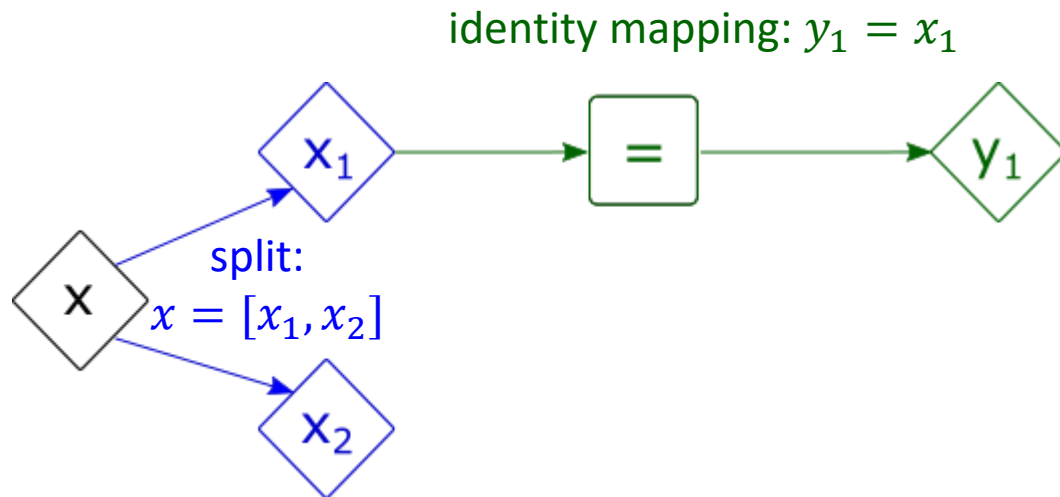
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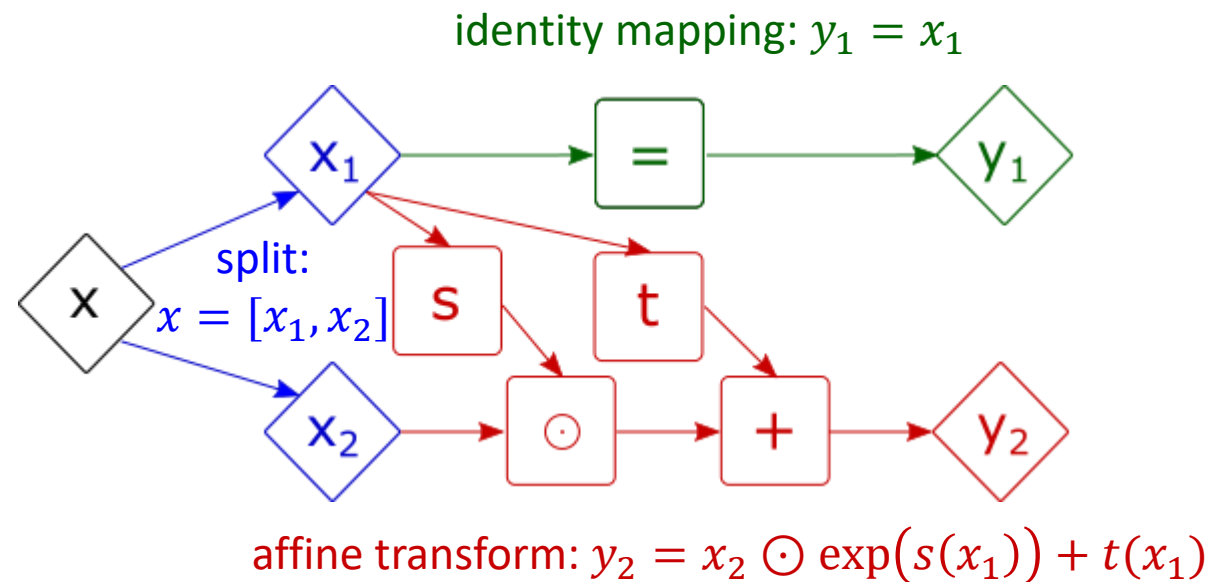
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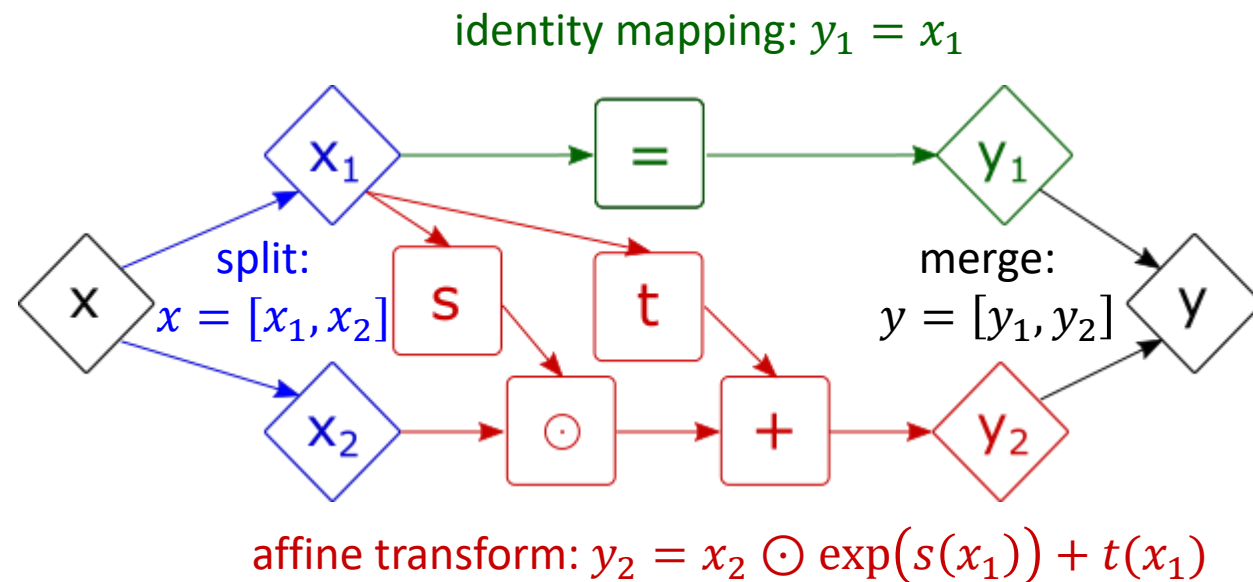
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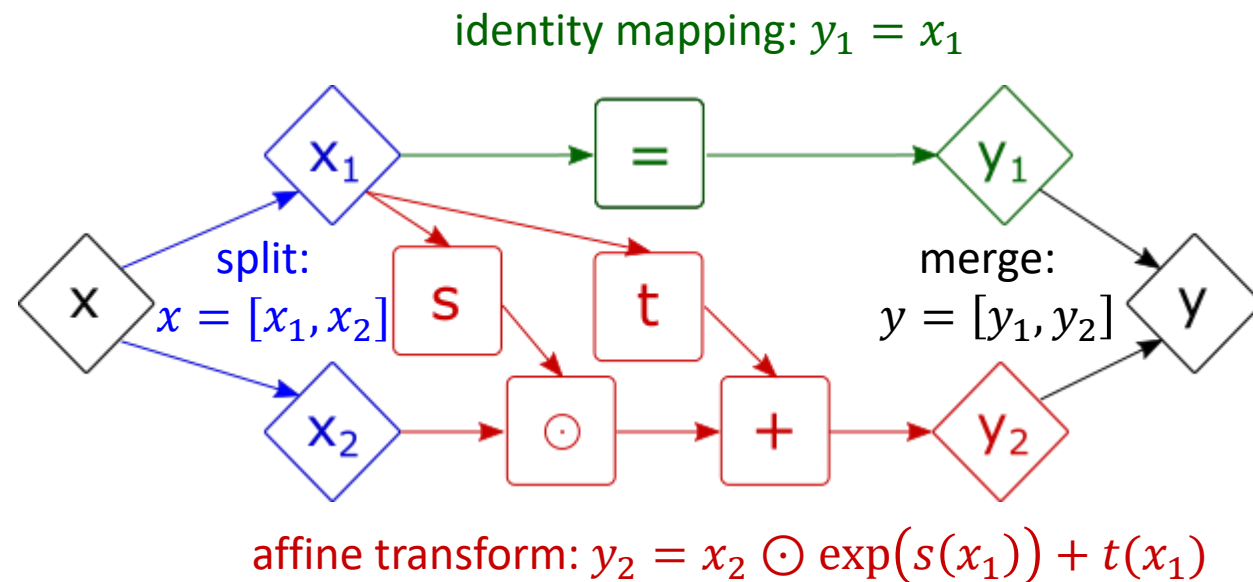
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Jacobian:

$$\frac{df_{\phi_k}}{dx} = \begin{pmatrix} I & 0 \\ dy_2/dx_1 & \text{diag}(\exp(s(x_1))) \end{pmatrix}$$

Log-determinant of Jacobian:

$$\Rightarrow \log \left| \det \left(\frac{df_{\phi}}{dx} \right) \right| = \sum_i s(x_1)_i$$

Auxiliary Variables & Mixture Distributions

- Construct $q(\theta)$ as a (hierarchical) mixture distribution

$$q(\theta) = \int q(\theta | a) q(a) da$$

- a is the auxiliary variable used to enrich the approximate posterior

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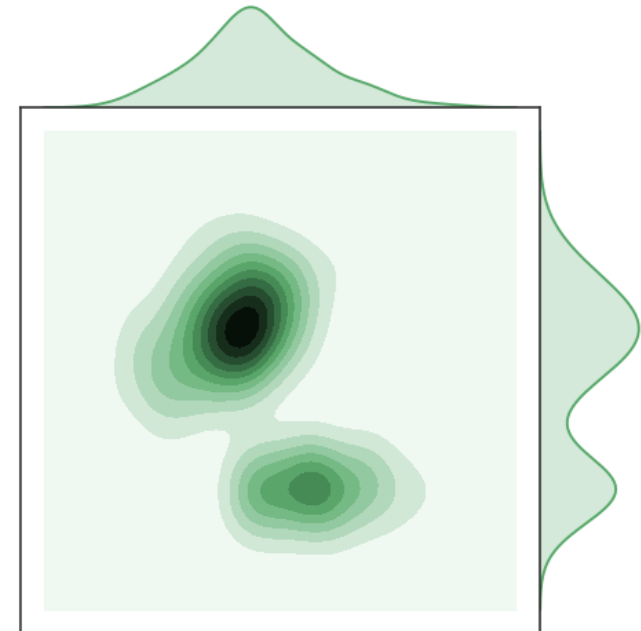
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- Example: Mixture of Gaussians

$$a \sim q(a) = \text{Categorical}(\pi_1, \dots, \pi_K)$$

$$\theta \sim q(\theta | a) = N(\theta; m_a, \Sigma_a)$$

Can be very flexible with many components!



Auxiliary Variables & Mixture Distributions

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- Now the variational lower-bound becomes intractable:

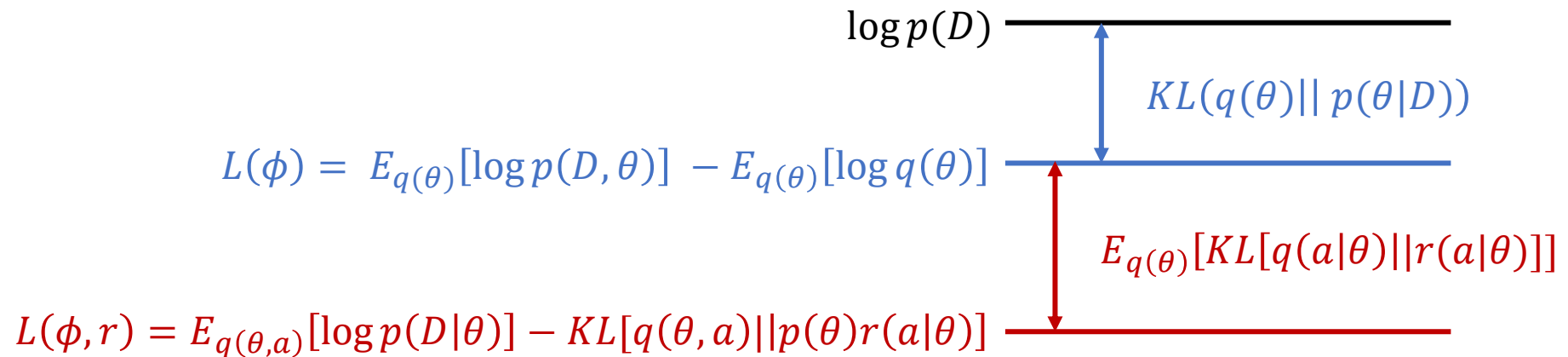
$$L(\phi) = \underbrace{E_{q(\theta)}[\log p(D, \theta)]}_{\text{Estimated by Monte Carlo:}} - \underbrace{E_{q(\theta)}[\log q(\theta)]}_{\text{Intractable density}}$$

Estimated by Monte Carlo:
 $a_k \sim q(a), \theta_k \sim q(\theta | a_k)$

Intractable density
 $q(\theta) = \int q(\theta|a)q(a) da$

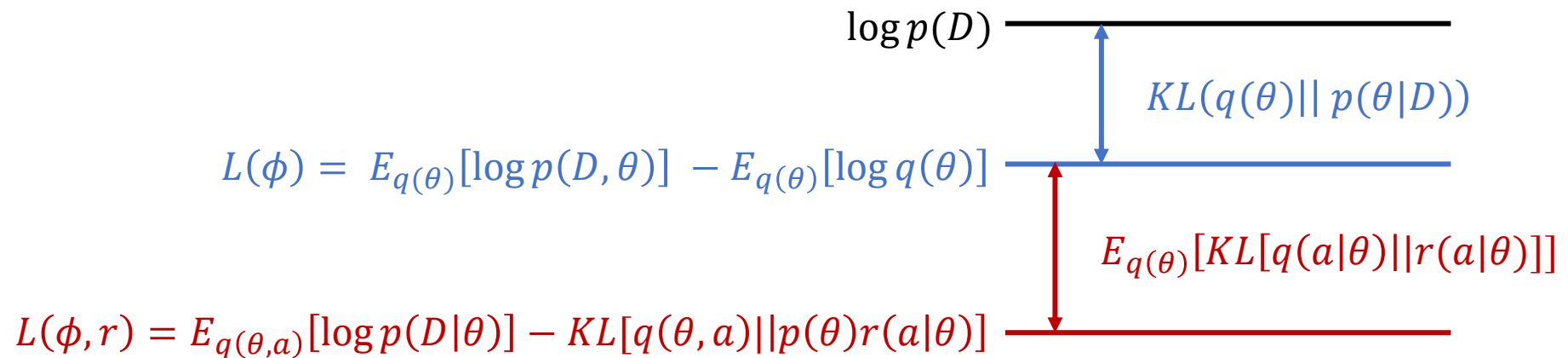
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- Solution: introducing an auxiliary variational lower-bound $L(\phi, r)$ with an auxiliary distribution $r(a|\theta)$:



Auxiliary Variables & Mixture Distributions

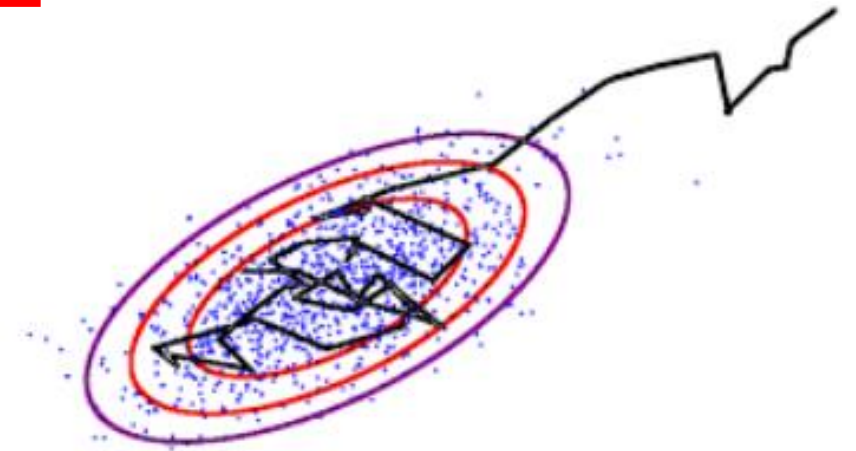
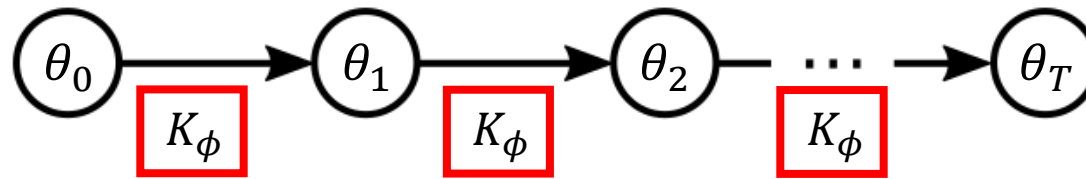
- Solution: introducing an auxiliary variational lower-bound $L(\phi, r)$ with an auxiliary distribution $r(a|\theta)$:



- Optimize $r(a|\theta)$ to close the gap!
- $L(\phi, r)$ estimated by Monte Carlo: $a_k \sim q(a), \theta_k \sim q(\theta | a_k)$

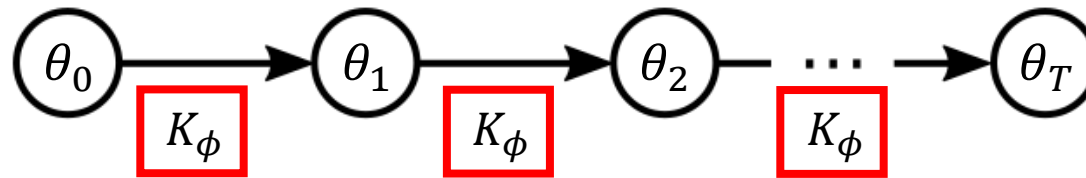
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Auxiliary Variables & Mixture Distributions

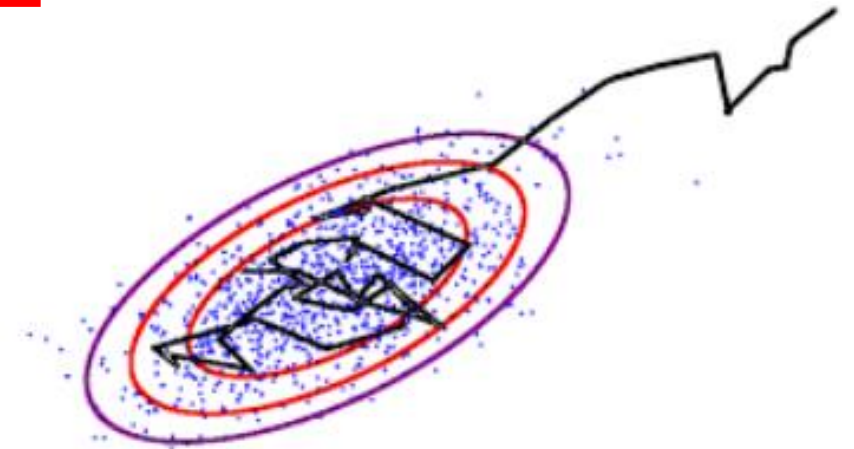
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learn the transition kernel with VI:

$$\theta := \theta^T, a = \{\theta^{0:T-1}\}$$

$$q(\theta^T) = \int q_0(\theta^0) \prod_{t=1}^T K_\phi(\theta^t | \theta^{t-1}) d\theta^{0:T-1}$$



Implicit Approximate Posteriors

- Two quantities computed in (approximate) Bayesian inference:

approximate Bayesian predictive

$$p(y^*|x^*, D) \approx E_{q(\theta)}[p(y^*|x^*, \theta)]$$

approximate posterior moments

$$E_{q(\theta)}[F(\theta)]$$

Implicit Approximate Posteriors

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$$p(y^*|x^*, D) \approx E_{q(\theta)}[p(y^*|x^*, \theta)]$$

$$\approx \frac{1}{K} \sum_k^K p(y^*|x^*, \theta_k), \theta_k \sim q(\theta)$$

approximate posterior moments

$$E_{q(\theta)}[F(\theta)]$$

$$\approx \frac{1}{K} \sum_k^K F(\theta_k), \theta_k \sim q(\theta)$$

Computed with Monte Carlo estimates

Only require fast sampling from q !
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implicit distributions

Mohamed and Lakshminarayanan. Learning in Implicit Generative Models. arXiv 2016

Li and Liu. Wild Variational Inference. AABI 2016

Huszár. Variational Inference using Implicit Distributions. arXiv 2017

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Mescheder et al. Adversarial Variational Bayes: Unifying Variational Autoencoders and Generative Adversarial Networks. ICML 2017

Tran et al. Hierarchical Implicit Models and Likelihood-Free Variational Inference. NeurIPS 2017

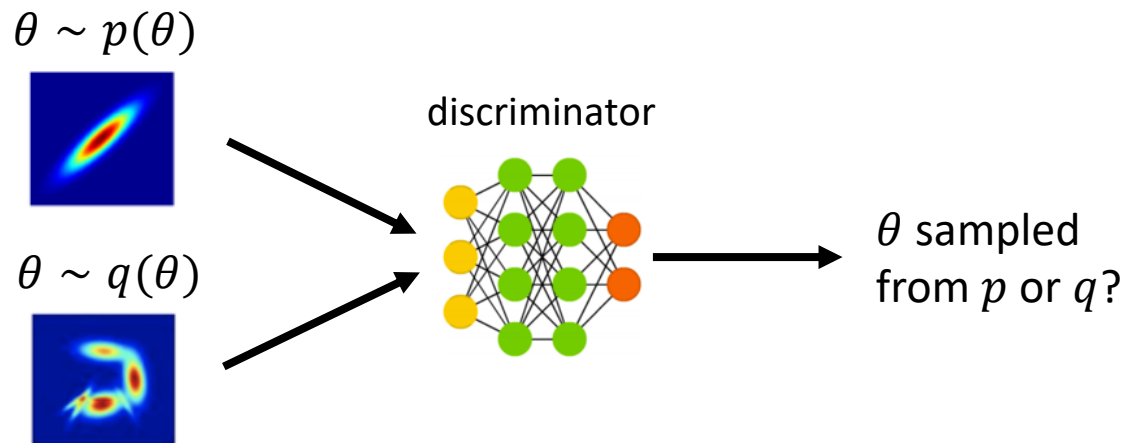
Li and Turner. Gradient Estimators for Implicit Models. ICLR 2018

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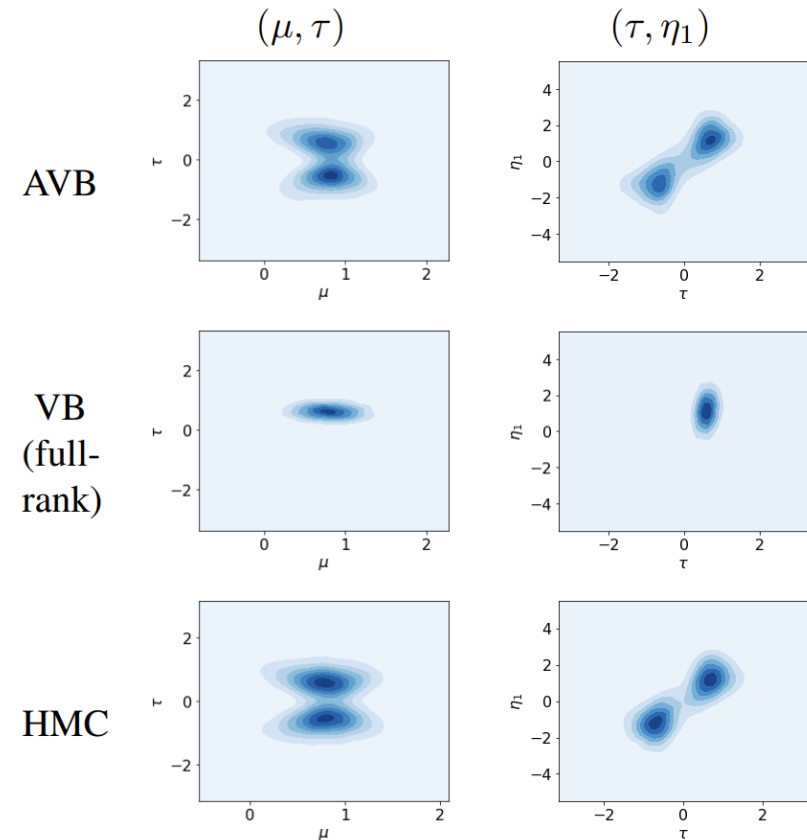
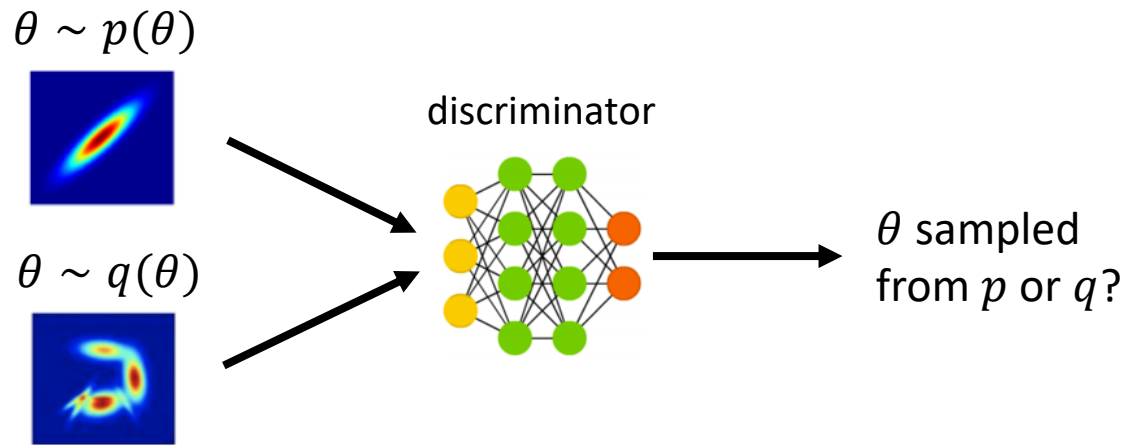
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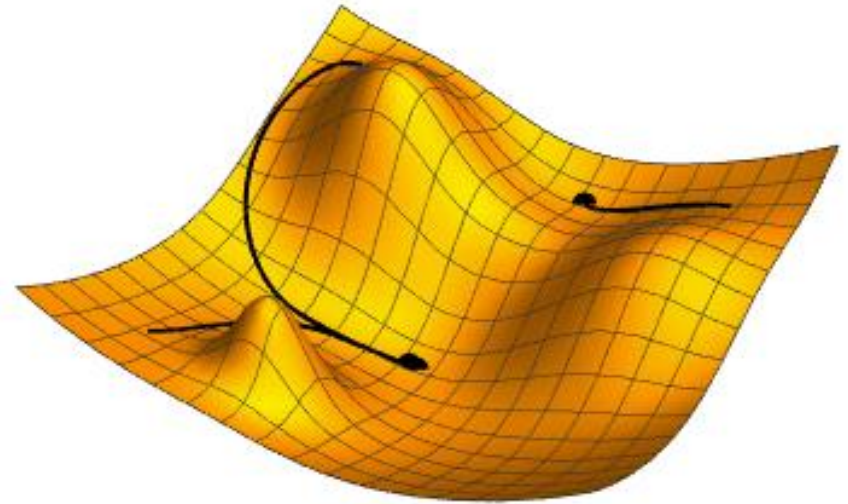
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Objective Functions

For fitting the approximate posterior



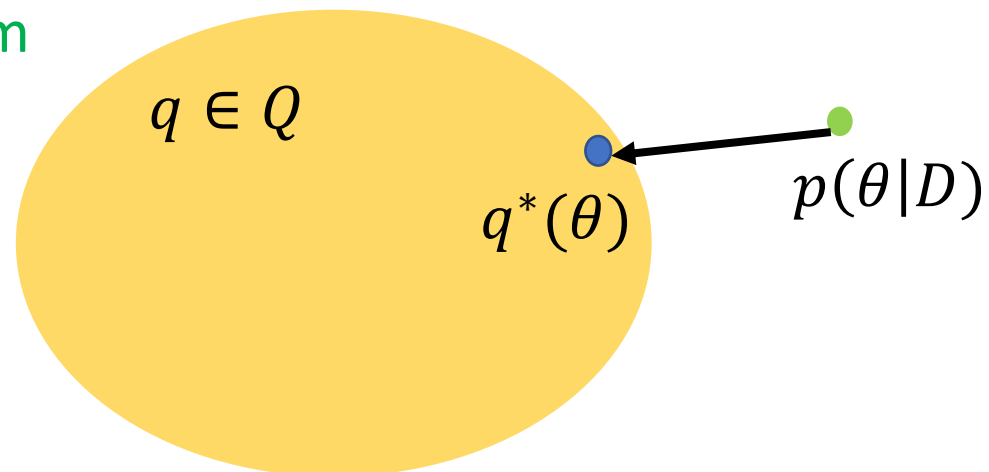
VI Ingredients

$$\mathbf{L} = E_{\theta \sim q_{\phi}} \left[\log \frac{p(D, \theta)}{q_{\phi}(\theta)} \right] = \log p(D) - \mathbf{KL}[q_{\phi} || p]$$

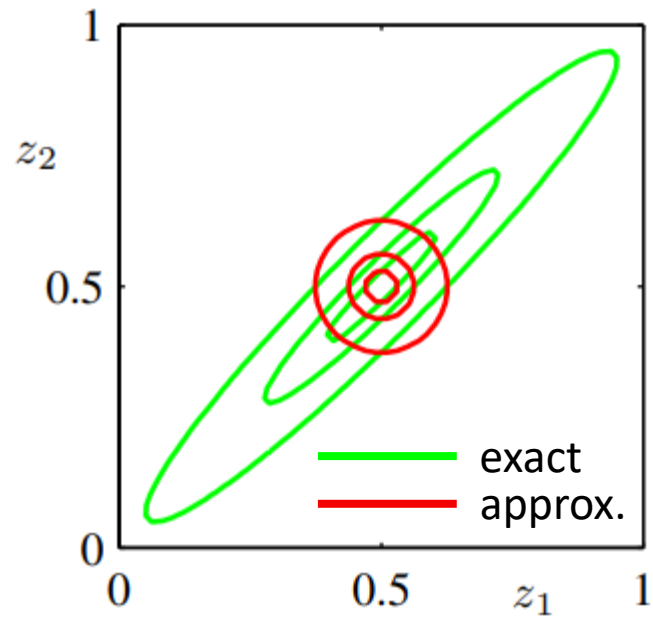
p : your model design

q : your choice of variational distribution, e.g. mean field, flow based

\mathbf{KL} : defines the algorithm



Does It Work?



VI underestimate the uncertainty

(Rényi) α -Divergence

$$\alpha > 0, \alpha \neq 1$$

$$D_\alpha[p||q] = \frac{1}{\alpha - 1} \log \int p(\theta)^\alpha q^{1-\alpha} d\theta$$

$$\alpha = 1$$

$$D_1[p||q] = \lim_{\alpha \rightarrow 1} D_\alpha(p|q) = KL(p||q)$$

VI with α -Divergence

ELBO

$$L = E_{\theta \sim q_\phi} \left[\log \frac{p(D, \theta)}{q_\phi(\theta)} \right] = \log p(D) - \boxed{KL[q_\phi || p]}$$

Variational Rényi bound:

$$L_\alpha = \frac{1}{1 - \alpha} E_{\theta \sim q_\phi} \left[\left(\log \frac{p(D, \theta)}{q_\phi(\theta)} \right)^{1 - \alpha} \right] = \log p(D) - \boxed{D_\alpha[q_\phi || p]}$$

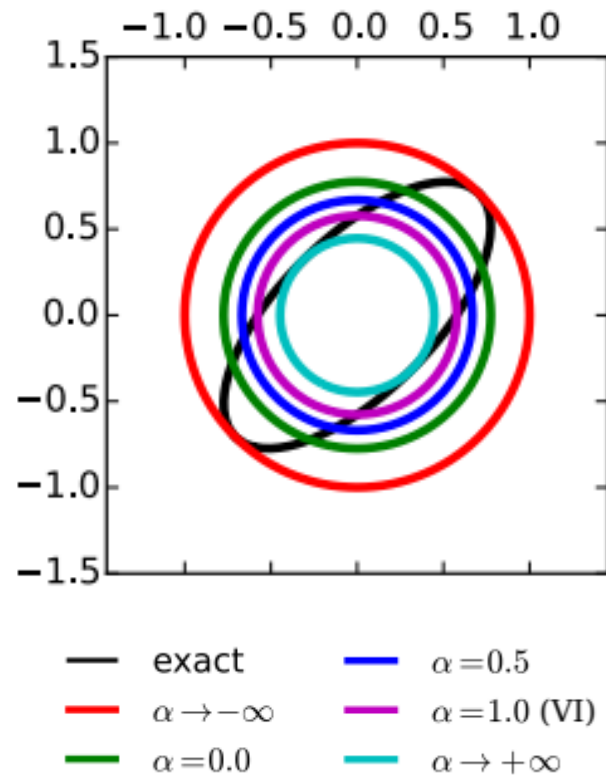
$$\lim_{\alpha \rightarrow 1} L_\alpha = L$$

Li and Turner. Rényi Divergence Variational Inference. NeurIPS 2016

Dieng et al. Variational Inference via χ -Upper Bound Minimization. NeurIPS 2017

Minka, Tom. Divergence measures and message passing. Technical report, Microsoft Research, 2005.

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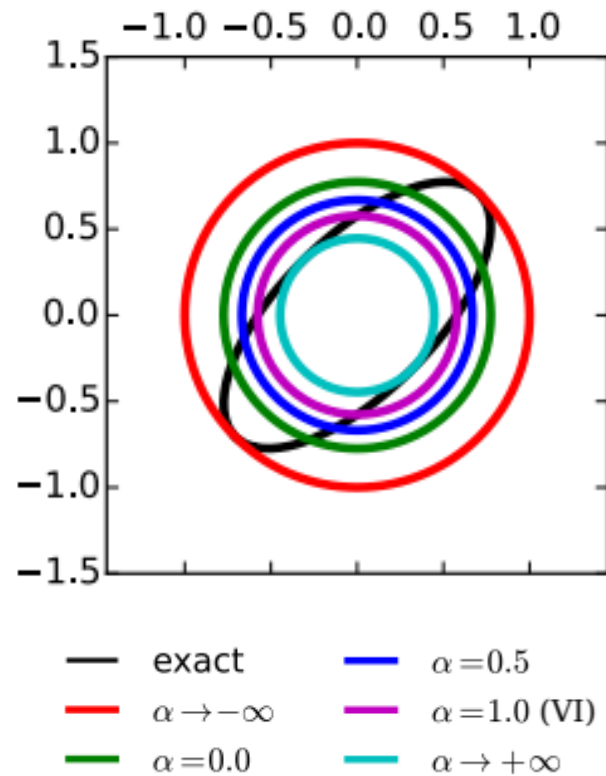


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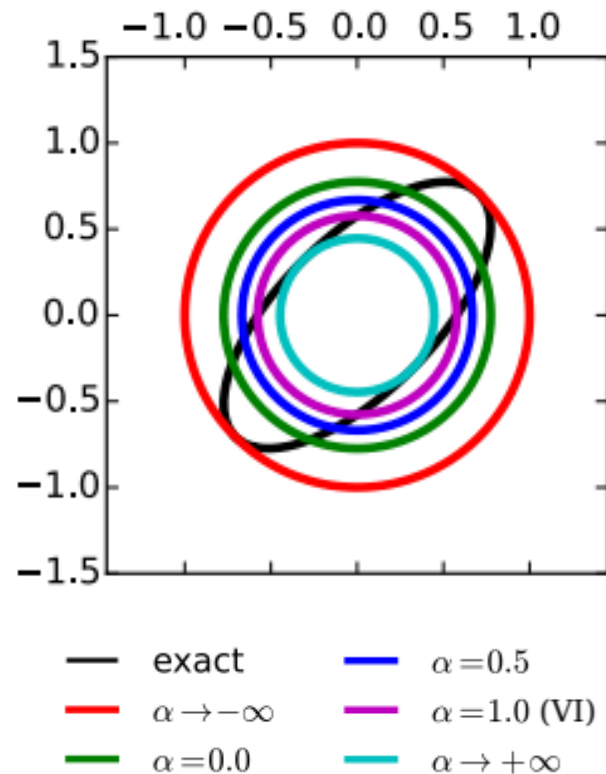
How to choose alpha?

Li and Turner. Rényi Divergence Variational Inference. NeurIPS 2016

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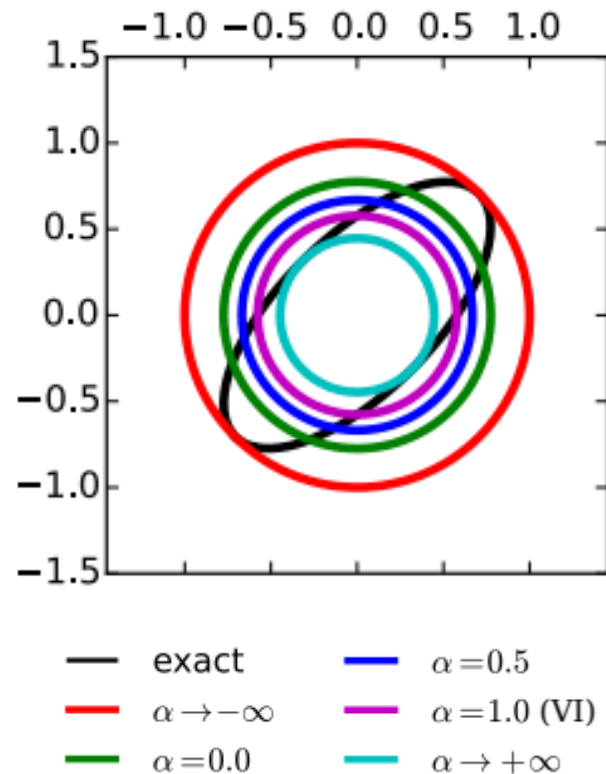
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Too small or too big alpha leads to extremely big variances

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Revisiting Perturbation Theory for VI

$$V(x, z) \equiv \log q_\lambda(z) - \log p(x, z)$$
$$\log p(x) = \log \left(E_{z \sim q_\lambda} \left[\frac{p(x, z)}{q_\lambda(z)} \right] \right) \stackrel{V(x, z)}{=} \log \left(E_{z \sim q_\lambda} \left[e^{-\beta V(x, z)} \right] \right) \Big|_{\beta=1}$$

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Taylor expansion around $\beta = 1$:

$$\log p(x) \approx \boxed{E_{q_\lambda}[-V]} + \frac{1}{2} \left[(V - E_{q_\lambda}[-V])^2 \right] - \frac{1}{3!} \left[(V - E_{q_\lambda}[-V])^3 \right] \\ + \frac{1}{4!} \left[(V - E_{q_\lambda}[-V])^4 \right] - \dots$$

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$$E_{q_\lambda}[-V(x, z)] = E_{q_\lambda}[\log p(x, z) - \log q_\lambda(z)]$$

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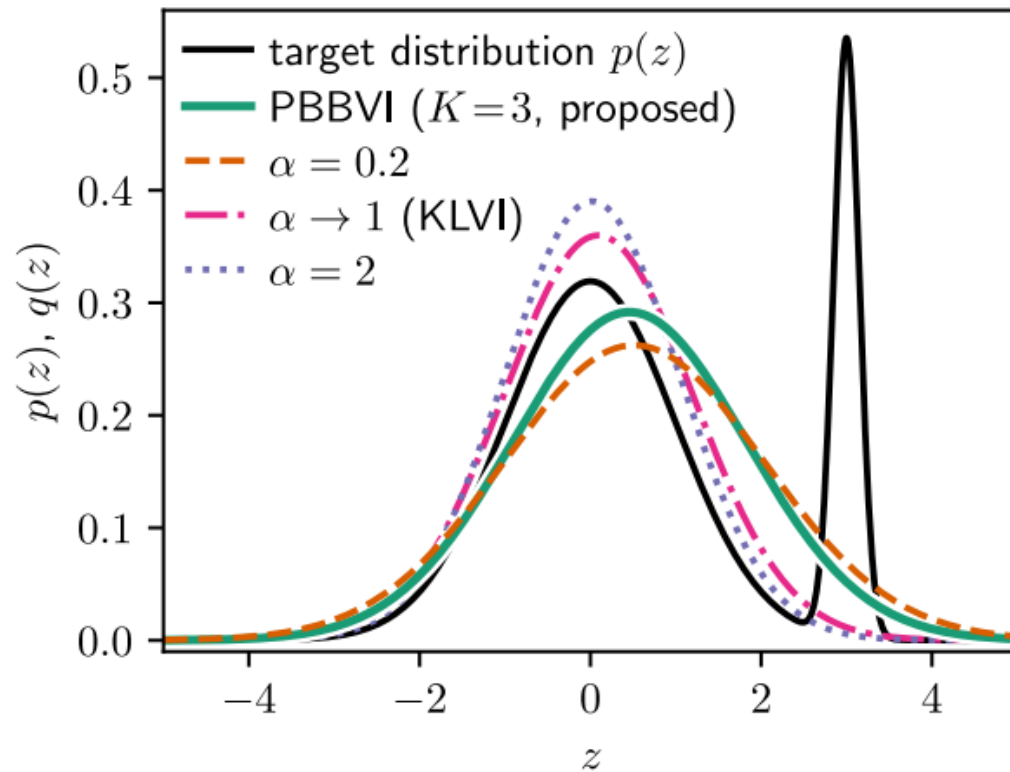
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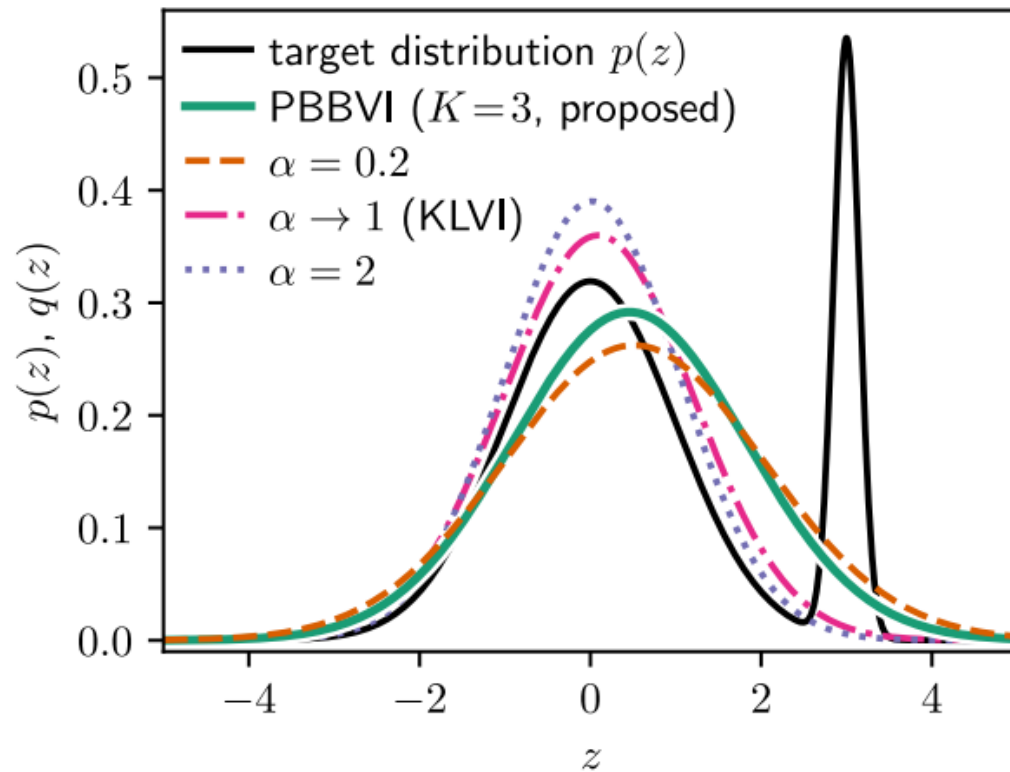
Truncation at any odd number term provides a bound.

Behaviour of PBBVI



- Better uncertainty estimation than KLVI
- Better bias-variance trade-off comparing to α -VI

Behaviour of PBBVI



- Better uncertainty estimation than KLVI
- Better bias-variance trade-off comparing to α -VI

Where do we truncate?
Is it flexible enough?

F-Divergence

$$D_f[p||q_\phi] = E_{\theta \sim q_\phi} \left[f \left(\frac{p(\theta)}{q_\phi(\theta)} \right) - f(1) \right]$$

F-Divergence

$$D_f[p||q_\phi] = E_{\theta \sim q_\phi} \left[f \left(\frac{p(\theta)}{q_\phi(\theta)} \right) - f(1) \right]$$

$$f(t) = -\log t \quad \longrightarrow \quad KL(q||p)$$

$$f(t) = t \log t \quad \longrightarrow \quad KL(p||q)$$

$$f(t) = \frac{t^\alpha}{\alpha(\alpha - 1)} \quad \longrightarrow \quad D_\alpha(p||q)$$

Integral Probability Metric (IPM)

- Using a test function to describe difference:

$$D[q(z), p(z|x)] = \sup_{f \in F} |E_{q(z)}[f(z)] - E_{p(z|x)}[f(z)]|$$

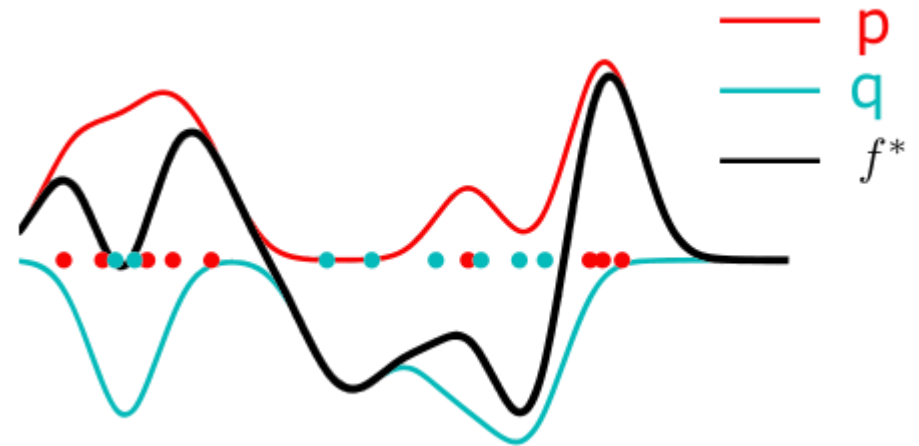


Figure adapted, source: Dougal Sutherland

Integral Probability Metric (IPM)

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$$D[q(z), p(z|x)] = \sup_{f \in F} |E_{q(z)}[f(z)] - E_{p(z|x)}[f(z)]|$$

- Stein discrepancy: only requires $z \sim q(z)$ and $\nabla_z \log p(z|x) = \nabla_z \log p(z, x)$

$$S[q(z), p(z|x)] = \sup_{f \in F_q} |E_{q(z)}[\nabla_z \log p(z, x)^\top f(z) + \nabla_z^\top f(z)]|$$

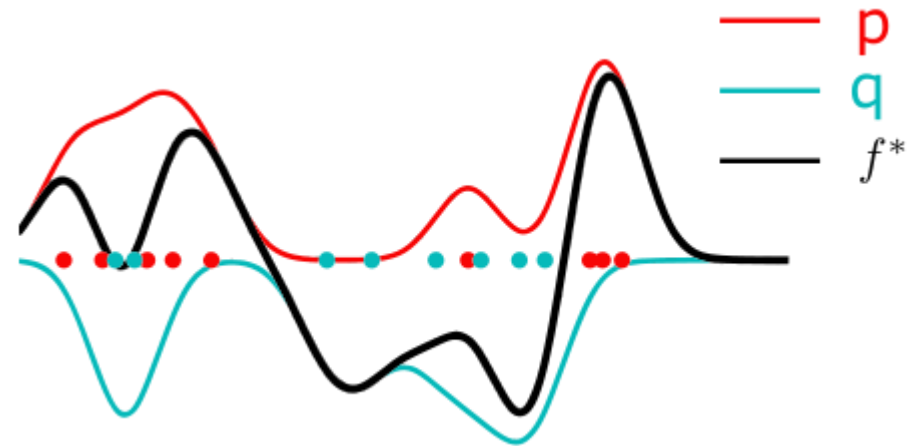
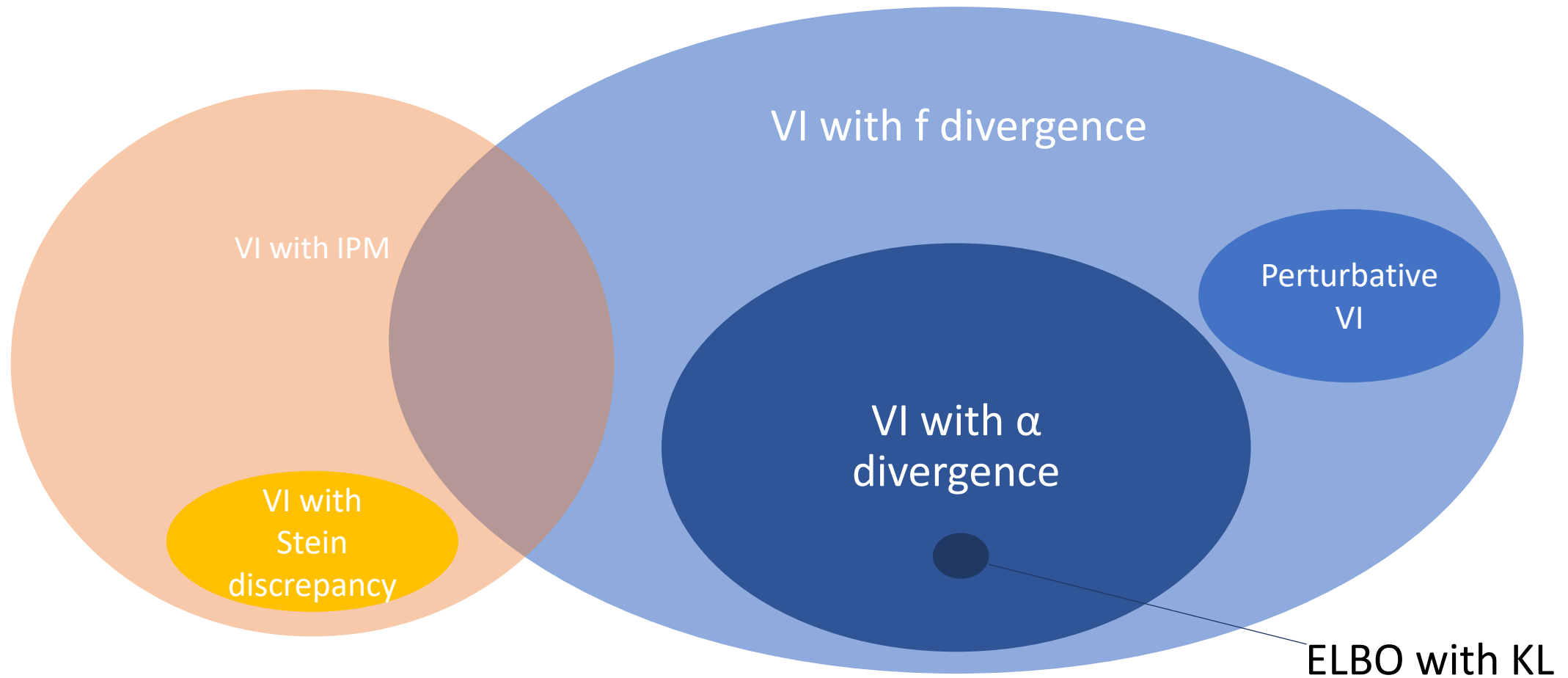
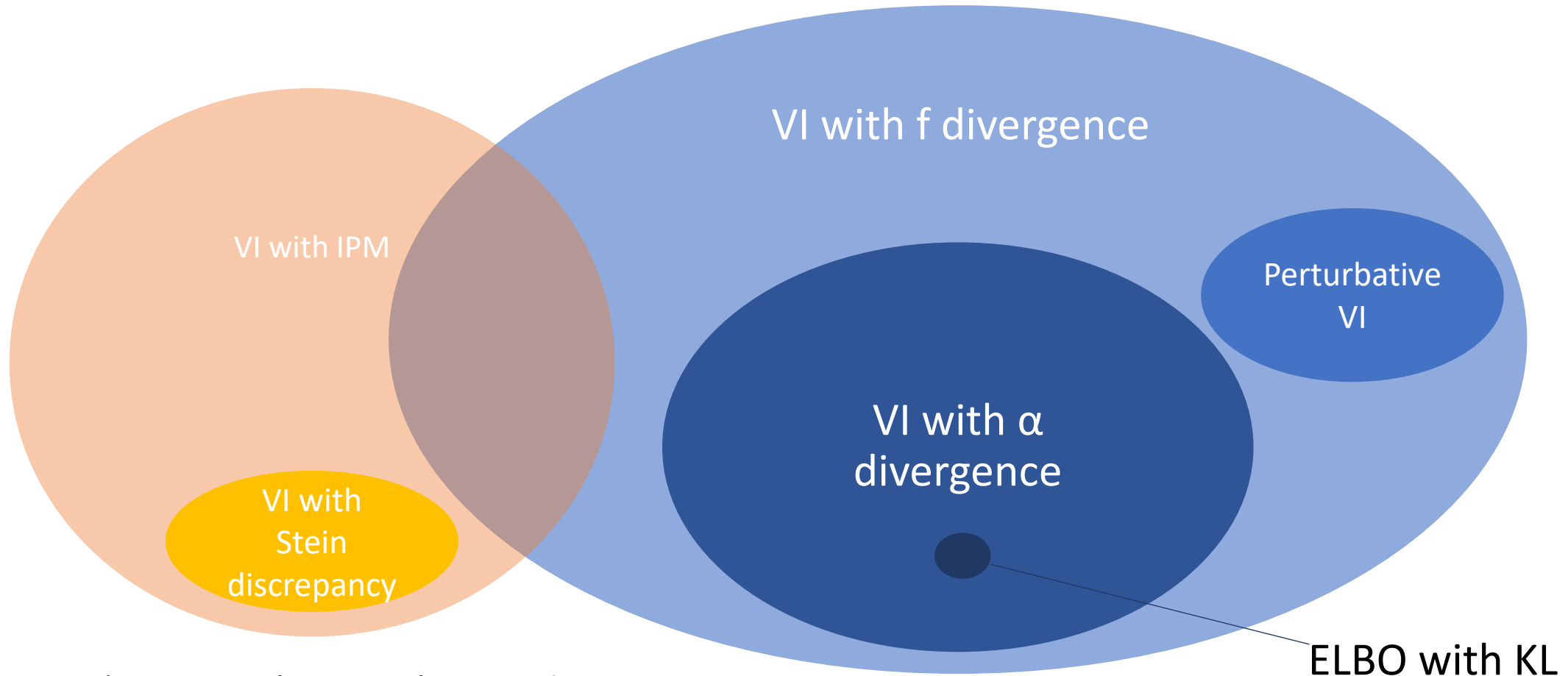


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Looking Back



How to Choose the Inference Algorithm?



Choose divergence by meta-learning!

Improved Monte Carlo Bounds

- Importance weighted auto-encoder (IWAE) bound:

$$L_K(\phi) = E_{z_1, \dots, z_K \sim q(z)} \left[\log \frac{1}{K} \sum_{k=1}^K \frac{p(x, z_k)}{q(z_k)} \right]$$

Importance sampling estimate of $p(x)$


Improved Monte Carlo Bounds

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$\log p(x)$ 

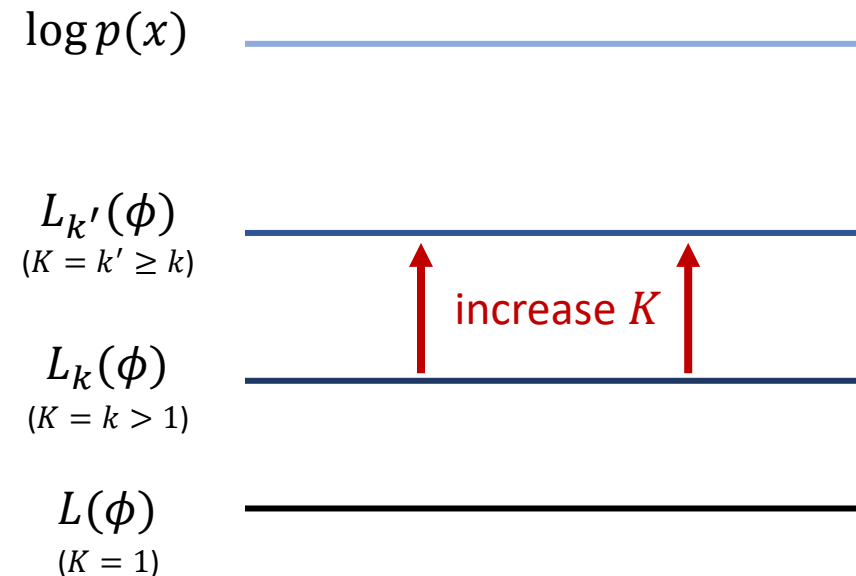
$L(\phi)$
($K = 1$) 

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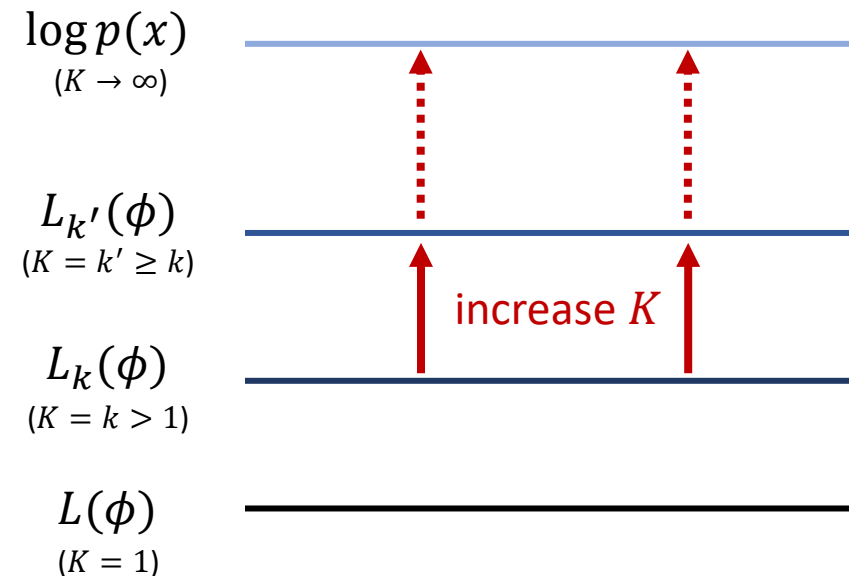


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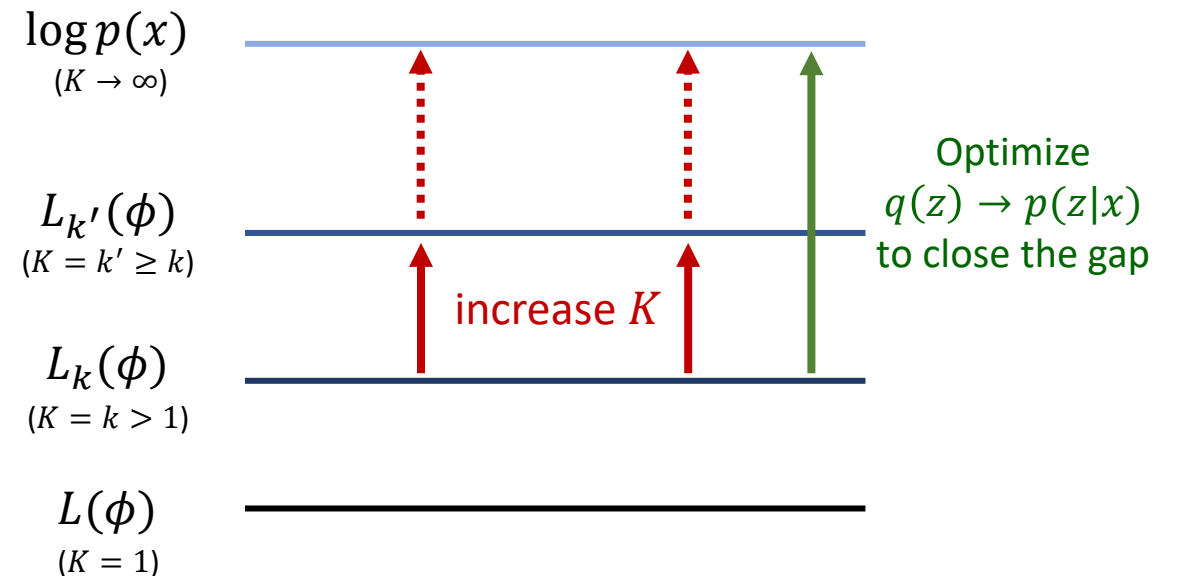


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Importance sampling estimate of $p(x)$



Improved Monte Carlo Bounds

- Constructing lower-bounds from an estimator R of the marginal:

$$E_{q(h)}[R(h, x)] = p(x) \Rightarrow \underline{E_{q(h)}[\log R(h, x)] \leq \log p(x)}$$

Jensen's inequality

Improved Monte Carlo Bounds

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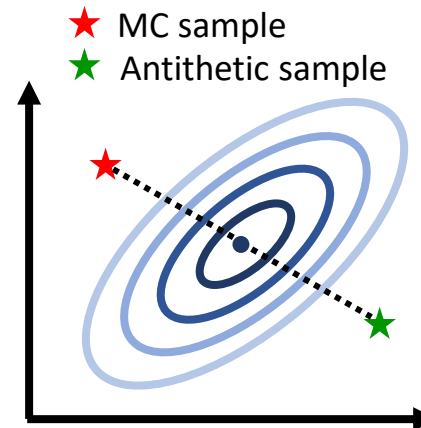
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- Fit q using existing Monte Carlo estimators of $p(x)$

- Example: antithetic sampling with Gaussian $q(z)$:

$$R(z, x) = \frac{p(x, z) + p(x, T(z))}{2q(z)}, \quad T(z) = \mu_q - (z - \mu_q)$$



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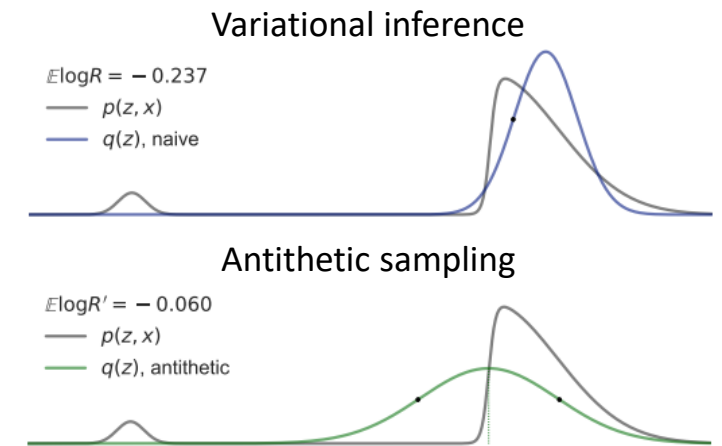
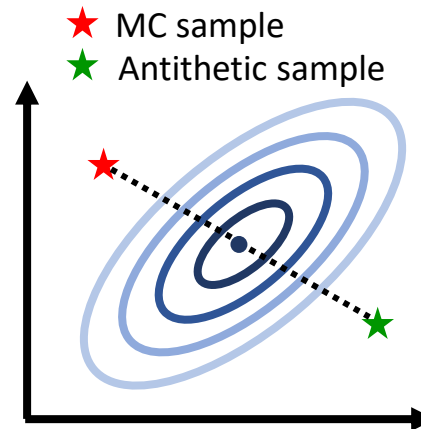
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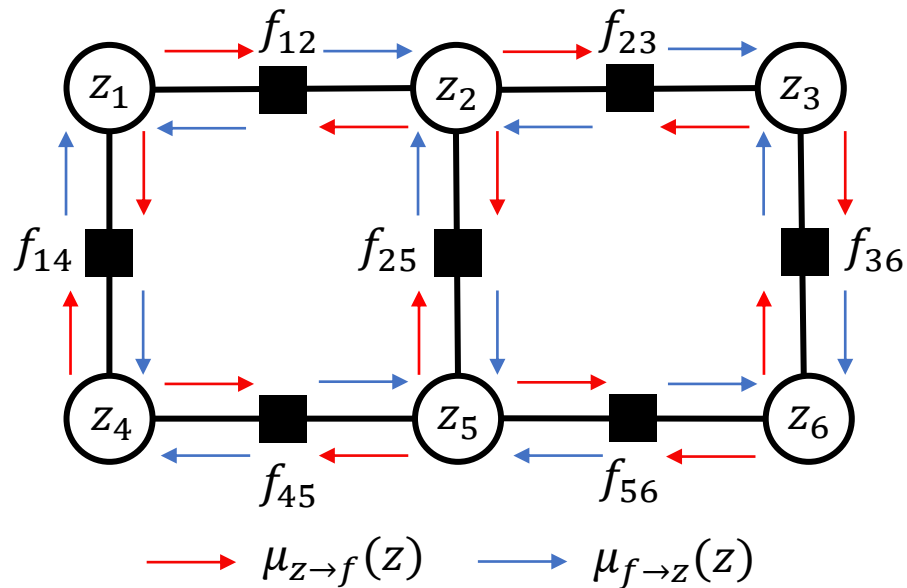
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Free-energy as an Objective

- Bethe free-energy & message passing:




- Both q and the inference algorithm are defined by the *factor graph*
- Optimal q achieved at the fixed point of the *Bethe free energy*

Landscape of Advances



q design

e.g. mean-field: $q(\theta) = \prod_i q(\theta_i)$

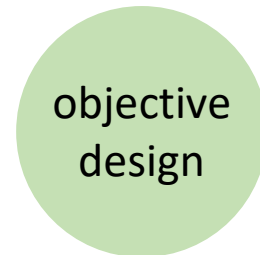
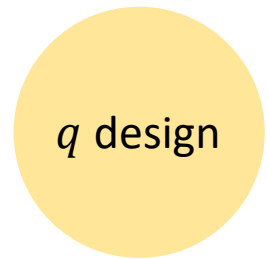


objective
design


variational lower-bound:

$$L(\phi) = E_{q(\theta)}[\log p(D|\theta)] - KL[q(\theta)||p(\theta)]$$

Landscape of Advances



scale-up



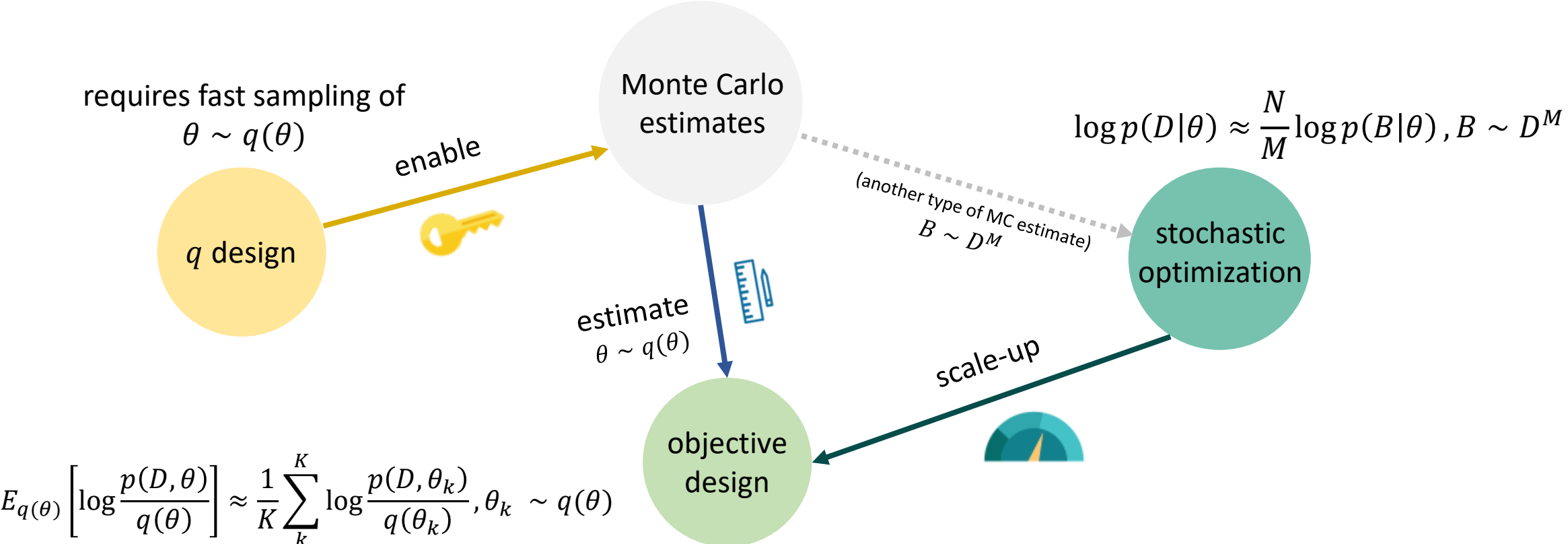
stochastic
optimization

mini-batch training:

$$\log p(D|\theta) \approx \frac{N}{M} \log p(B|\theta), B \sim D^M$$



Landscape of Advances



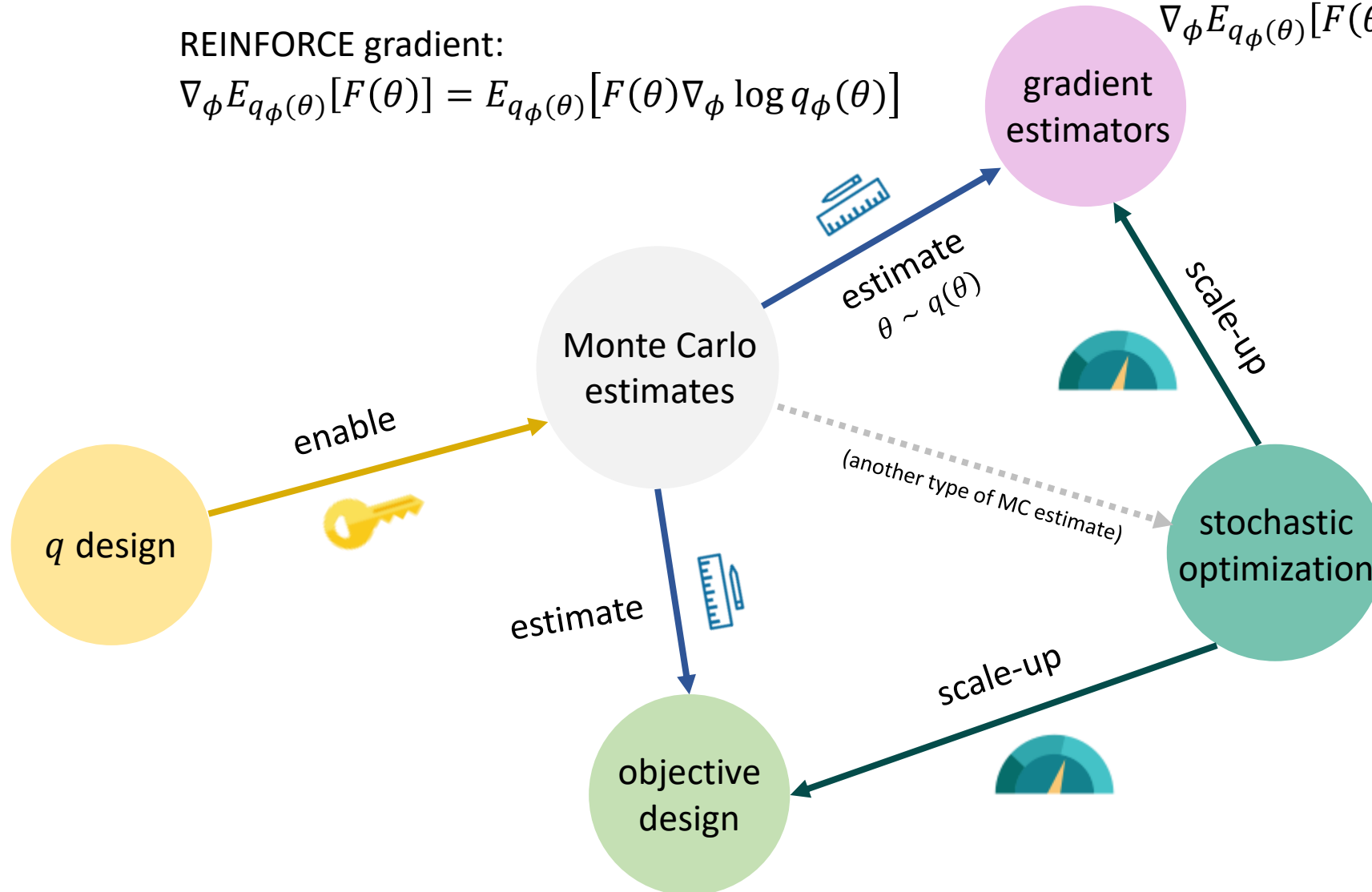
Landscape of Advances

REINFORCE gradient:

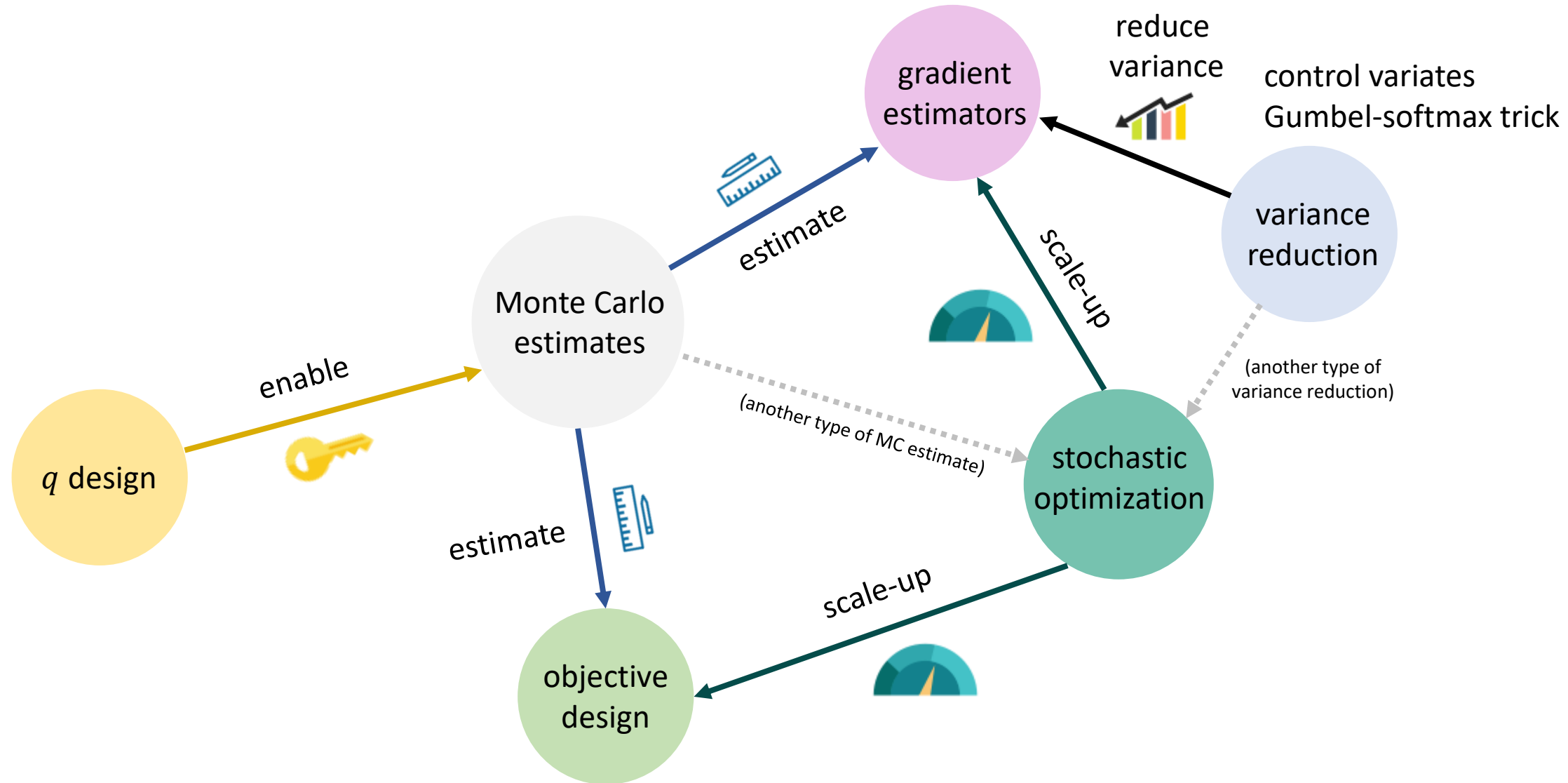
$$\nabla_{\phi} E_{q_{\phi}(\theta)}[F(\theta)] = E_{q_{\phi}(\theta)}[F(\theta) \nabla_{\phi} \log q_{\phi}(\theta)]$$

reparam. trick:

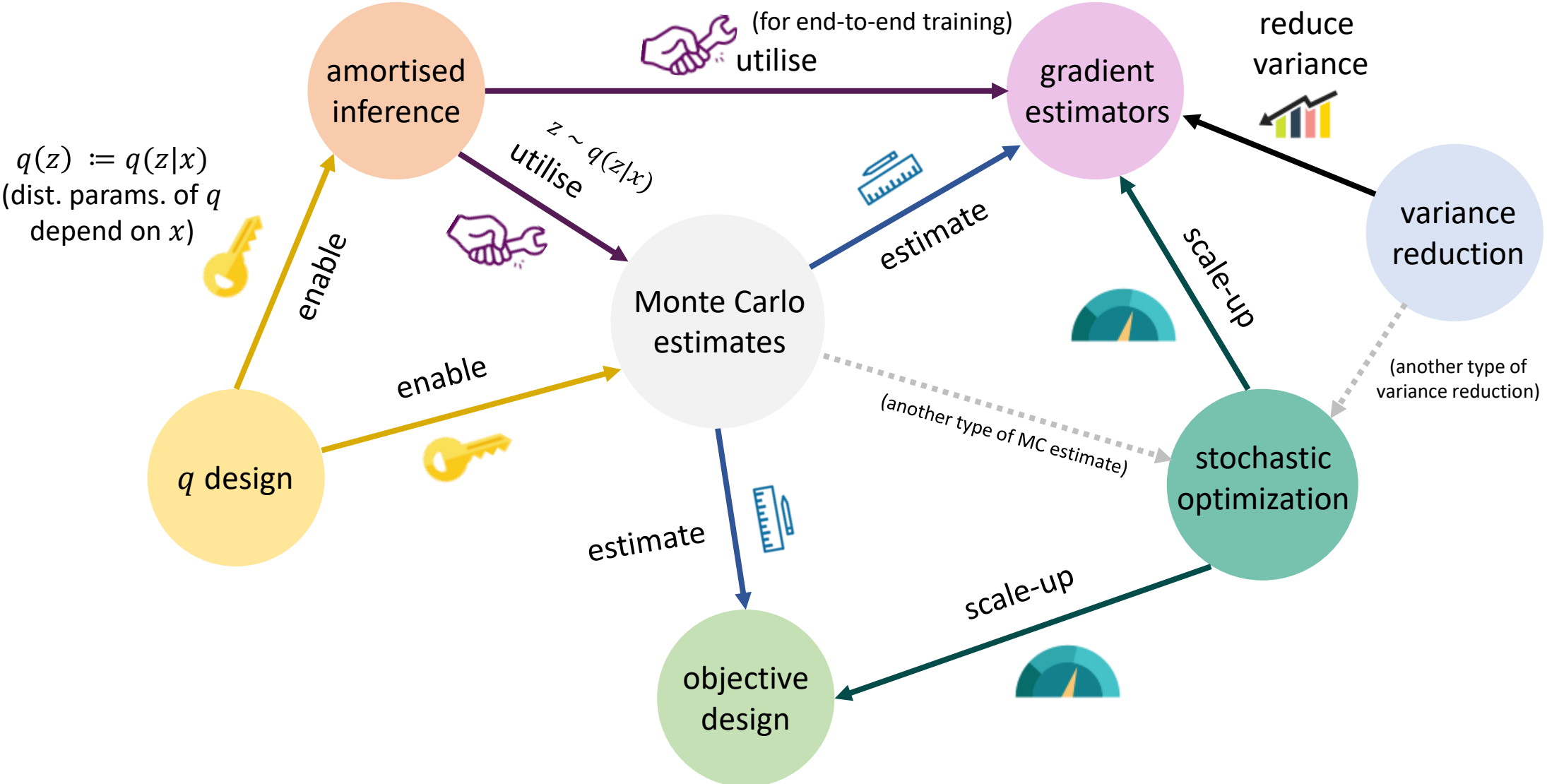
$$\nabla_{\phi} E_{q_{\phi}(\theta)}[F(\theta)] = E_{p(\epsilon)}[\nabla_{\phi} F(g_{\phi}(\epsilon))]$$



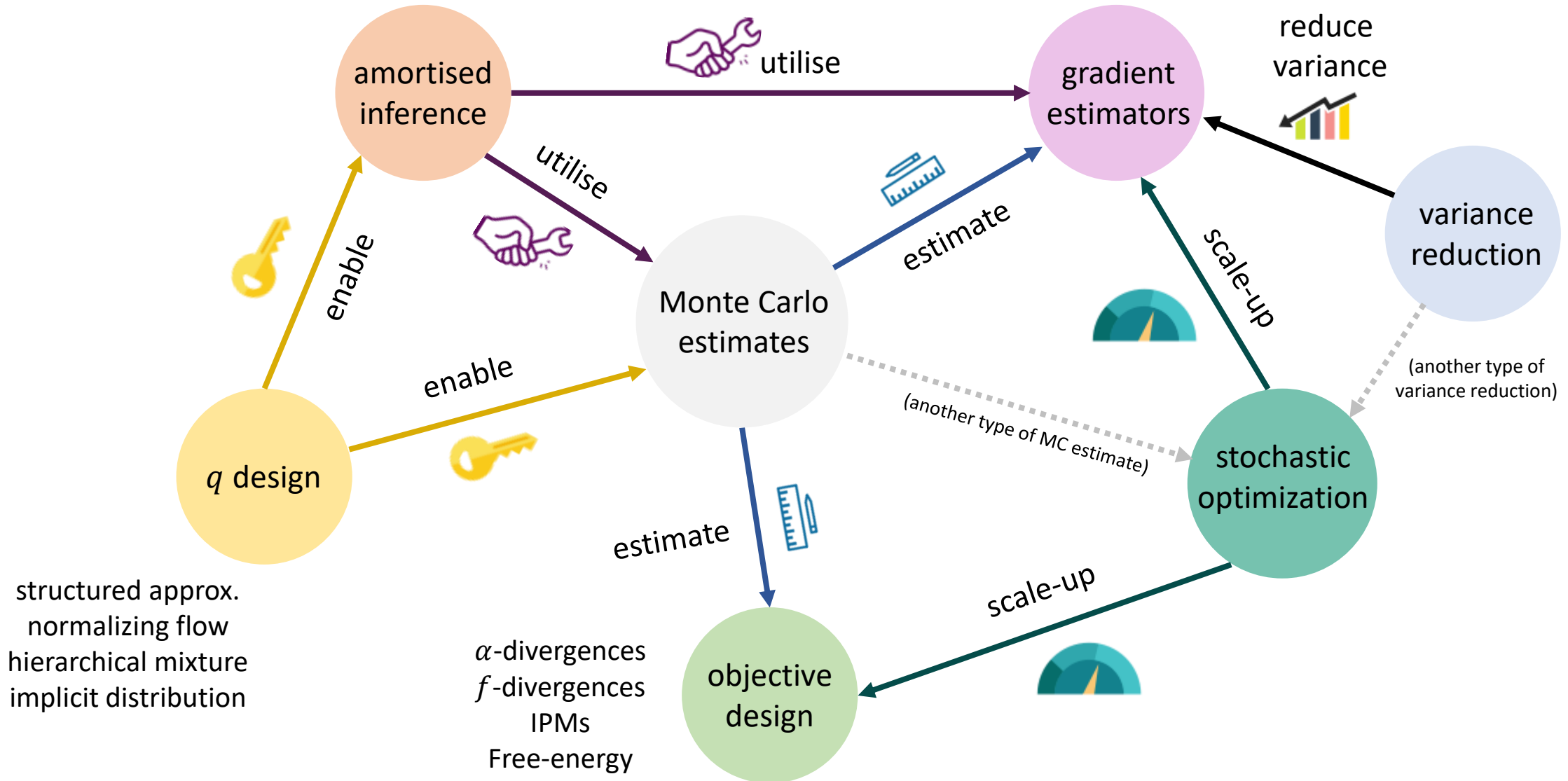
Landscape of Advances



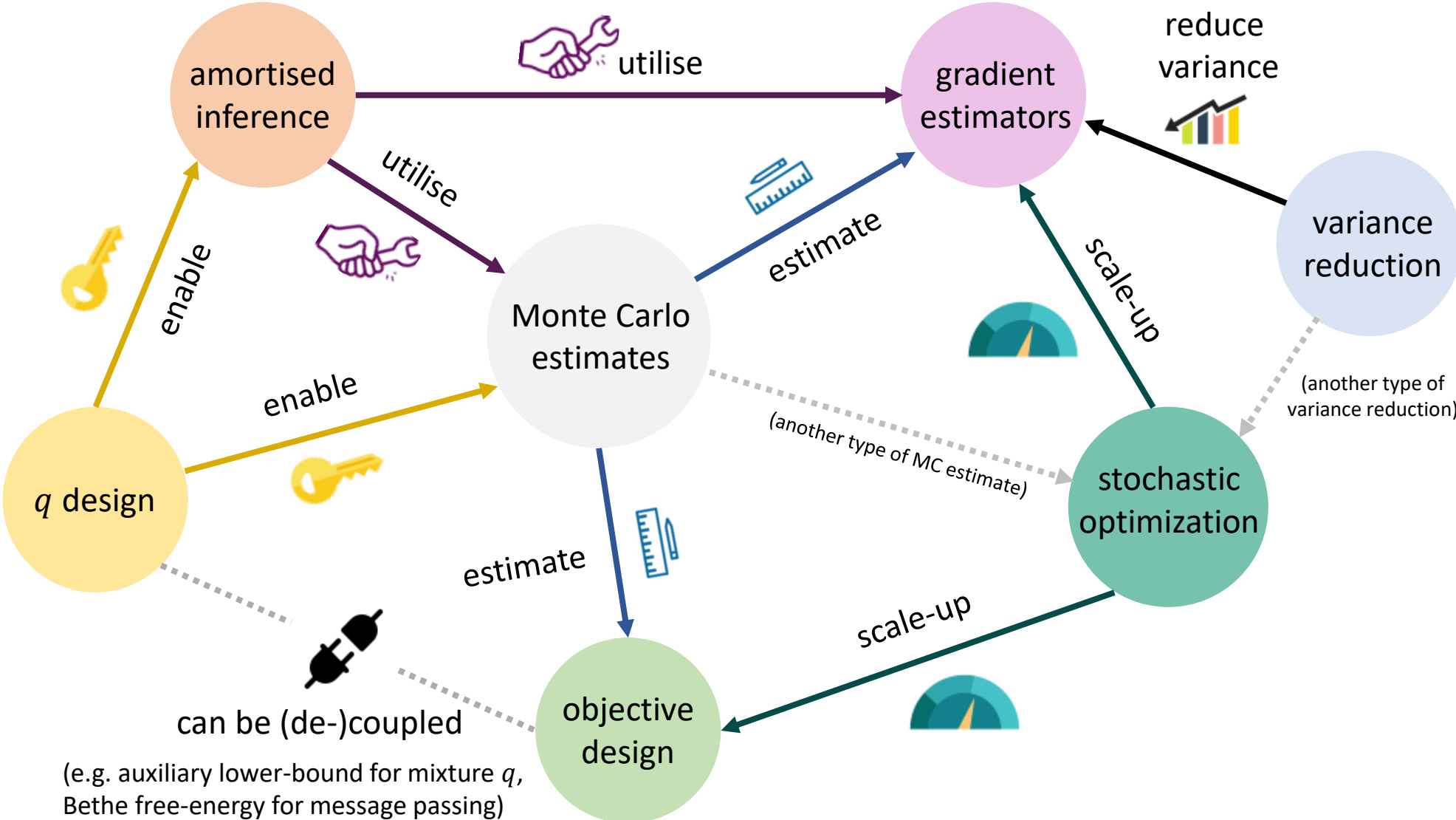
Landscape of Advances



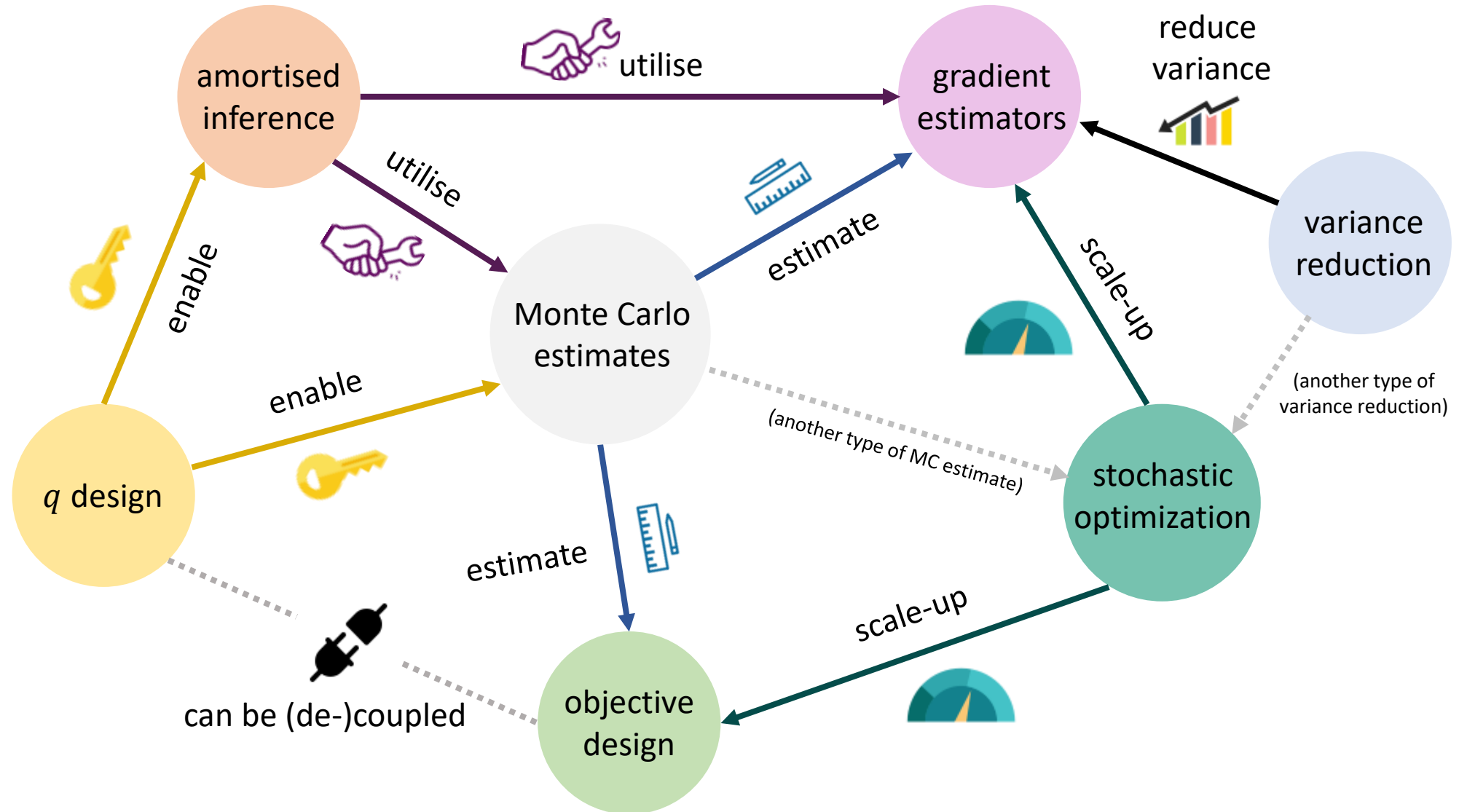
Landscape of Advances

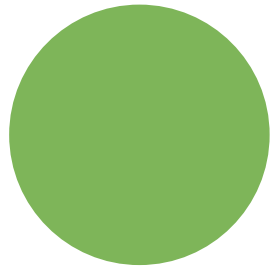


Landscape of Advances




Landscape of Advances







Part III: Applications

- Bayesian neural networks
 - Generative models for decision making
 - Future directions
- 

Why Estimating Uncertainty in DL?

- Models are often over-parameterised
 - E.g. BERT, GPT-3 in NLP
 - E.g. ResNet-152 for vision tasks

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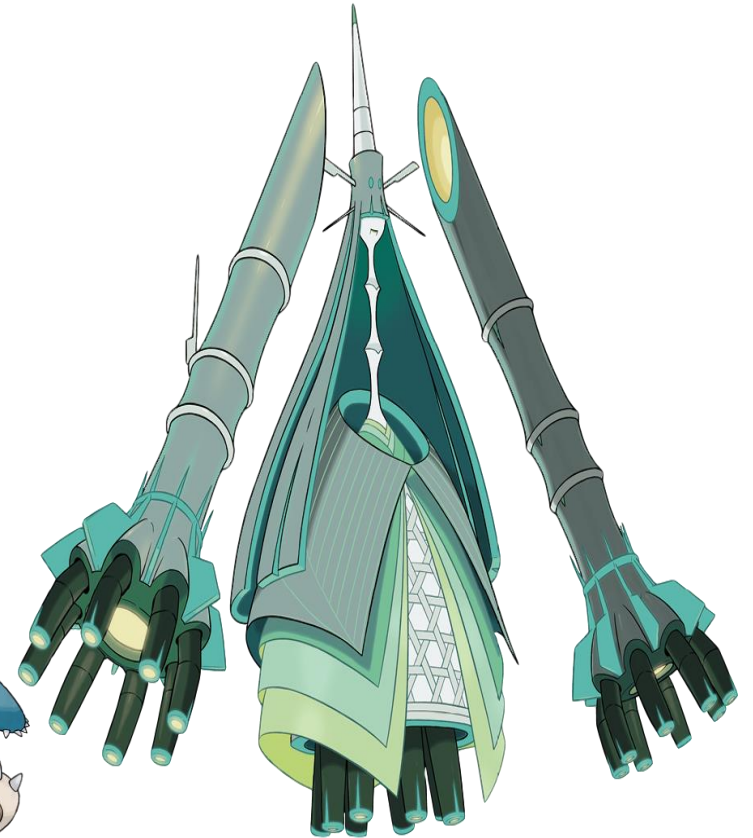
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ResNet-152
(~60 million)



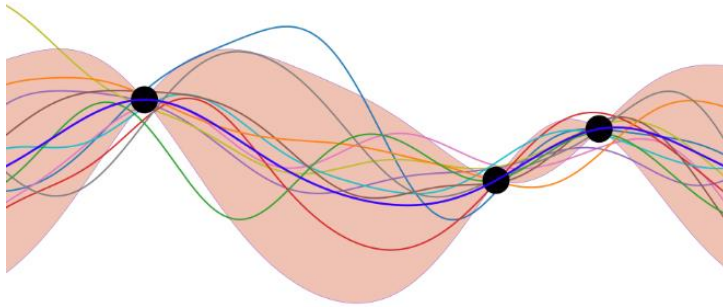
BERT
(~110 million)



GPT-3
(~175 billion)

Why Estimating Uncertainty in DL?

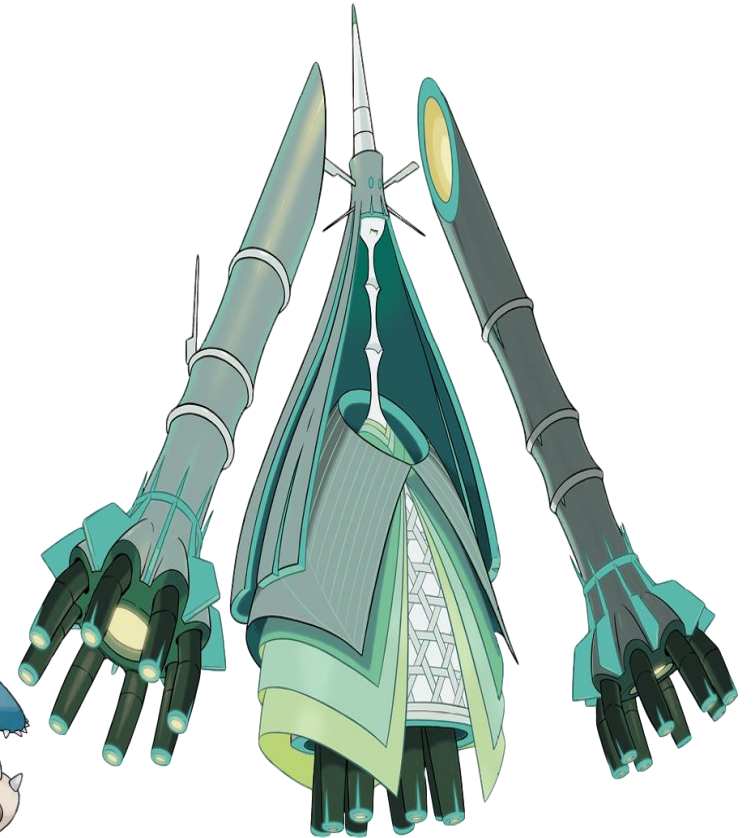
- Models are often over-parameterised
 - E.g. BERT, GPT-3 in NLP
 - E.g. ResNet-152 for vision tasks
- Multiple parameter settings can fit the same data
 - They might provide different predictions on test data



ResNet-152
(~60 million)



BERT
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GPT-3
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Why Estimating Uncertainty in DL?

- Critical tasks need uncertainty estimates to assist decision making
 - Inform end users when uncertain, for safe decision making



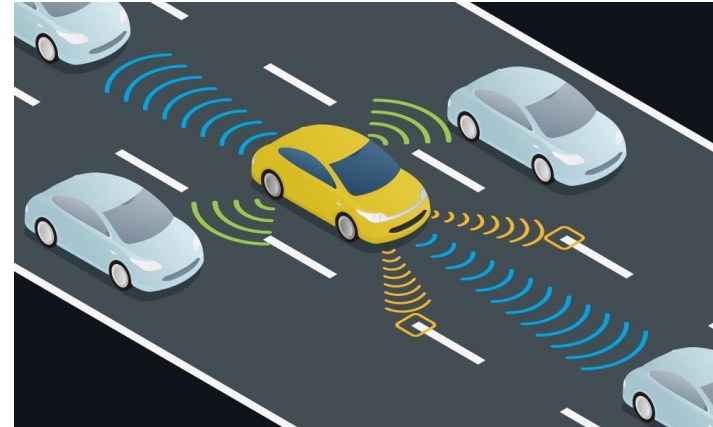
Healthcare AI

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Healthcare AI



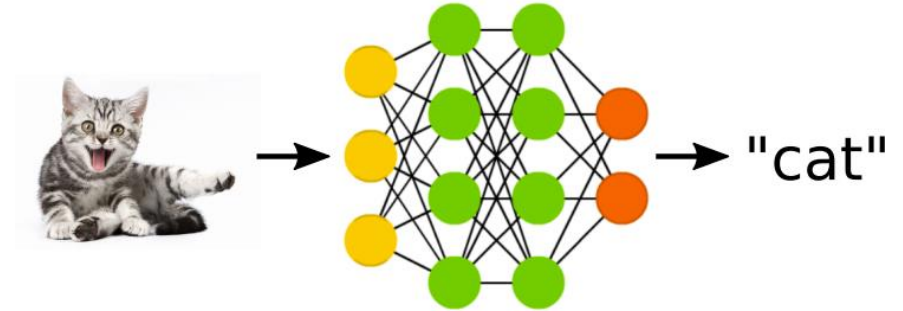
Autonomous driving

Bayesian Neural Network (BNN) 101

Classifying different types of animals:

- x : input image; y : output label
- Build a neural network with parameters θ :

$$p(y|x, \theta) = \textit{softmax}(f_{\theta}(x))$$

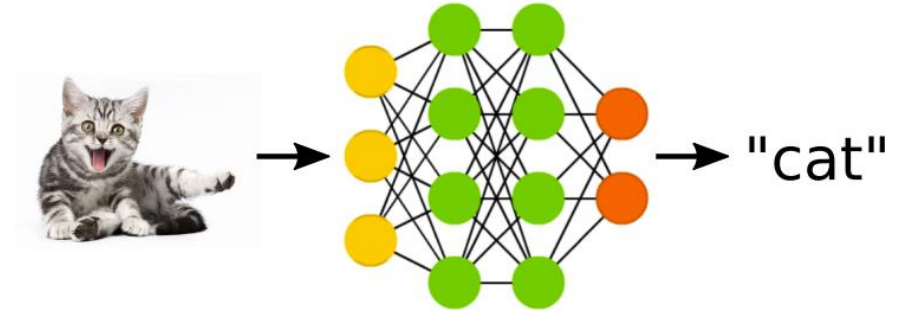


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A typical neural network (with non-linearity $g(\cdot)$):

$$f_{\theta}(x) = W^L g(W^{L-1} g(\dots g(W^1 x + b^1)) + b^{L-1}) + b^L,$$

$$h^l = g(W^l h^{l-1} + b^l), h^1 = g(W^1 x + b^1).$$

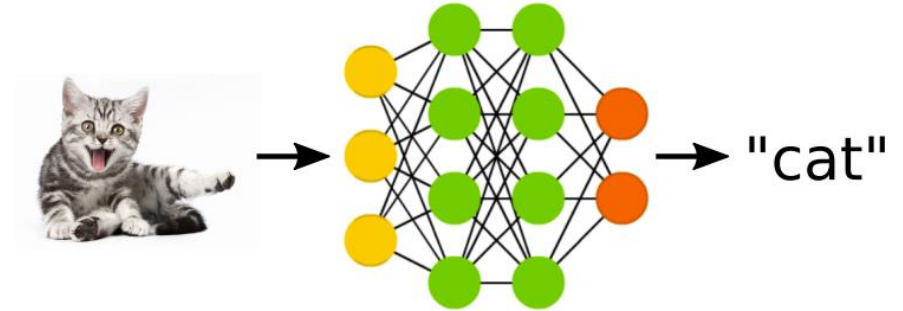
Neural network parameters: $\theta = \{W^l, b^l\}_{l=1}^L$

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Typical deep learning solution:

- Optimize θ to obtain a point estimates (MLE):

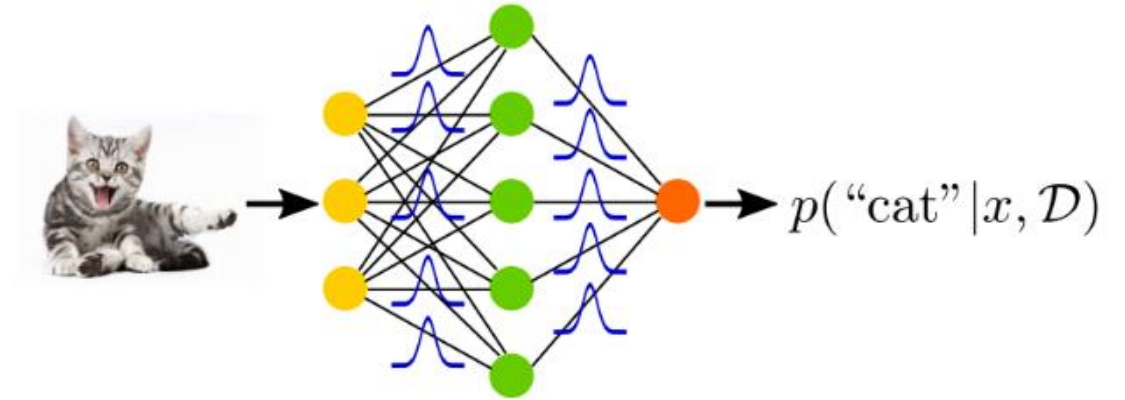
$$\theta^* = \text{argmax} \log p(D | \theta),$$
$$\log p(D | \theta) = \sum_{n=1}^N \log p(y_n | x_n, \theta), D = \{(x_n, y_n)\}_{n=1}^N$$

- Prediction: using $p(y^* | x^*, \theta^*)$

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Bayesian solution:

- Put a prior $p(\theta)$ on network parameters θ , e.g. Gaussian prior

$$p(\theta) = N(\theta; 0, \sigma^2 I)$$

- Compute the posterior distribution $p(\theta | D)$:

$$p(\theta | D) \propto p(D | \theta) p(\theta)$$

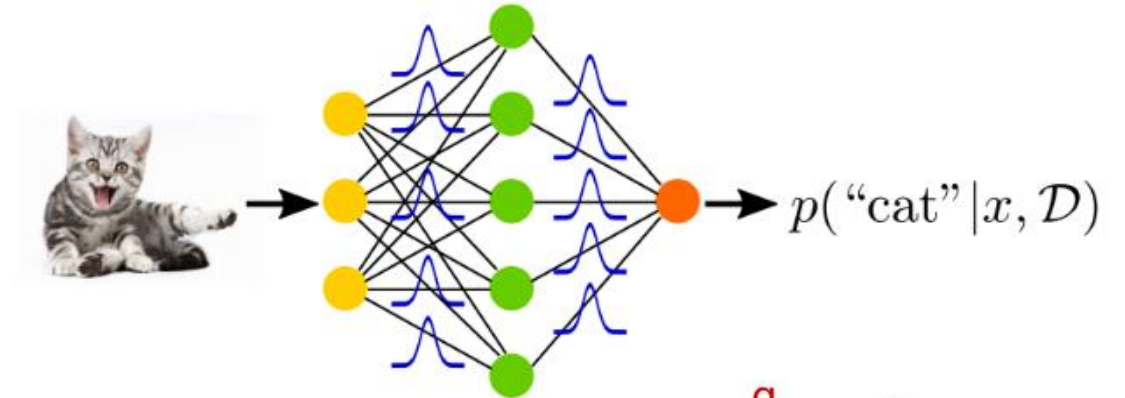
- Bayesian predictive inference:

$$p(y^* | x^*, D) = E_{p(\theta | D)}[p(y^* | x^*, \theta)]$$

Bayesian Neural Network (BNN) 101

Classifying different types of animals:

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Approximate (Bayesian) inference solution:

- Exact posterior intractable, use approximate posterior:

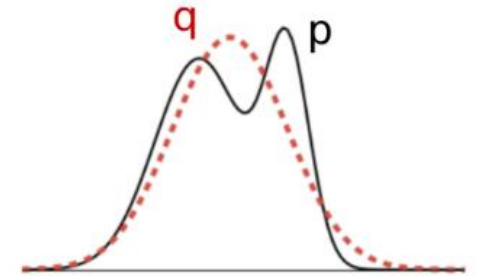
$$q(\theta) \approx p(\theta | D)$$

- Approximate Bayesian predictive inference:

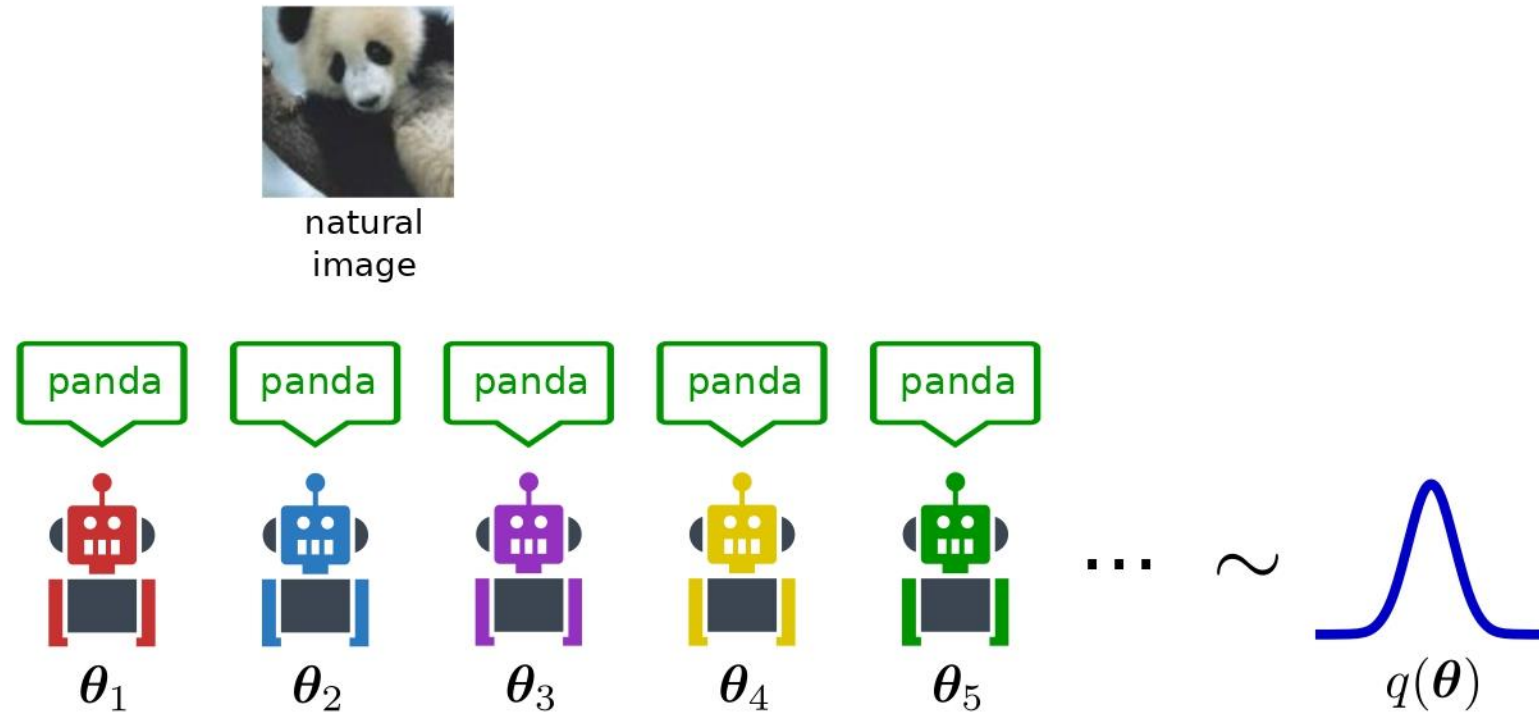
$$p(y^* | x^*, D) \approx E_{q(\theta)}[p(y^* | x^*, \theta)]$$

- Monte Carlo approximation:

$$p(y^* | x^*, D) \approx \frac{1}{K} \sum_{k=1}^K p(y^* | x^*, \theta_k), \quad \theta_k \sim q(\theta)$$

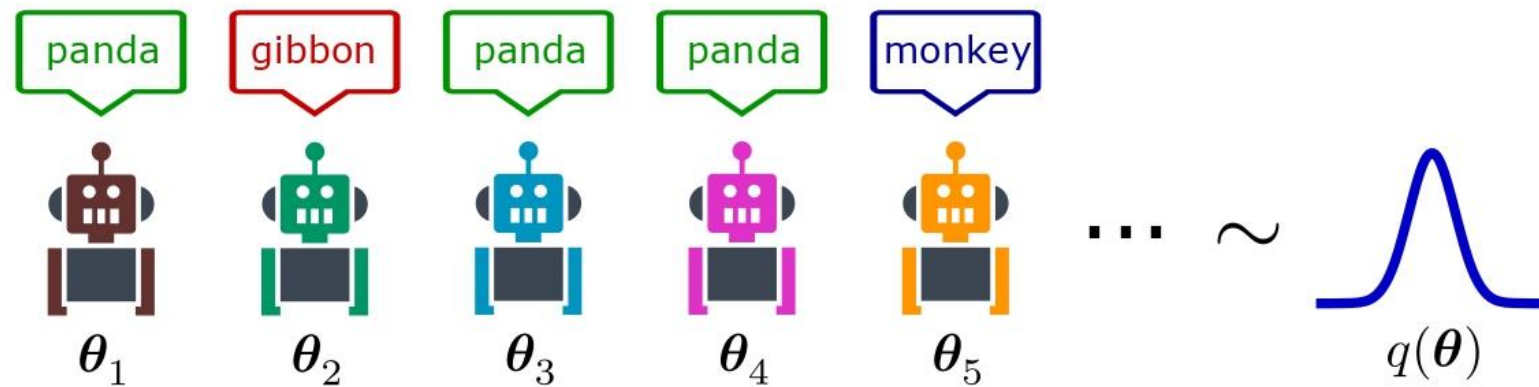


Bayesian Neural Network (BNN) 101



Prediction on in-distribution data:
ensemble over networks, using weights sampled from $q(\theta)$

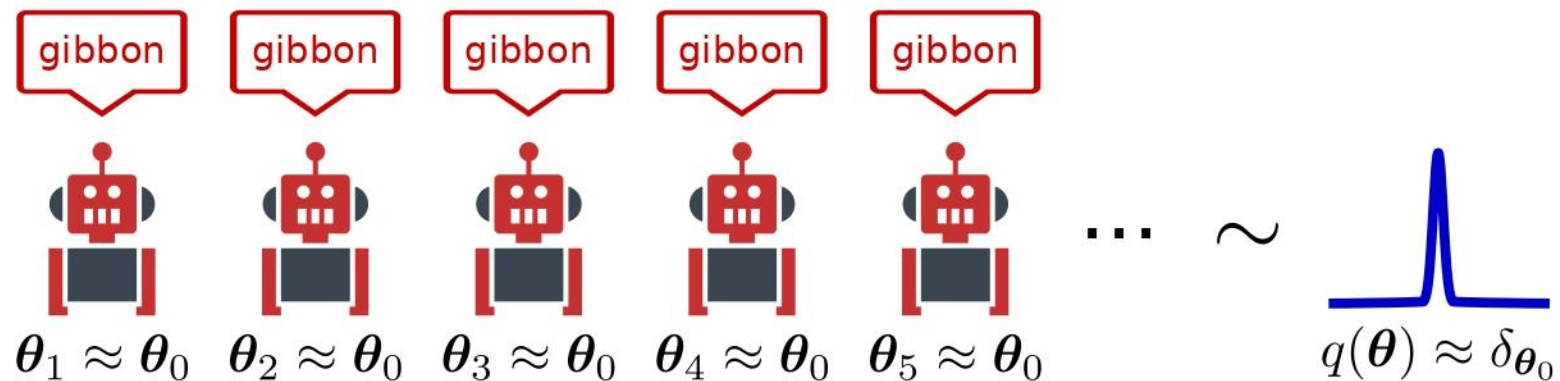
Bayesian Neural Network (BNN) 101



Prediction on OOD/noisy/adversarial data:

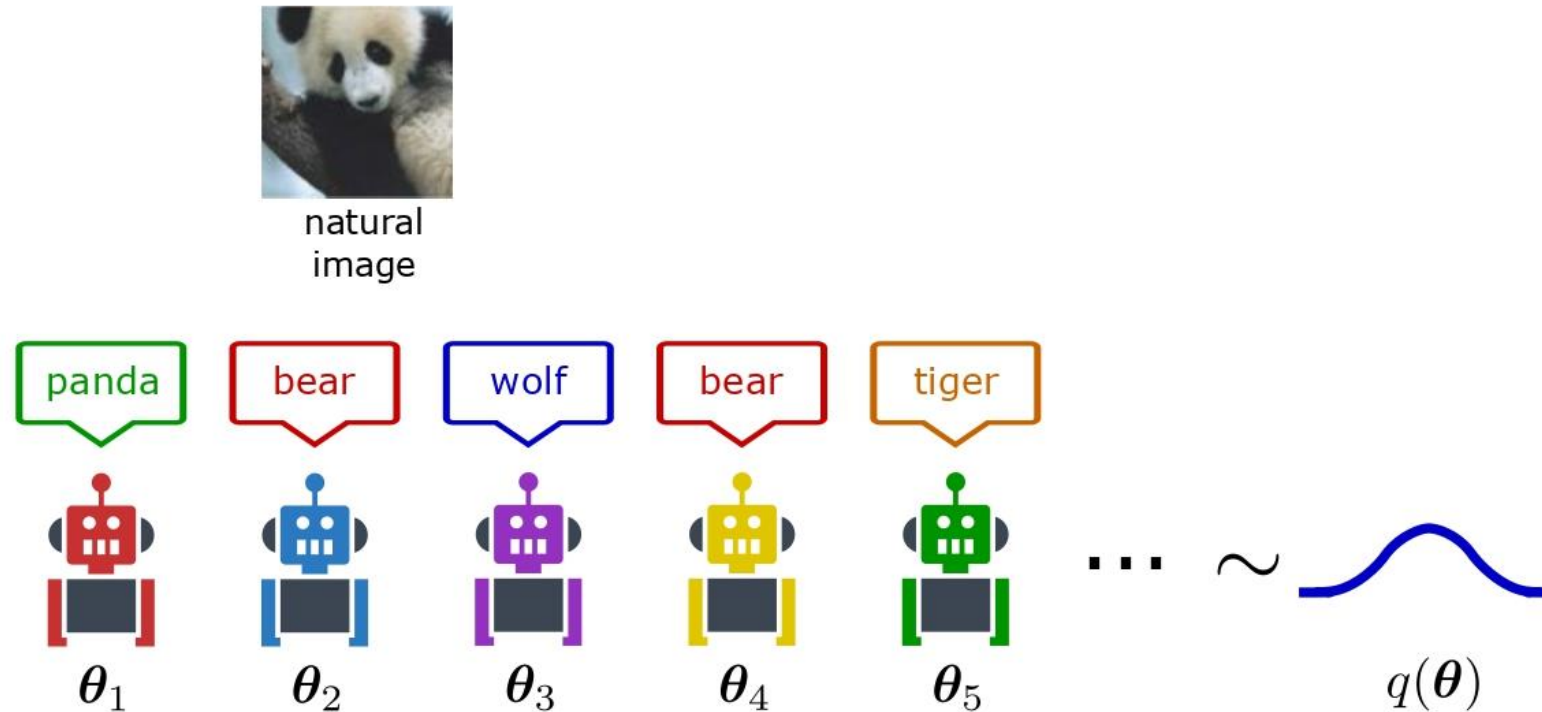
Disagreement (i.e. uncertainty) exists over networks sampled from $q(\theta)$

Bayesian Neural Network (BNN) 101



Prediction on OOD/noisy/adversarial data **when $q(\theta)$ is over-confident:**
Return **confidently wrong answers** (close to point estimate)

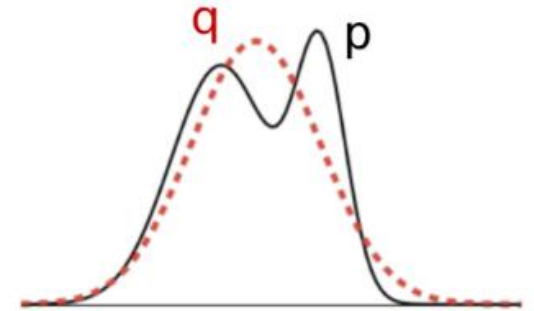
Bayesian Neural Network (BNN) 101



Prediction on in-distribution data **when $q(\theta)$ is under-confident:**
Low accuracy in prediction tasks (less desirable)

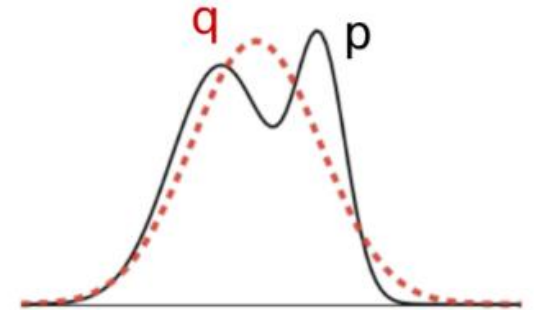
Approximate Inference in BNNs

- Key steps of approximate inference in BNNs
 1. Construct the $q(\theta) \approx p(\theta | D)$ distribution
 - Simple distributions: e.g. Mean-field Gaussian
 - Structured approximations, e.g. low-rank Gaussians
 - Others (non-Gaussian)



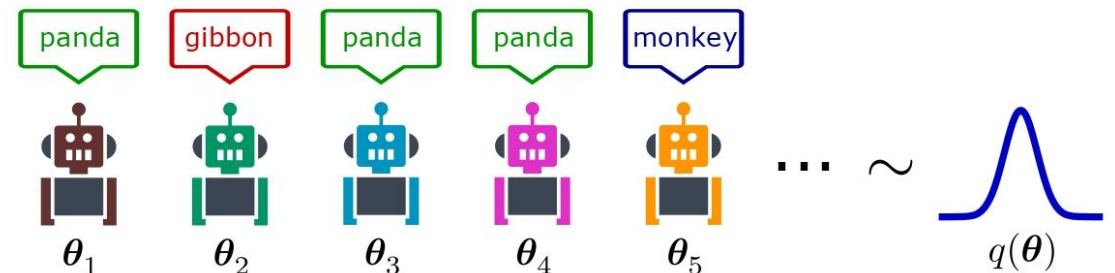
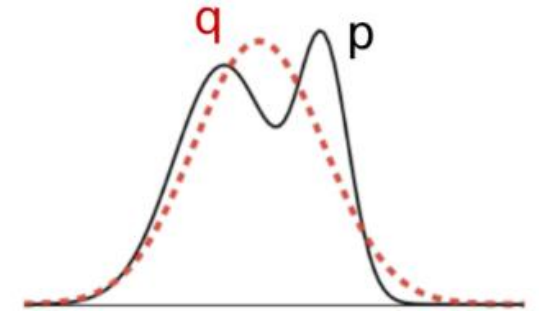
Approximate Inference in BNNs

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 2. Fit the $q(\theta)$ distribution
 - E.g. with variational inference



Approximate Inference in BNNs

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 - Others (non-Gaussian)
 2. Fit the $q(\theta)$ distribution
 - E.g. with variational inference
 3. Compute prediction with Monte Carlo approximations



Approximate Inference in BNNs

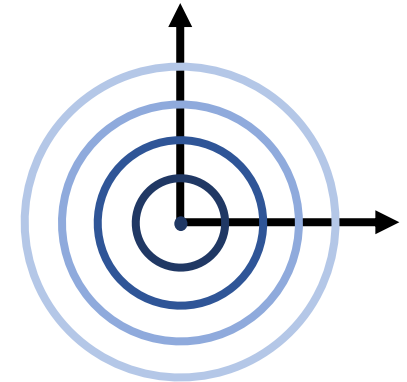
- Step 1: construct the $q(\theta) \approx p(\theta | D)$ distribution
 - Example: Mean-field Gaussian distribution:

$$q(\theta) = \prod_{l=1}^L q(W^l) q(b^l)$$

$$q(W_l) = \prod_{ij} q(W_{ij}^l), \quad q(W_{ij}^l) = N(W_{ij}^l; M_{ij}^l, V_{ij}^l)$$

$$q(b^l) = \prod_i q(b_i^l), \quad q(b_i^l) = N(b_i^l; m_i^l, v_i^l)$$

- Variational parameters: $\phi = \{M_{ij}^l, \log V_{ij}^l, m_i^l, \log v_i^l\}_{l=1}^L$



Approximate Inference in BNNs

- Step 2: fit the $q(\theta)$ distribution:

- Variational inference: $\phi^* = \operatorname{argmax} L(\phi)$

$$L(\phi) = E_{q(\theta)}[\log p(D | \theta)] - KL[q(\theta) || p(\theta)]$$

Approximate Inference in BNNs

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- First scalable technique: **Stochastic optimization**

- i.i.d. assumption of data: $\log p(D | \theta) = \sum_{n=1}^N \log p(y_n | x_n, \theta)$
- Enable mini-batch training with $\{(x_m, y_m)\} \sim D^M$:

$$L(\phi) \approx \frac{N}{M} \sum_{m=1}^M E_{q(\theta)}[\log p(y_m | x_m, \theta)] - KL[q(\theta) || p(\theta)]$$

Approximate Inference in BNNs

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reweighting to ensure calibrated
posterior concentration

Approximate Inference in BNNs

- Step 2: fit the $q(\theta)$ distribution:
 - 2nd scalable technique: **Monte Carlo sampling**
 - $E_{q(\theta)}[\log p(y | x, \theta)]$ intractable even with Gaussian $q(\theta)$
 - **Solution: Monte Carlo estimate:**

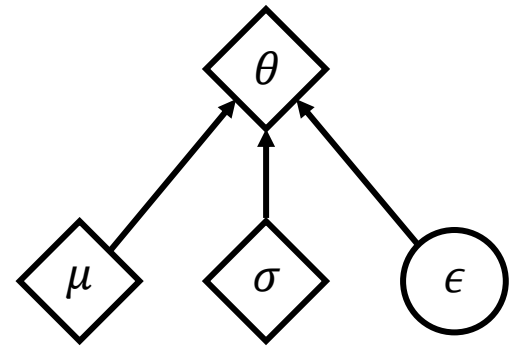
$$E_{q(\theta)}[\log p(y | x, \theta)] \approx \frac{1}{K} \sum_k^K \log p(y | x, \theta_k), \quad \theta_k \sim q(\theta)$$

Approximate Inference in BNNs

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- **Reparameterization trick** to sample mean-field Gaussians:
 $\theta_k \sim q(\theta) \Leftrightarrow \theta_k = m_\theta + \sigma_\theta \epsilon_k, \epsilon_k \sim N(0, I)$



Approximate Inference in BNNs

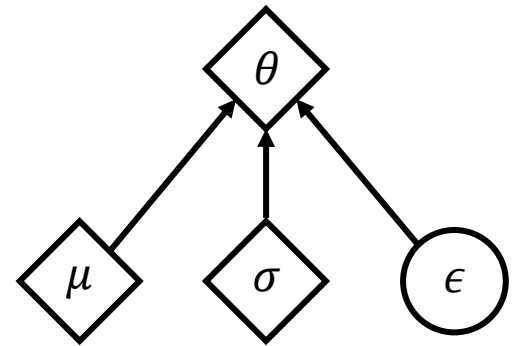
- Step 2: fit the $q(\theta)$ distribution:
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$$E_{q(\theta)}[\log p(y | x, \theta)] \approx \frac{1}{K} \sum_k^K \log p(y | x, \theta_k), \quad \theta_k \sim q(\theta)$$

- **Reparameterization trick** to sample mean-field Gaussians:

$$\theta_k \sim q(\theta) \Leftrightarrow \theta_k = m_\theta + \sigma_\theta \epsilon_k, \quad \epsilon_k \sim N(0, I)$$

$$\Rightarrow E_{q(\theta)} [\log p(y | x, \theta)] \approx \frac{1}{K} \sum_k^K \log p(y | x, m_\theta + \sigma_\theta \epsilon_k), \quad \epsilon_k \sim N(0, I)$$



Approximate Inference in BNNs

- Combining both steps:

$$L(\phi) \approx \frac{N}{M} \sum_{m=1}^M \frac{1}{K} \sum_{k=1}^K \log p(y_m | x_m, \theta_k) - \underline{KL[q(\theta) || p(\theta)]}, \theta_k \sim q(\theta)$$

analytic between two Gaussians

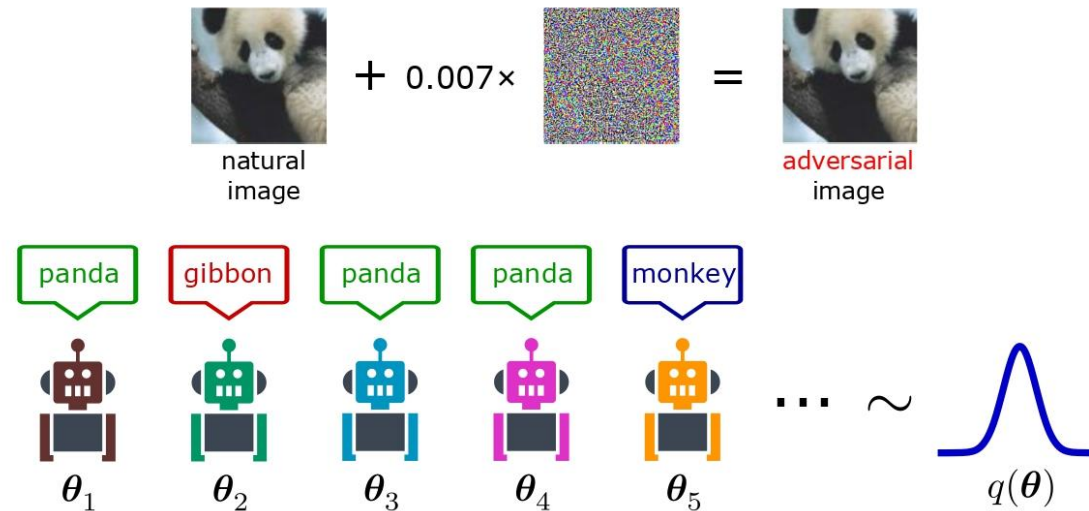
(if not, can also be estimated with Monte Carlo)

Approximate Inference in BNNs

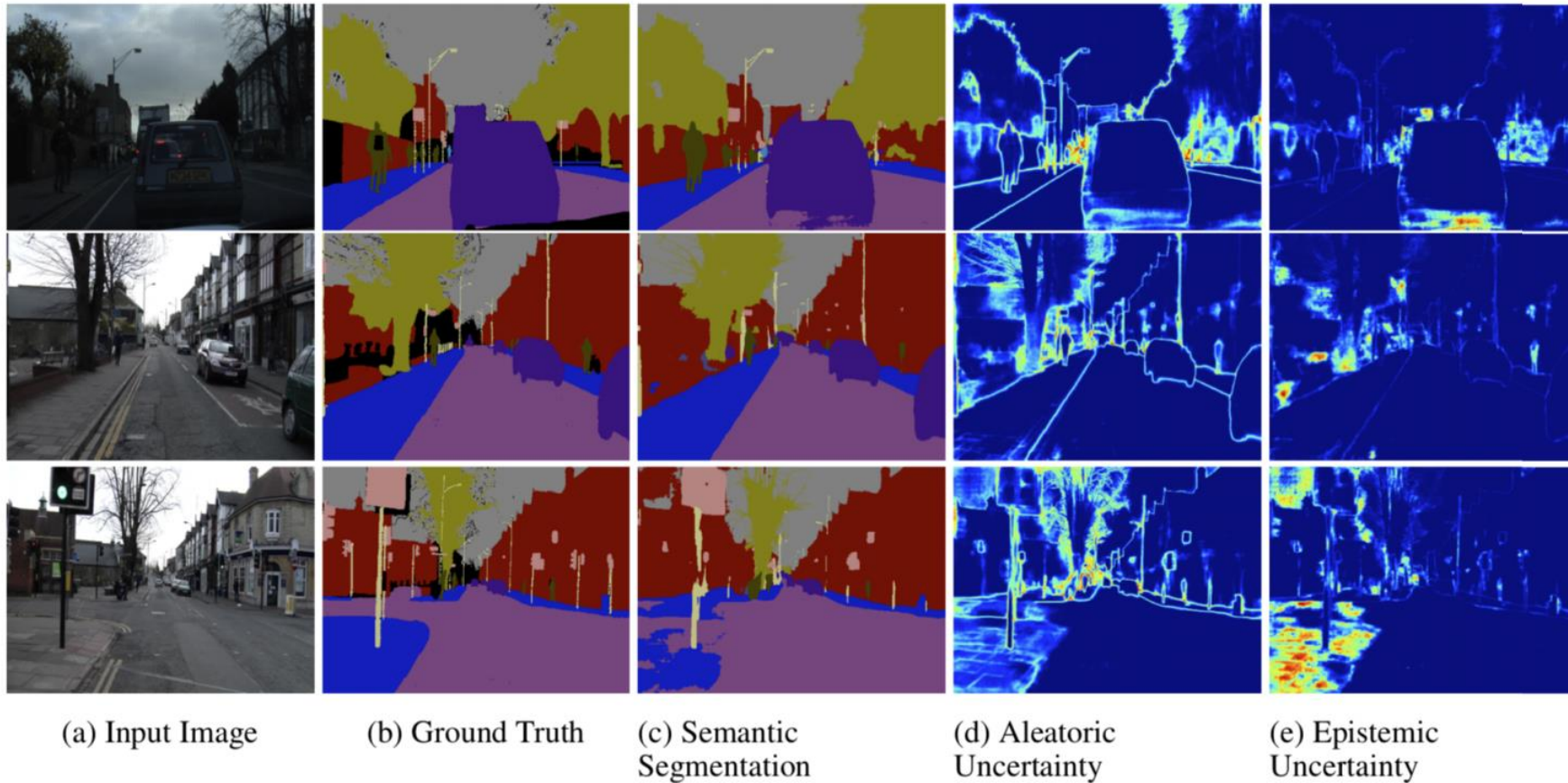
- Step 3: compute prediction with Monte Carlo approximations:

$$p(y^* | x^*, D) \approx \frac{1}{K} \sum_{k=1}^K p(y^* | x^*, \theta_k), \quad \underline{\theta_k \sim q(\theta)}$$

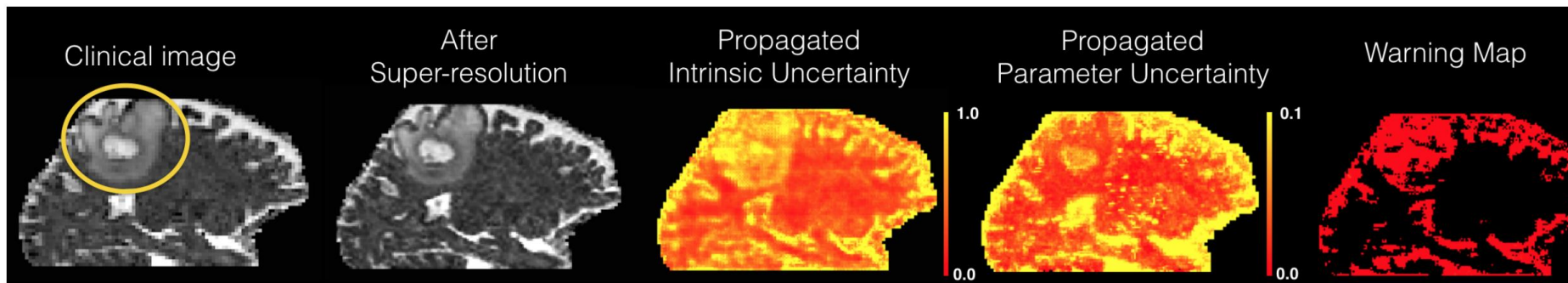
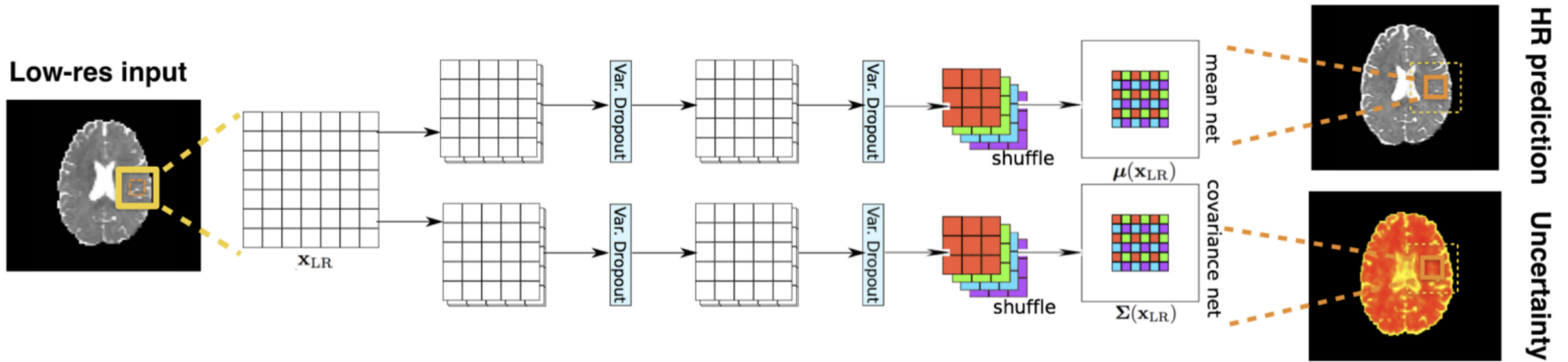
Mean-field Gaussian case:
 $\theta_k = m_\theta + \sigma_\theta \epsilon_k, \epsilon_k \sim N(0, I)$



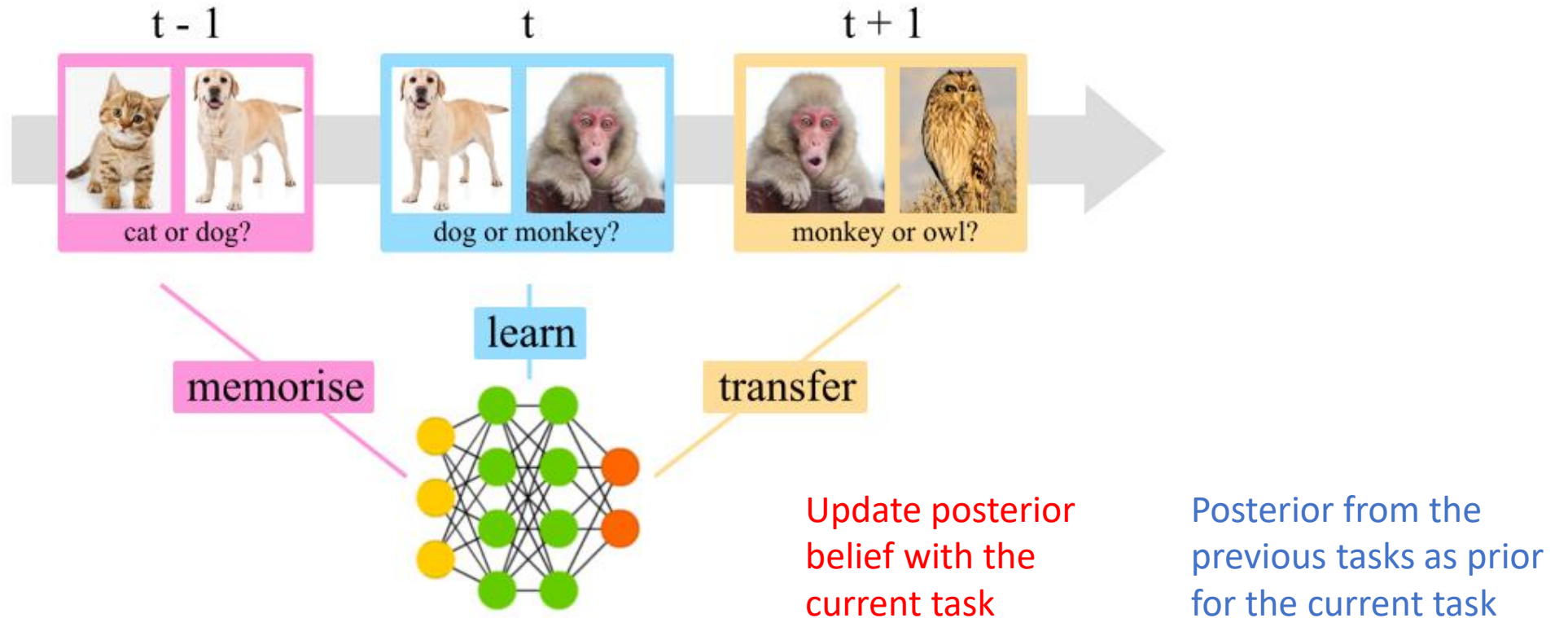
Applications of BNNs: Image Segmentation



Applications of BNNs: Super Resolution



Applications of BNNs: Continual Learning



$$L_{VCL}^t(q_t(\theta)) = E_{q_t(\theta)}[\log p(D_t | \theta)] - KL[q_t(\theta) || q_{t-1}(\theta)]$$

Recent Progress in BNNs: Inference

$$\text{SGD: } \theta_{t+1} = \theta_t - \eta \nabla_{\theta_t} \tilde{U}(\theta_t)$$

$$\text{SGLD: } \theta_{t+1} = \theta_t - \eta \nabla_{\theta_t} \tilde{U}(\theta_t) + \sqrt{2\eta} \epsilon, \quad \epsilon \sim N(0, I)$$

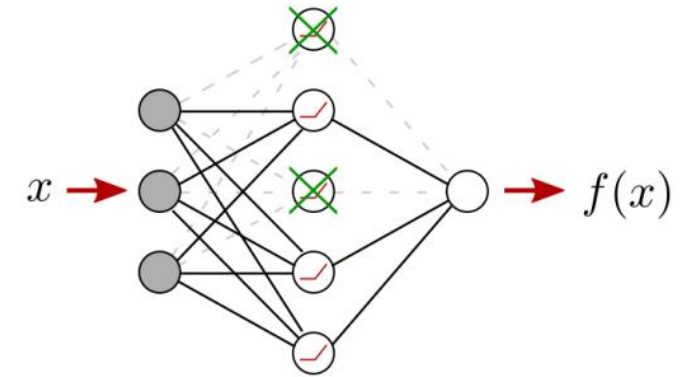
Stochastic gradient MCMC

Recent Progress in BNNs: Inference

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Stochastic gradient MCMC



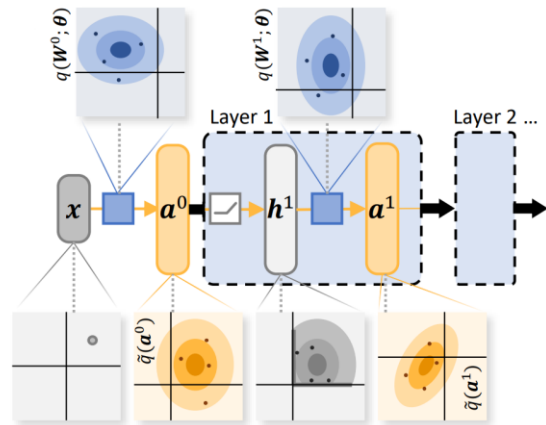
Monte Carlo dropout

Recent Progress in BNNs: Inference

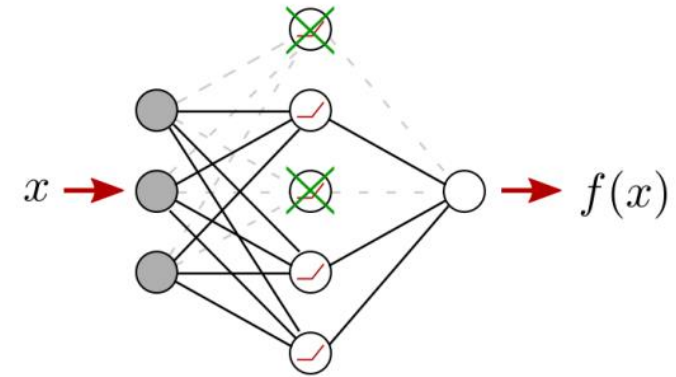
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Deterministic approximations



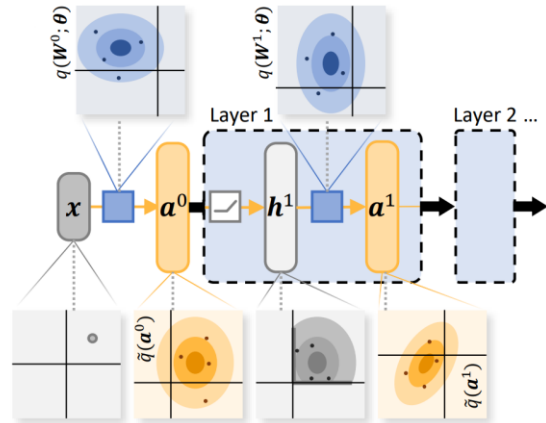
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Recent Progress in BNNs: Inference

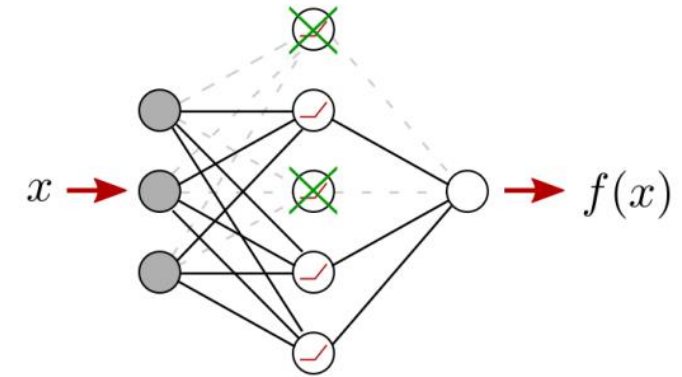
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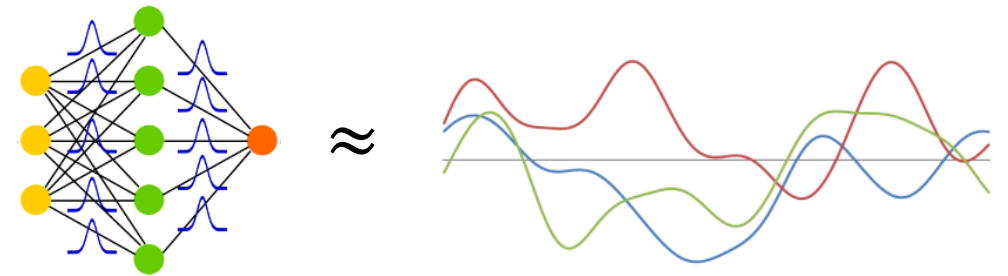
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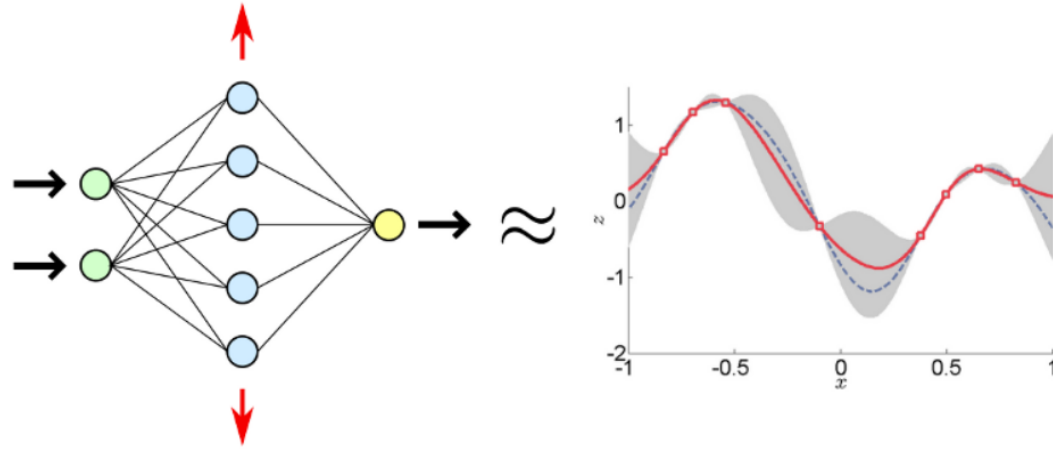


Monte Carlo dropout



Function space approximate inference

Recent Progress in BNNs: Theory



Connections to GPs:

- BNN with very wide hidden layers
 \approx Gaussian process
- Width limit convergence: in both
prior (Neal's result) and posterior

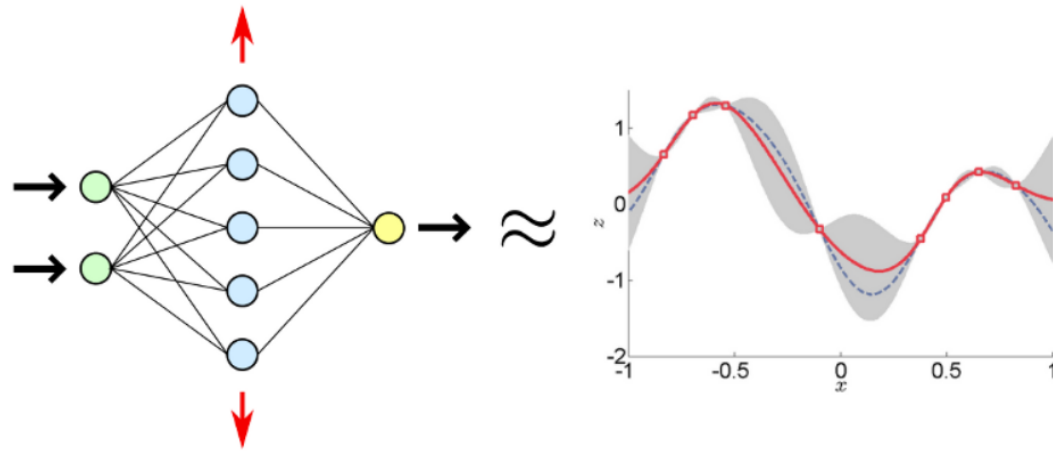
Neal. Bayesian Learning for Neural Networks. PhD Thesis, 1996

Matthews et al. Gaussian Process Behaviour in Wide Deep Neural Networks. ICLR 2018

Lee et al. Deep Neural Networks as Gaussian Processes. ICLR 2018

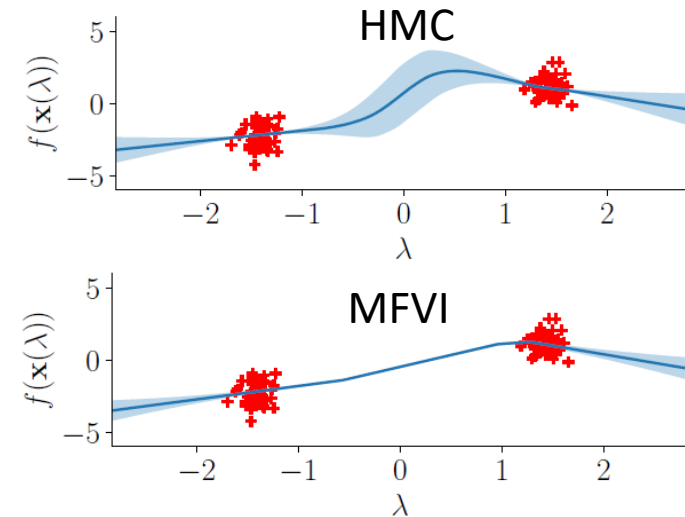
Hron et al. Exact posterior distributions of wide Bayesian neural networks. 2020

Recent Progress in BNNs: Theory



Connections to GPs:

- BNN with very wide hidden layers \approx Gaussian process
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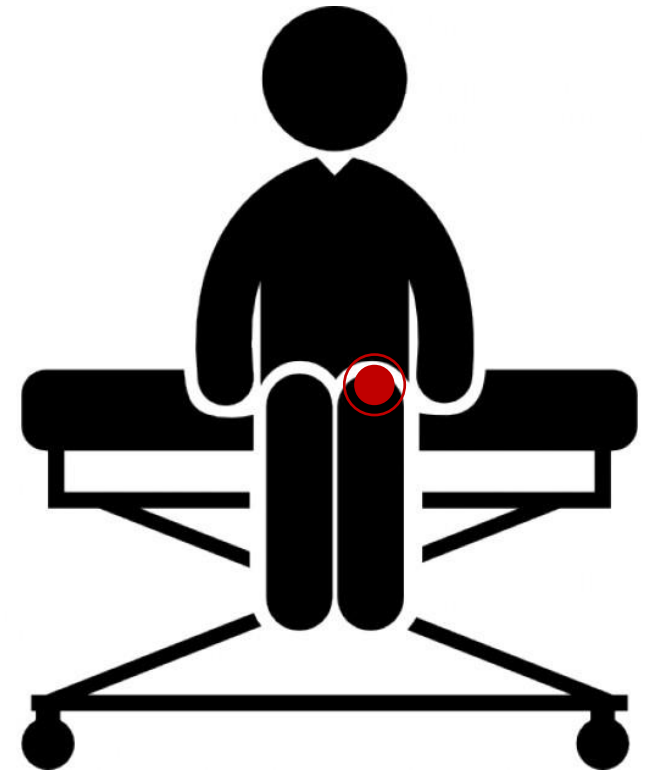


Approx. vs exact inference:

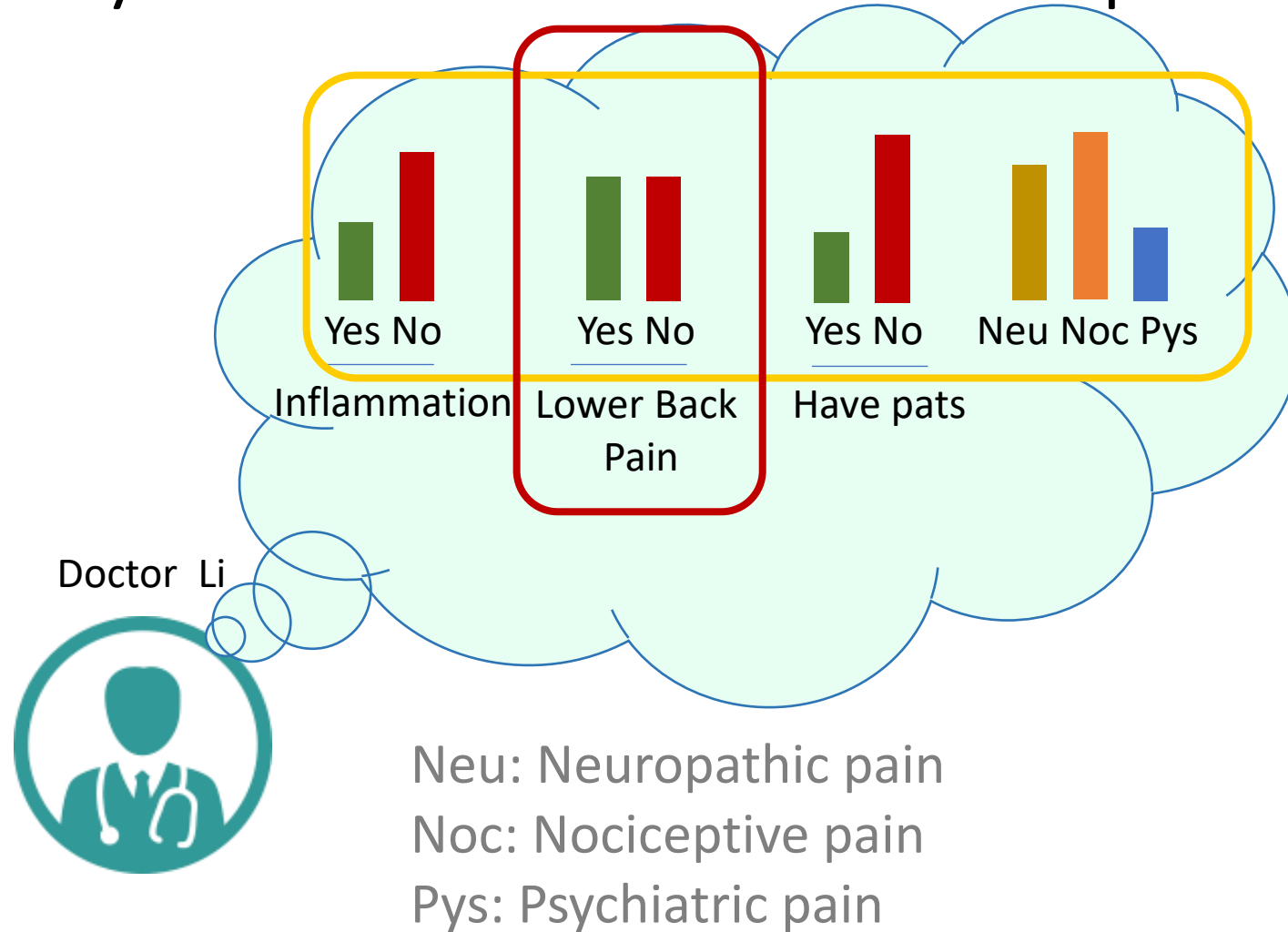
- Theoretical limitation of MFVI in shallow BNNs with ReLU activations
- Empirically deep BNNs with MFVI still fails in certain cases

Dynamic Information Acquisition

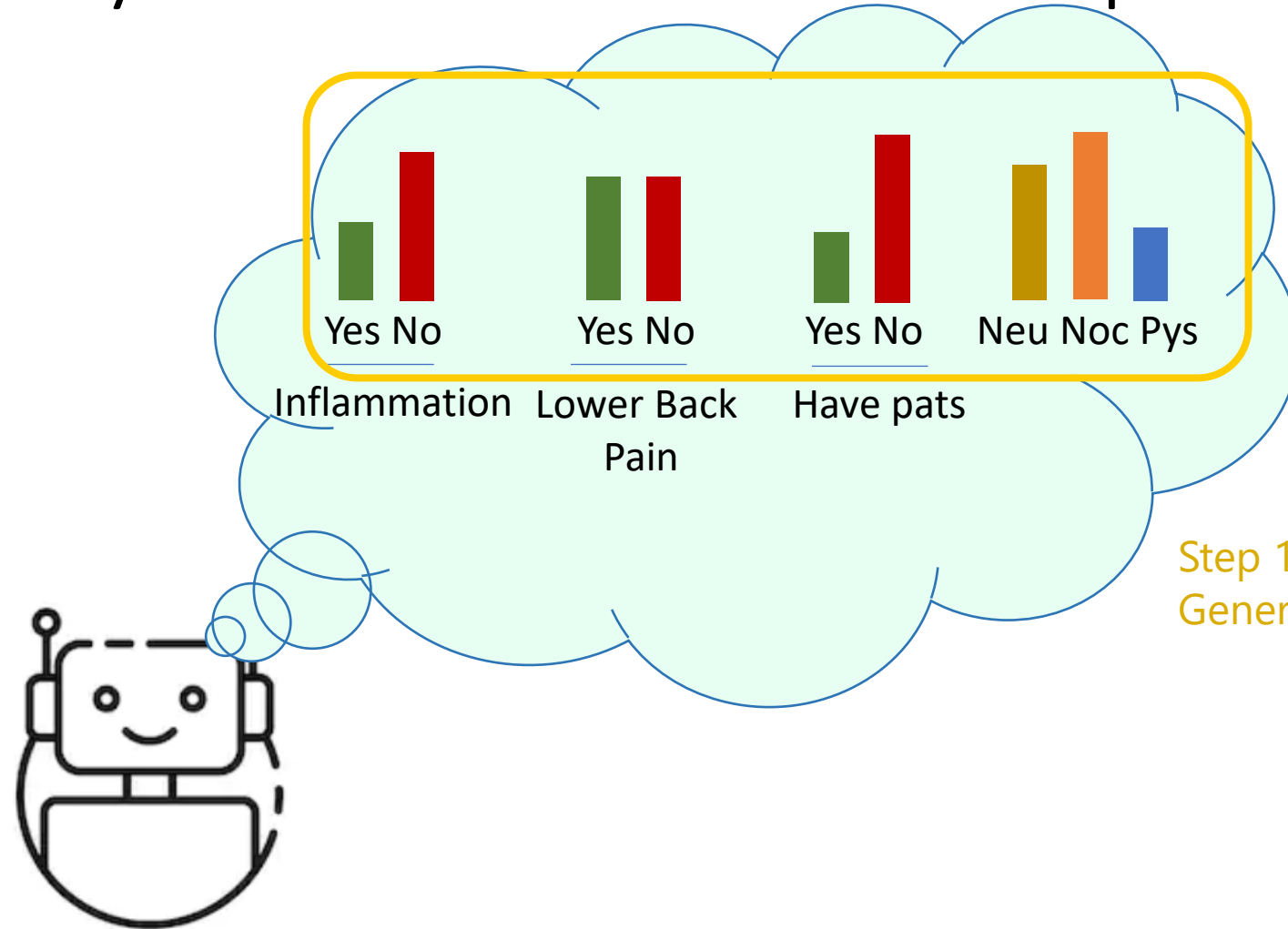
Doctor Li



Dynamic Information Acquisition

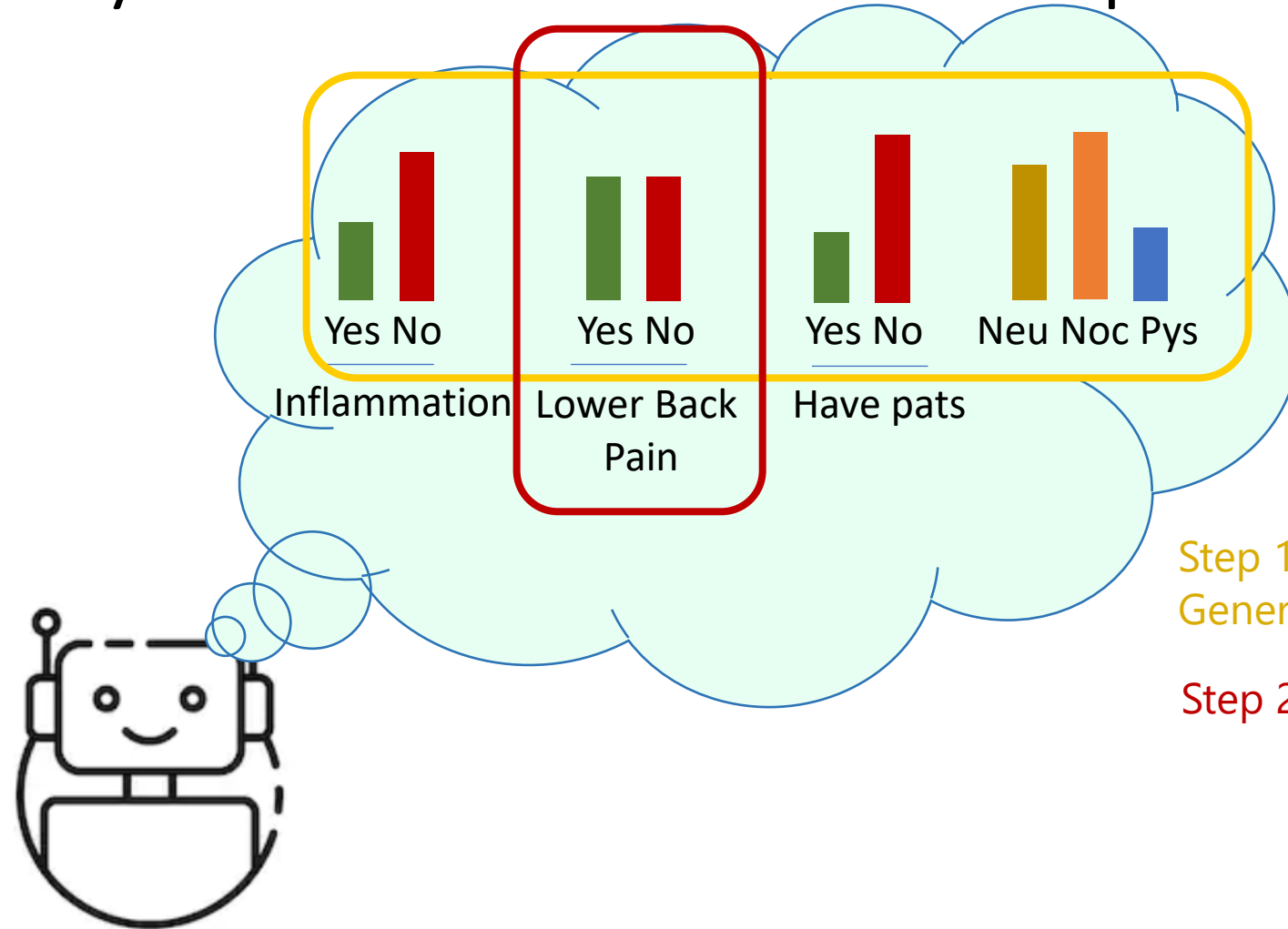


Dynamic Information Acquisition



Step 1: Missing Value Prediction with Deep Generative Model

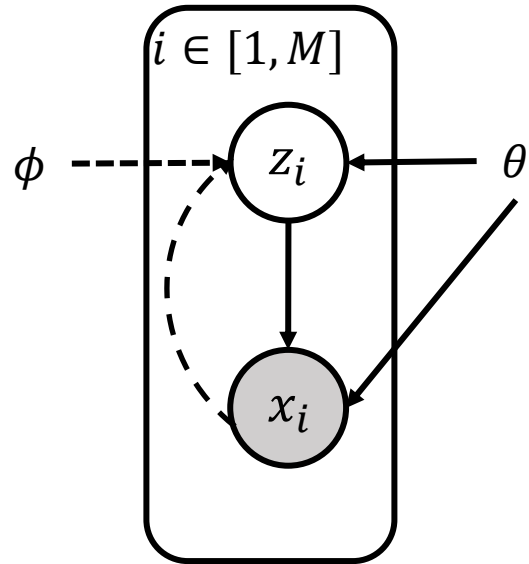
Dynamic Information Acquisition



Step 1: Missing Value Prediction with Deep Generative Model

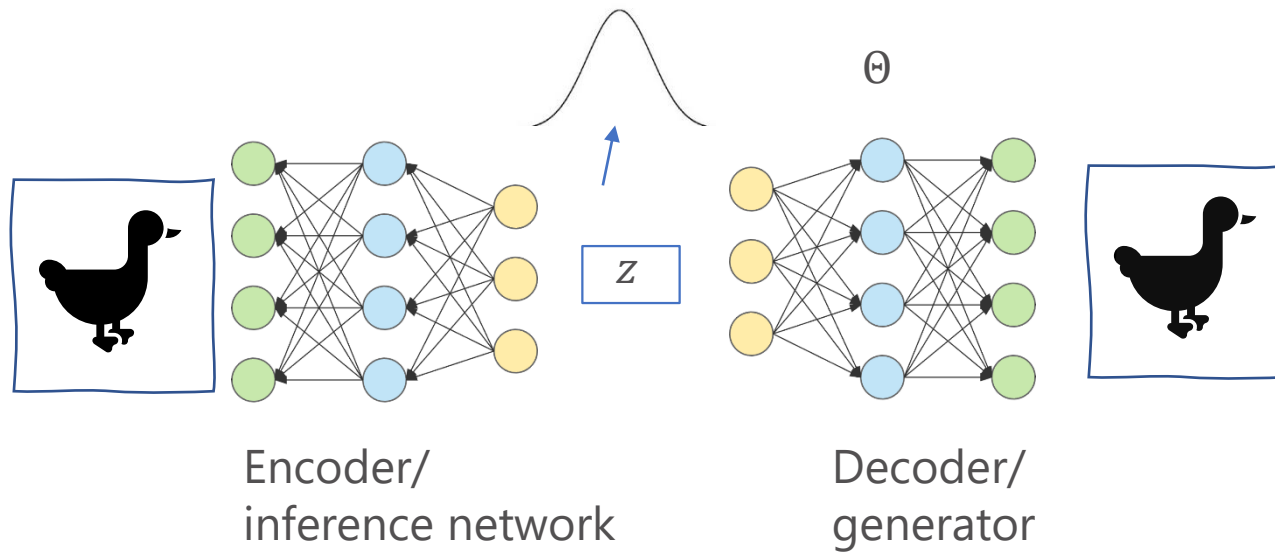
Step 2: Active element-wise information acquisition

A Deep Generative Model

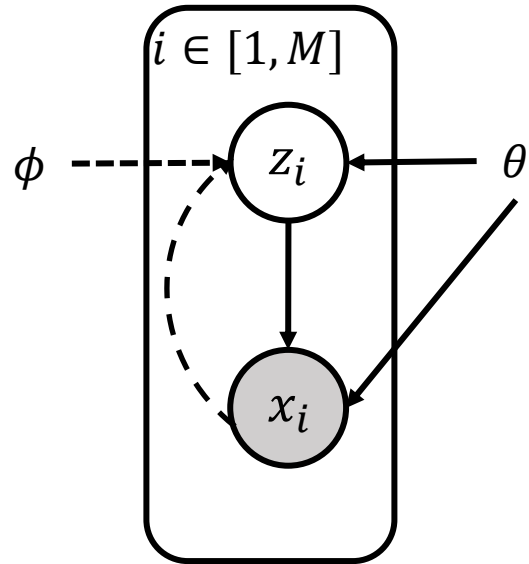


Variational Auto-encoder (VAE)

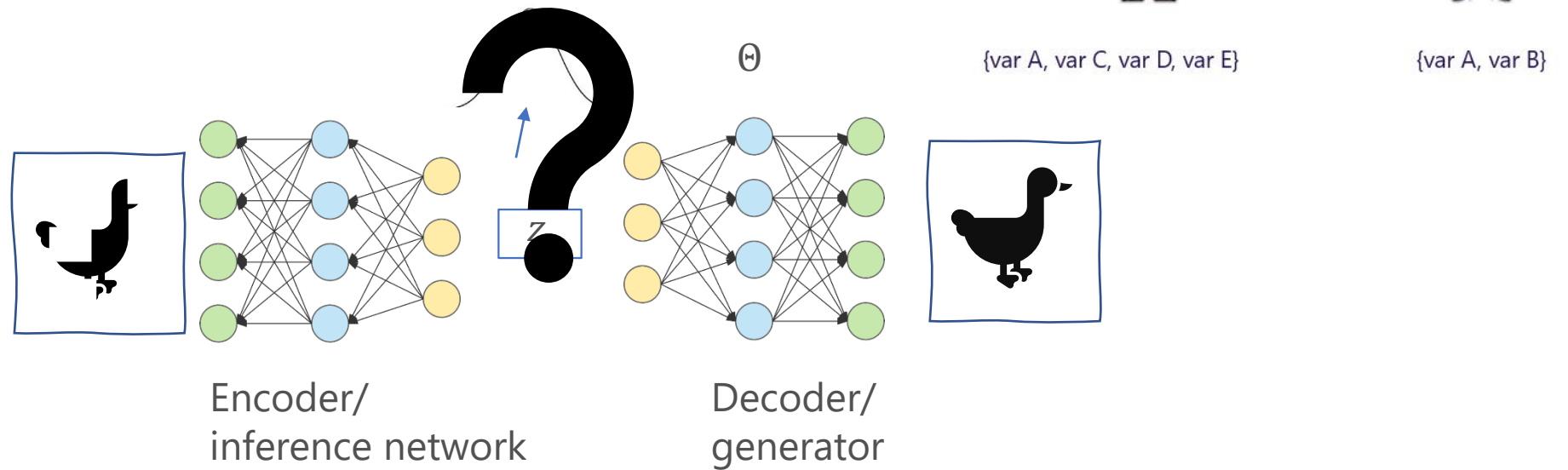
$$L_{amortized} = \log p(\mathbf{x}) - KL(q(\mathbf{z} | \mathbf{x}) | p(\mathbf{z} | \mathbf{x}))$$



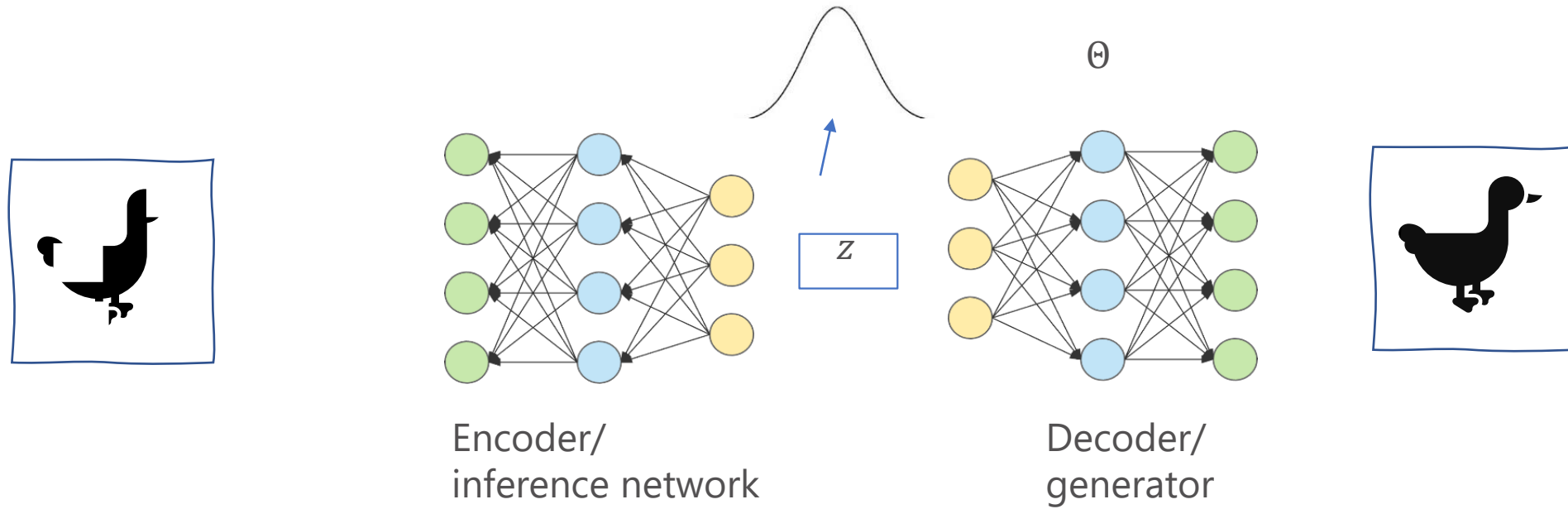
Challenges



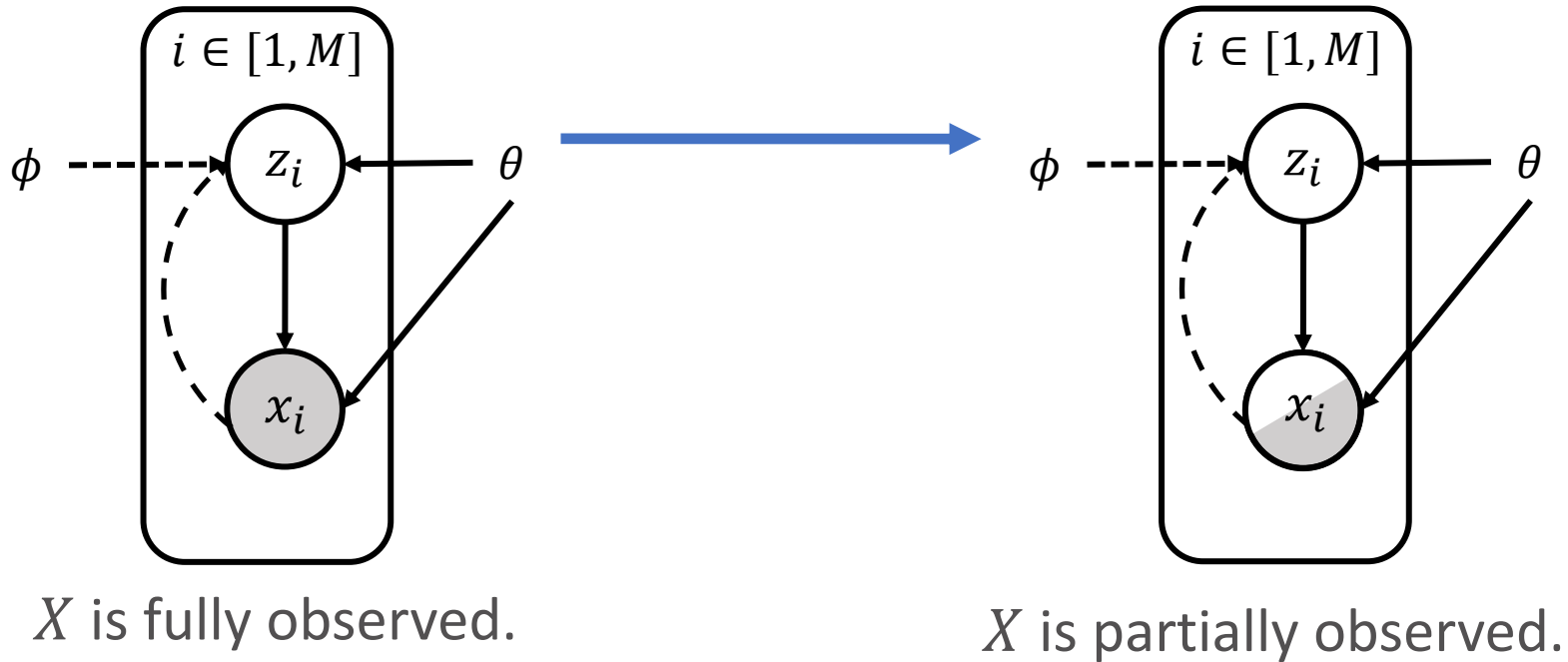
- Utilizing Deep Learning for scalable inference
- **Cannot handle partial observation**



VAE to Partial VAE

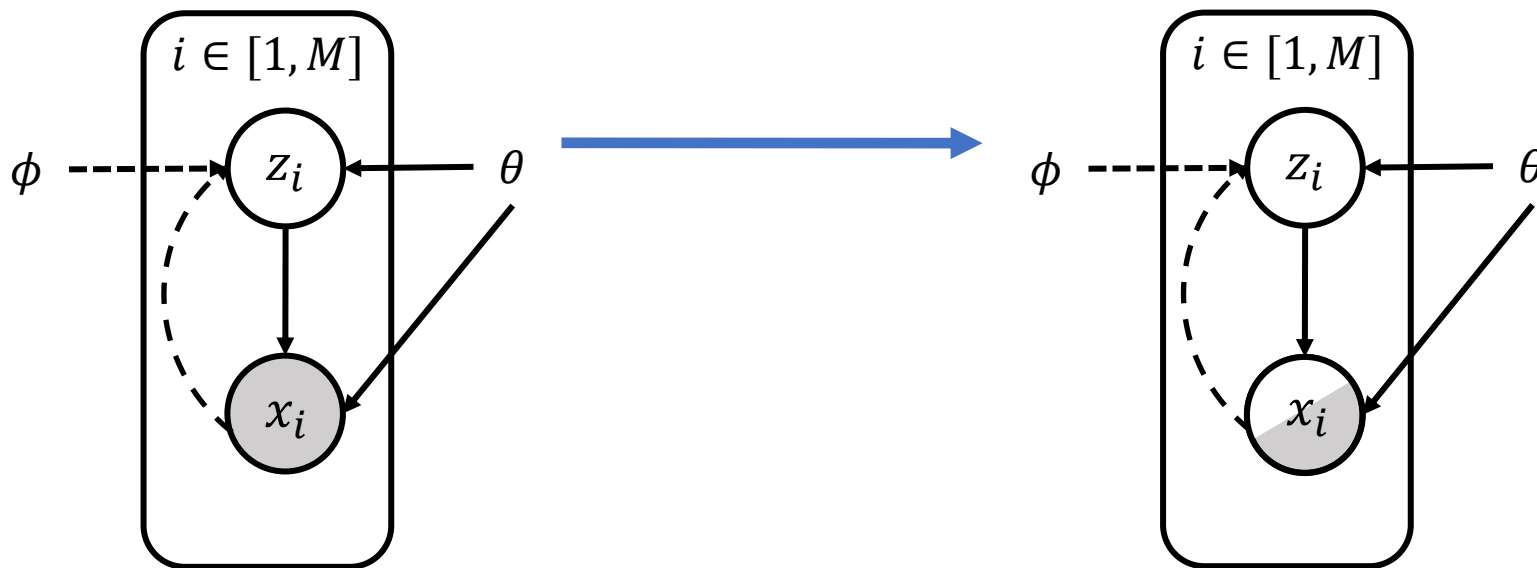


VAE to Partial VAE



We aim to infer the missing values X_U from the observed values X_O

VAE to Partial VAE



$$L_{amortized} = \log p(\mathbf{x}) - KL[q(\mathbf{z}|\mathbf{x})||p(\mathbf{z}|\mathbf{x})]$$

$$= E_{\mathbf{z} \sim q(\mathbf{z}|\mathbf{x})} [\log p_{\theta}(\mathbf{x}|\mathbf{z})] - KL[q(\mathbf{z}|\mathbf{x})||p(\mathbf{z})]$$

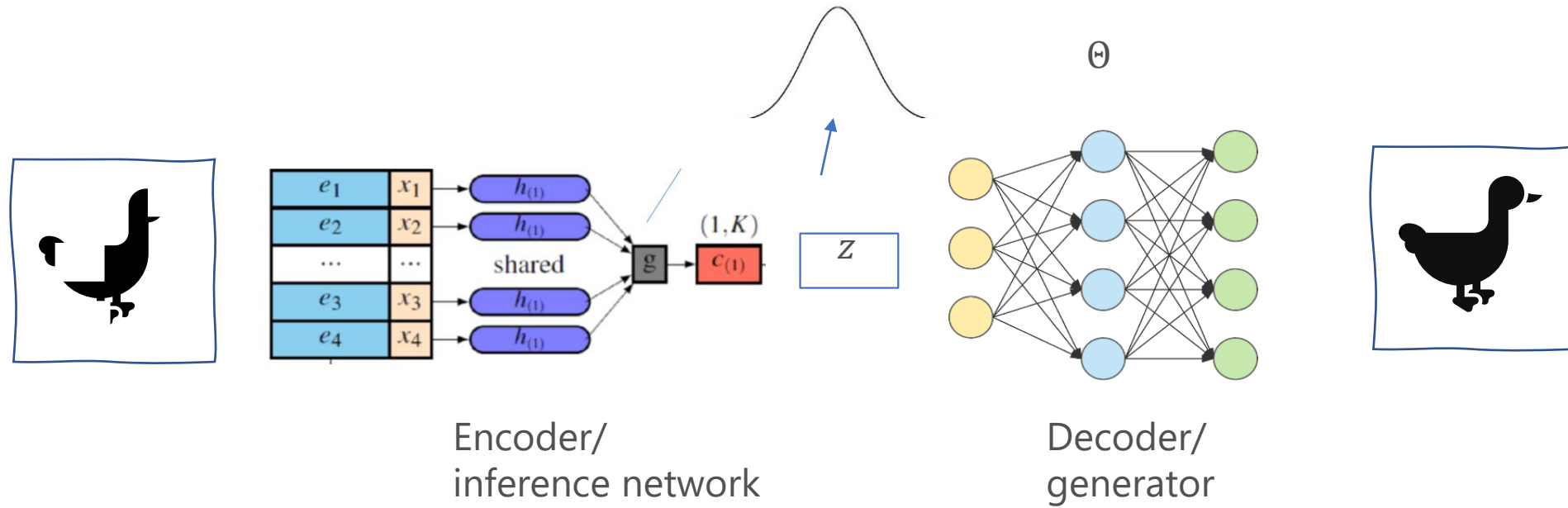
$$L_{amortized} = \log p(\mathbf{x}_o) - KL[q(\mathbf{z}|\mathbf{x}_o)||p(\mathbf{z}|\mathbf{x}_o)]$$

$$= E_{\mathbf{z} \sim q(\mathbf{z}|\mathbf{x}_o)} [\log p_{\theta}(\mathbf{x}|\mathbf{z})] - KL[q(\mathbf{z}|\mathbf{x}_o)||p(\mathbf{z})]$$

The ELBO still holds.

The challenge is how to design an inference net.

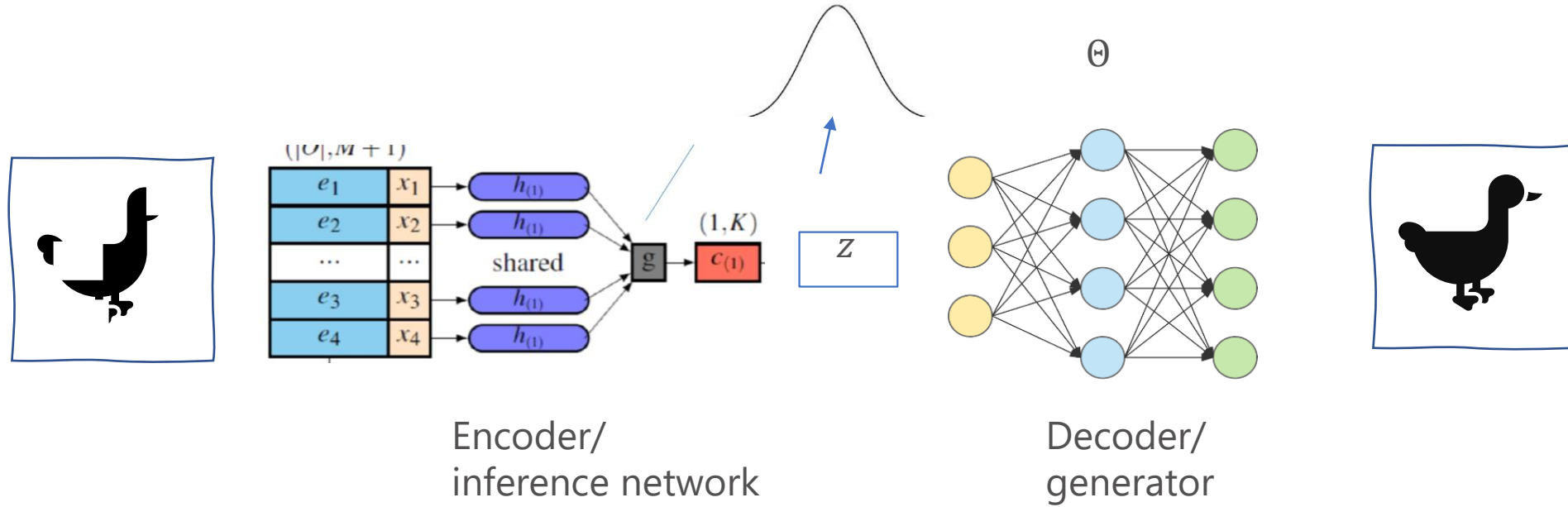
VAE to Partial VAE



$$\mathbf{c}(\mathbf{x}_O) := g(\vec{h}(\mathbf{s}_1), \vec{h}(\mathbf{s}_2), \dots, \vec{h}(\mathbf{s}_O))$$

Set Encoder

VAE to Partial VAE

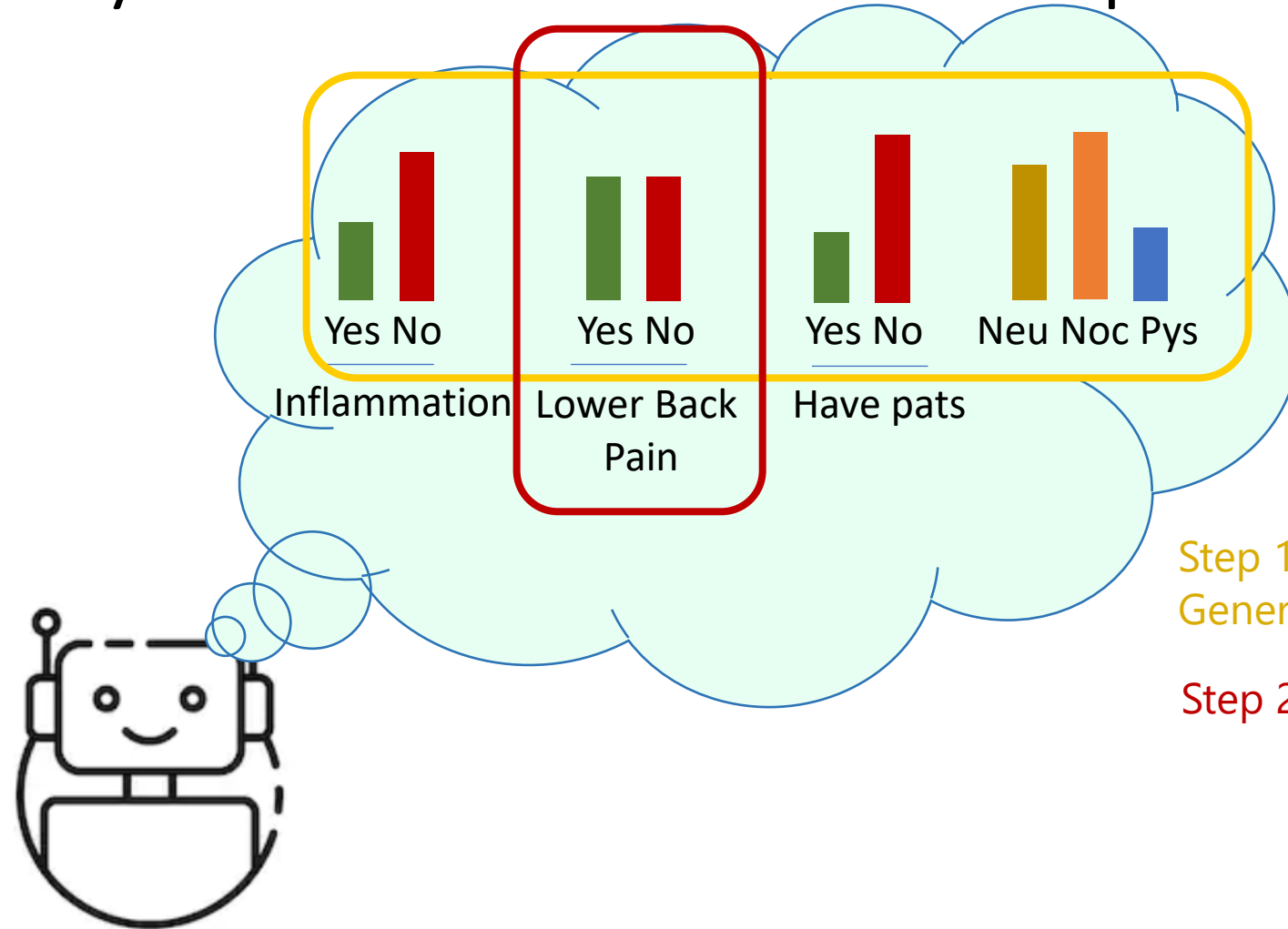


$$\mathbf{c}(\mathbf{x}_O) := g(\vec{h}(s_1), \vec{h}(s_2), \dots, \vec{h}(s_O))$$

Set Encoder

	A	B	C	D	E	F	
1							
2		cough	cold	temperature	marlaria	age	...
3	Alice	1	1			25	
4	Mary				0		
5	Kevin	0				47	

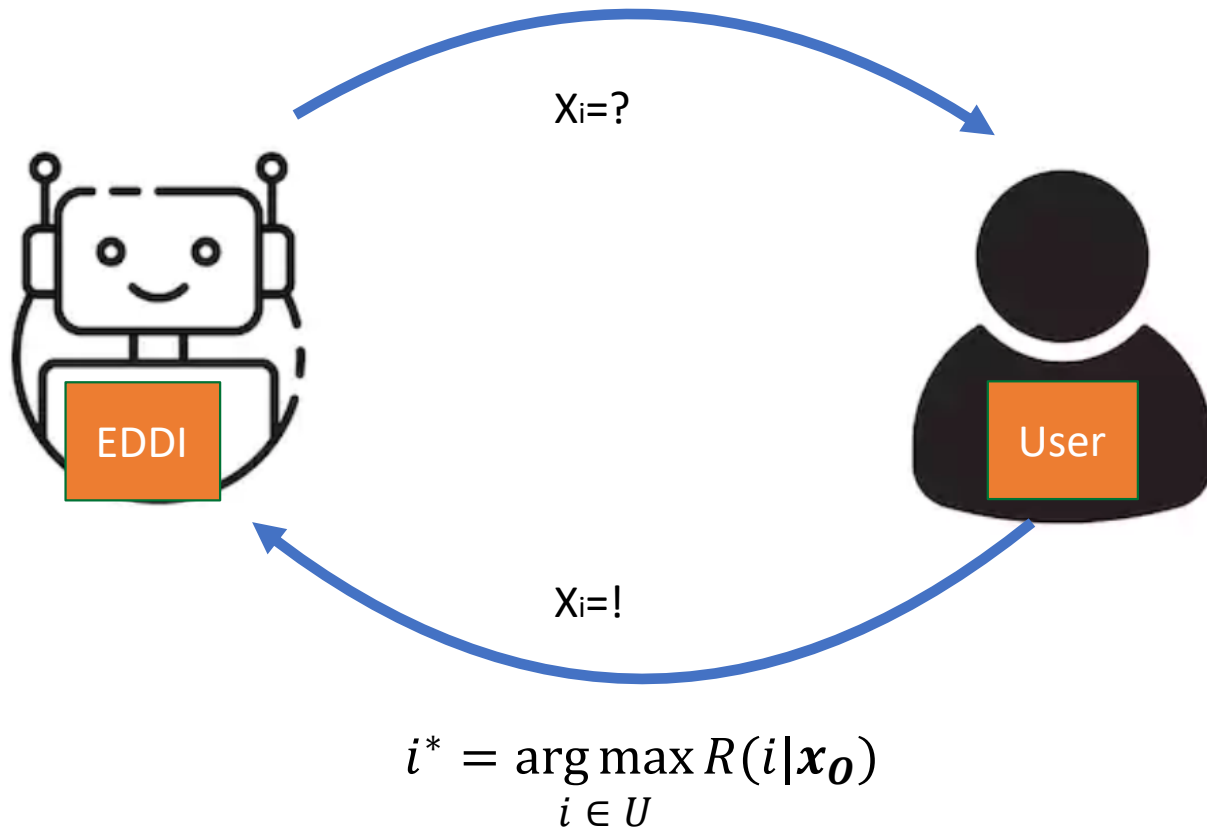
Dynamic Information Acquisition



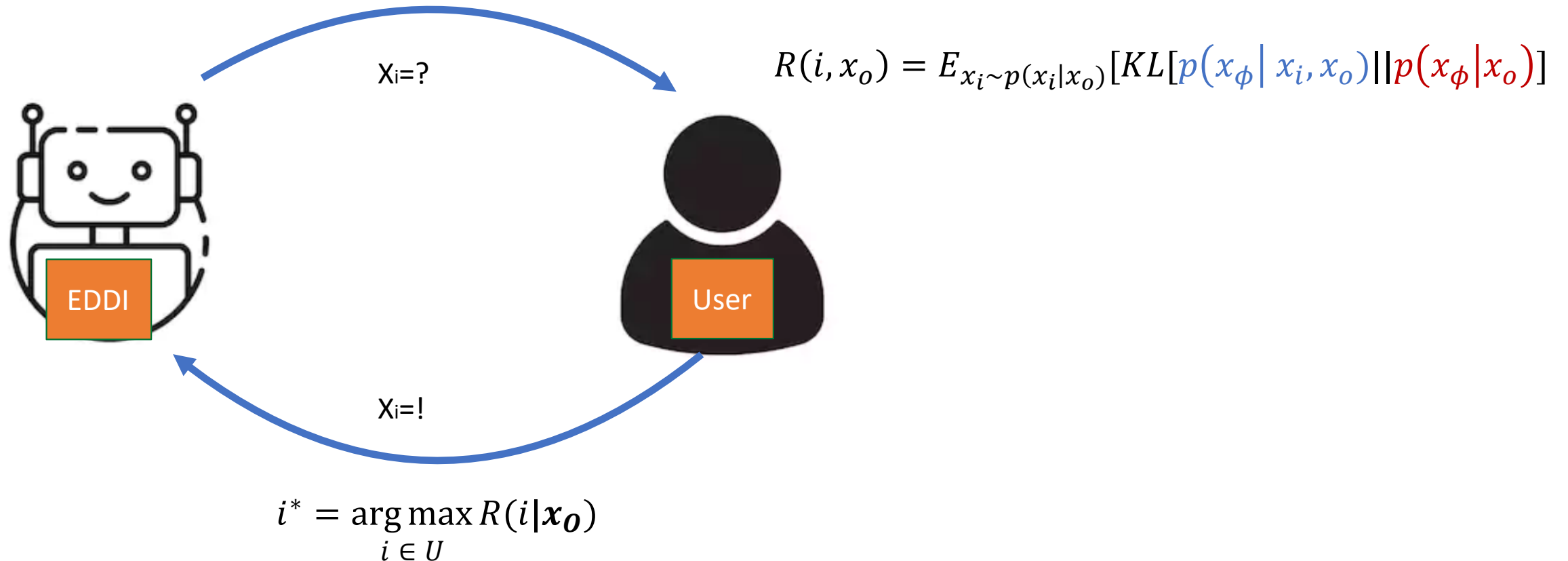
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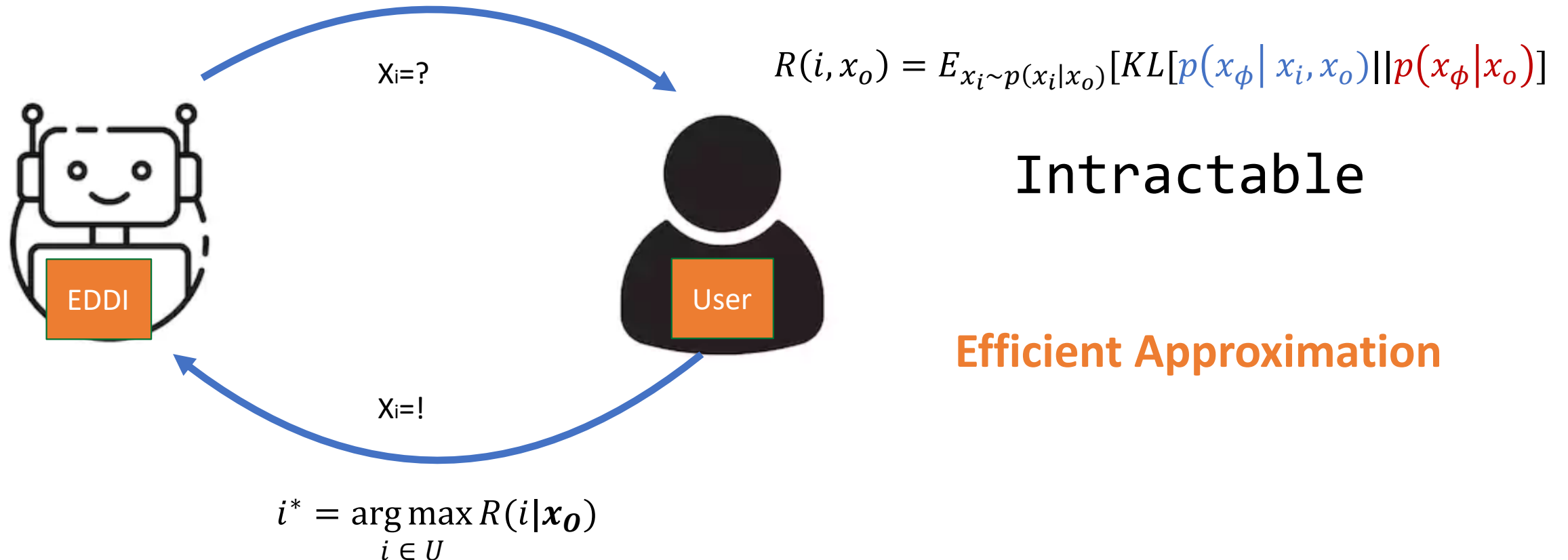
Actively Select the Next Variable



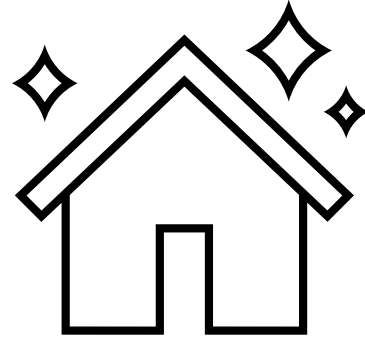
Actively Select the Next Variable



Actively Select the Next Variable



Predicting a House's Price



Kim (customer)



How far away is the
supermarket?



Ty (broker)

Our solution

Baseline

Model Questions

TARGET VARIABLE
median value of owner-occupied homes in \$1000's

QUESTION 1
index of accessibility to radial highways

1.00 24.00 ✓ 5.00

1.0000

full-value property-tax rate per \$10,000 0.9822

proportion of non-retail business acres per town 0.7785

pupil-teacher ratio by town 0.4598

$1000(B - 0.63)^2$ where B is the proportion of students by town 0.4401

% lower status of the population 0.1252

average number of rooms per dwelling 0.0172

per capita crime rate by town 0.0195

proportion of residential land zoned for lots over 25,000 sq.ft 0.0000

charles river dummy variable (true if tract bounds river false otherwise) 0.0000

nitrous oxides concentration (parts per 10 million) 0.0000

proportion of owner-occupied units built prior to 1940 0.0000

weighted distances to five Boston employment centres 0.0000

Target Variable

median value of owner-occupied homes in \$1000's

36.55

Random Questions

TARGET VARIABLE
median value of owner-occupied homes in \$1000's

QUESTION 1
weighted distances to five Boston employment centres

1.13 12.13 ✓ 2.65

0.4422

pupil-teacher ratio by town 0.4717

proportion of residential land zoned for lots over 25,000 sq.ft 0.0040

proportion of owner-occupied units built prior to 1940 0.3899

proportion of non-retail business acres per town 1.0000

per capita crime rate by town 0.1247

nitrous oxides concentration (parts per 10 million) 0.2050

index of accessibility to radial highways 0.8514

full-value property-tax rate per \$10,000 0.4286

charles river dummy variable (true if tract bounds river false otherwise) 0.1289

average number of rooms per dwelling 0.0000

$1000(B - 0.63)^2$ where B is the proportion of students by town 0.8540

% lower status of the population 0.0000

Target Variable

median value of owner-occupied homes in \$1000's

40.66

Our solution

Baseline

Model Questions Target Variable Random Questions Target Variable

TARGET VARIABLE
median value of owner-occupied homes in \$1000's

QUESTION 1
index of accessibility to radial highways



full-value property-tax rate per \$10,000 0.9822

proportion of non-retail business acres per town 0.7765

pupil-teacher ratio by town 0.4356

1000(B - 0.63)^2 where B is the proportion of students by town 0.4401

% lower status of the population 0.1237

average number of rooms per dwelling 0.0172

per capita crime rate by town 0.0155

proportion of residential land zoned for lots over 25,000 sq.ft. 0.0000

charles river dummy variable (true if tract bounds river false otherwise) 0.0000

nitrous oxides concentration (parts per 10 million) 0.0000

proportion of owner-occupied units built prior to 1940 0.0000

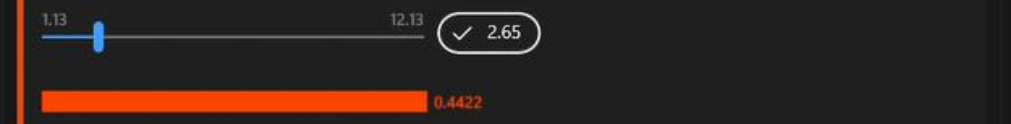
weighted distances to five Boston employment centres 0.0000

Target Variable



TARGET VARIABLE
median value of owner-occupied homes in \$1000's

QUESTION 1
weighted distances to five Boston employment centres



pupil-teacher ratio by town 0.4711

proportion of residential land zoned for lots over 25,000 sq.ft. 0.0000

propo 0.0000

propo 0.0000

per ca 0.1247

nitrous oxides concentration (parts per 10 million) 0.2090

index of accessibility to radial highways 0.8334

full-value property-tax rate per \$10,000 0.6298

charles river dummy variable (true if tract bounds river false otherwise) 0.1289

average number of rooms per dwelling 0.0000

1000(B - 0.63)^2 where B is the proportion of students by town 0.8549

% lower status of the population 0.0000

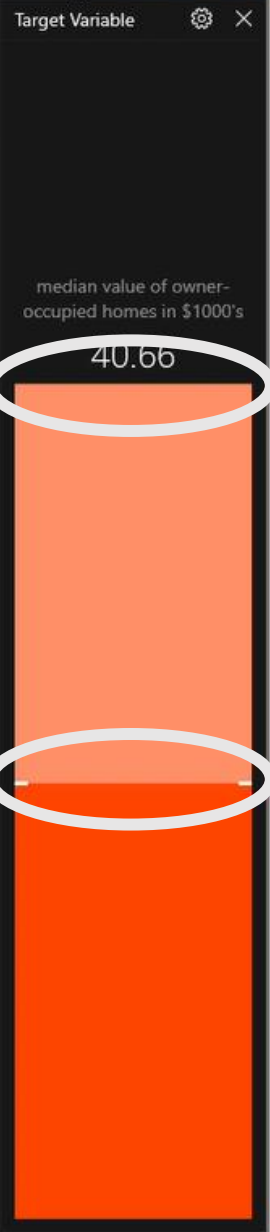
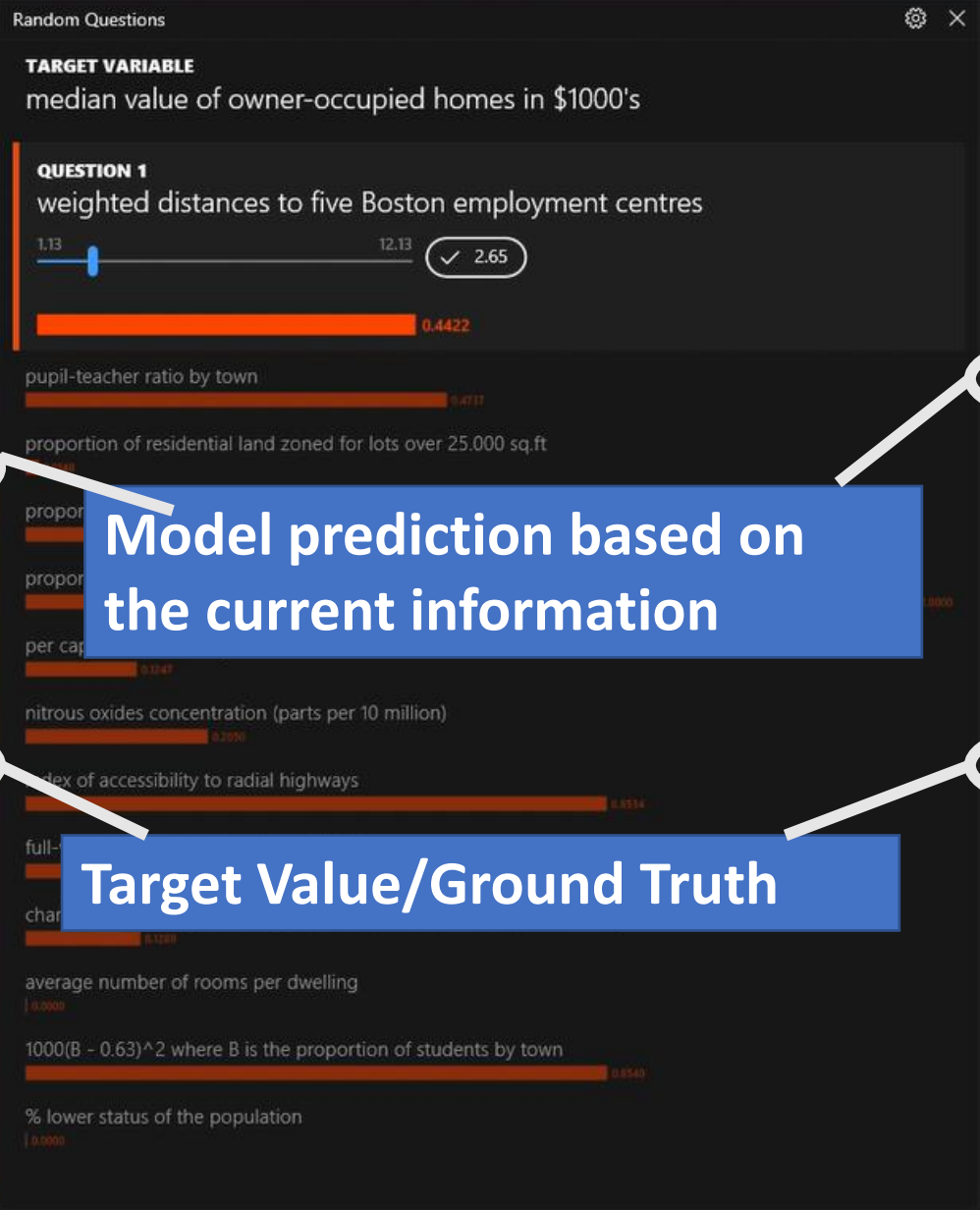
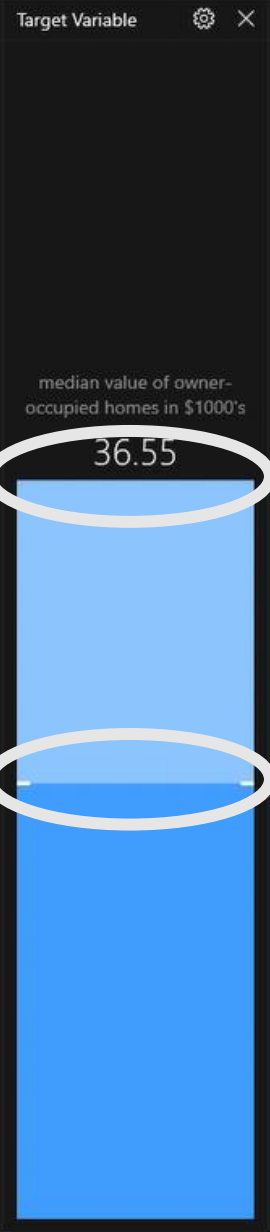
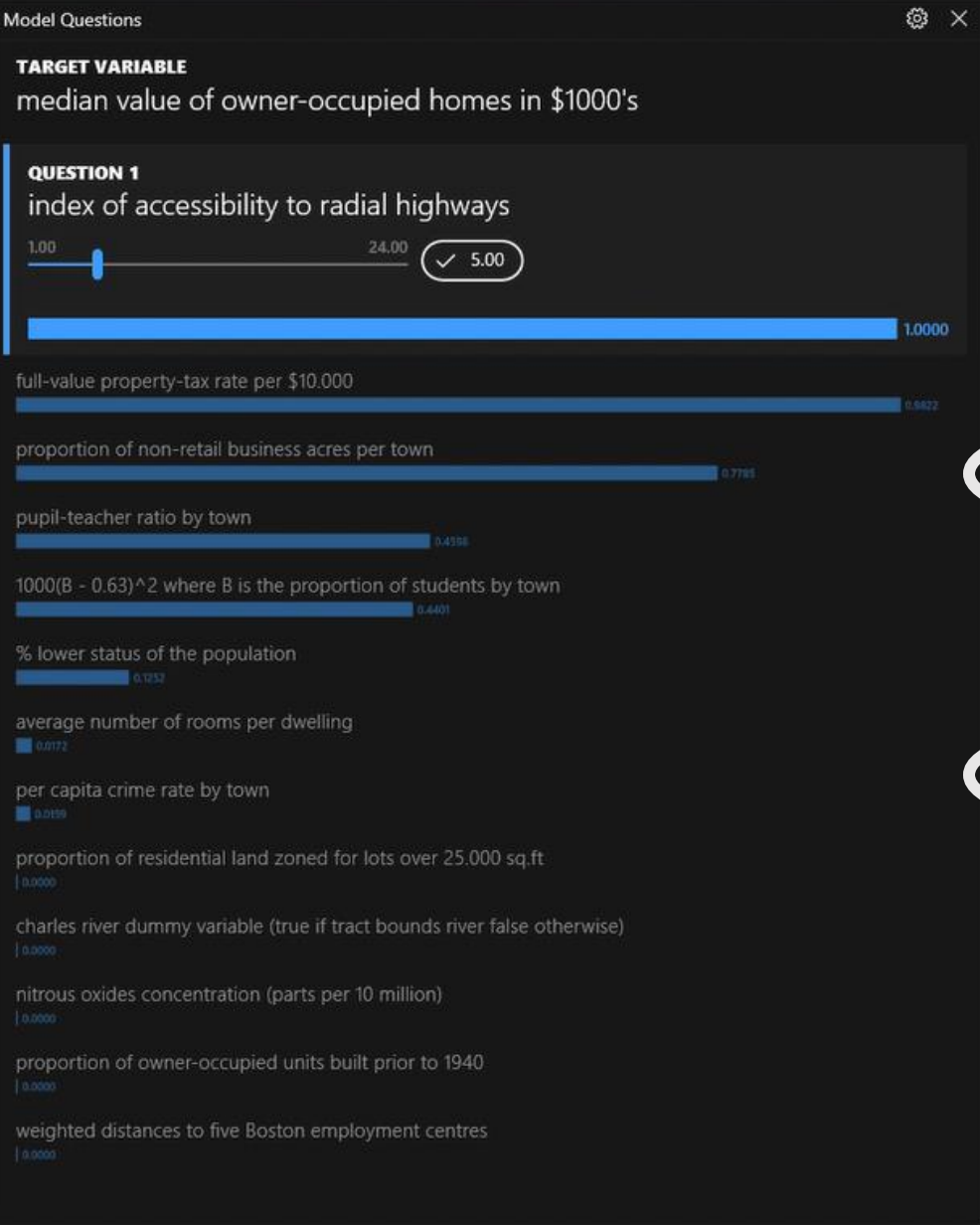
Target Variable



Model prediction based on the current information

Our solution

Baseline



Model prediction based on the current information

Target Value/Ground Truth

Our solution

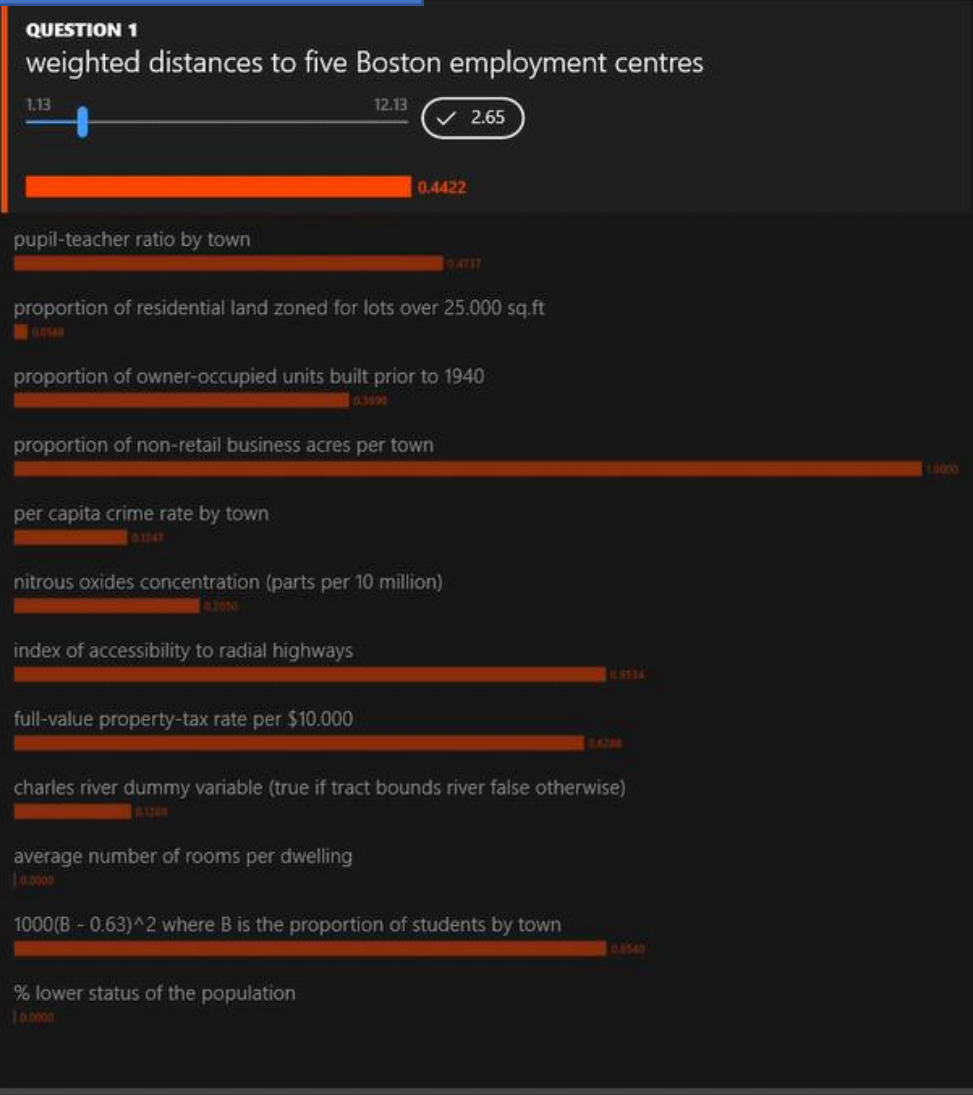
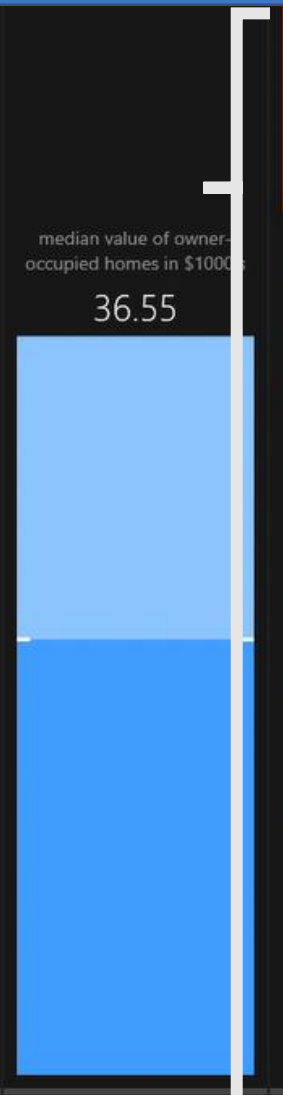
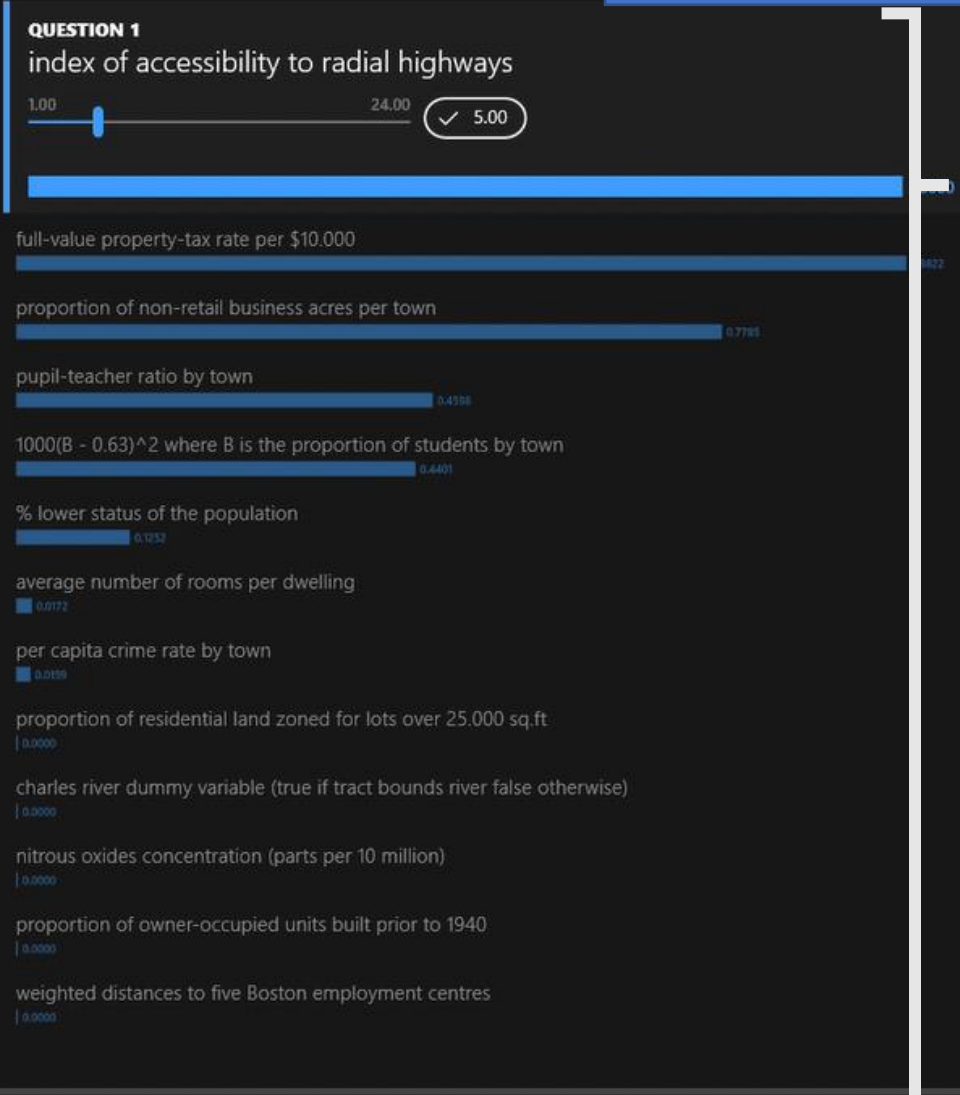
Baseline

List of questions that we could ask

Model Questions Target Variable Random Questions Target Variable

TARGET VARIABLE median value of owner-occupied homes in \$1000's

homes in \$1000's median value of owner-occupied homes in \$1000's



TARGET VARIABLE
median value of owner-occupied homes in \$1000's

QUESTION 1
index of accessibility to radial highways

1.00 24.00 ✓ 5.00

1.0000

full-value property-tax rate per \$10,000 0.5822

proportion of non-retail business acres per town 0.7785

pupil-teacher ratio by town 0.4398

$1000(B - 0.63)^2$ where B is the proportion of students by town 0.4401

% lower status of the population 0.1921

average number of rooms per dwelling 0.0172

per capita crime rate by town

charles river dummy variable (true if tract bounds river false otherwise) 0.0000

nitrous oxides concentration (parts per 10 million) 0.0000

proportion of owner-occupied units built prior to 1940 0.0000

weighted distances to five Boston employment centres 0.0000

List of questions that we could ask

TARGET VARIABLE
median value of owner-occupied homes in \$1000's

QUESTION 1
weighted distances to five Boston employment centres

1.13 12.13 ✓ 2.65

0.4422

pupil-teacher ratio by town 0.4737

proportion of residential land zoned for lots over 25,000 sq.ft 0.0148

proportion of owner-occupied units built prior to 1940 0.3898

proportion of non-retail business acres per town 1.0000

per capita crime rate by town 0.1847

nitrous oxides concentration (parts per 10 million) 0.2990

index of accessibility to radial highways 0.8534

full-value property-tax rate per \$10,000 0.6288

charles river dummy variable (true if tract bounds river false otherwise) 0.1288

average number of rooms per dwelling 0.0000

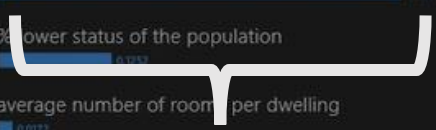
$1000(B - 0.63)^2$ where B is the proportion of students by town 0.8540

% lower status of the population 0.0000

median value of owner-occupied homes in \$1000's

40.66

Information reward



MODEL

Boston - Media Home Value

Refresh

GROUND TRUTH

C:\Users\mgrayson\Desktop\min-data-ai\boston\test.csv

Browse...

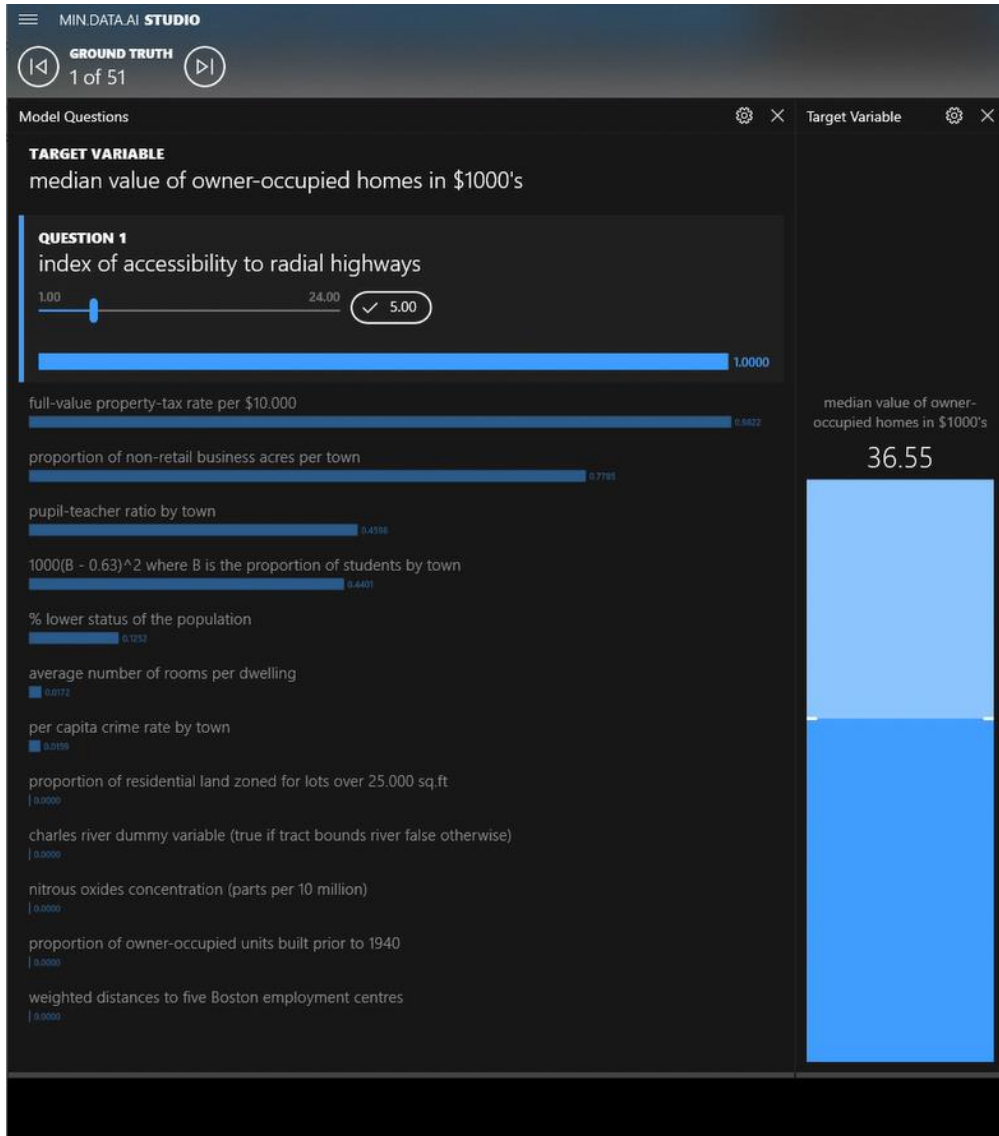
PANELS

✓ Model questions

✕ Random questions

GROUND TRUTH
1 of 51

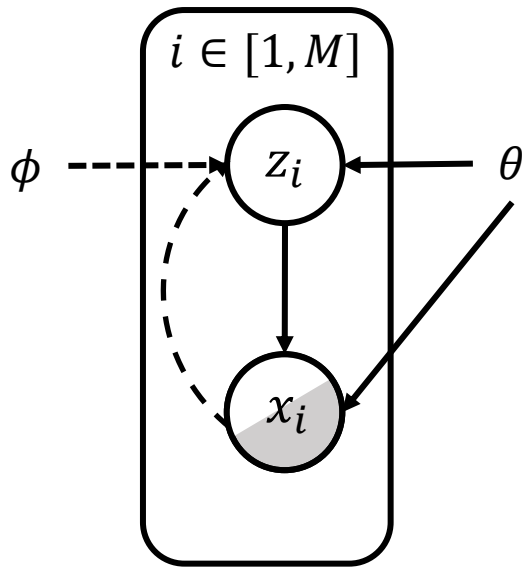




- List of questions from survey (e.g. US veteran) for mental health monitoring
- List of technical questions from interview for recruiting
- List of medical tests for diagnosis
-

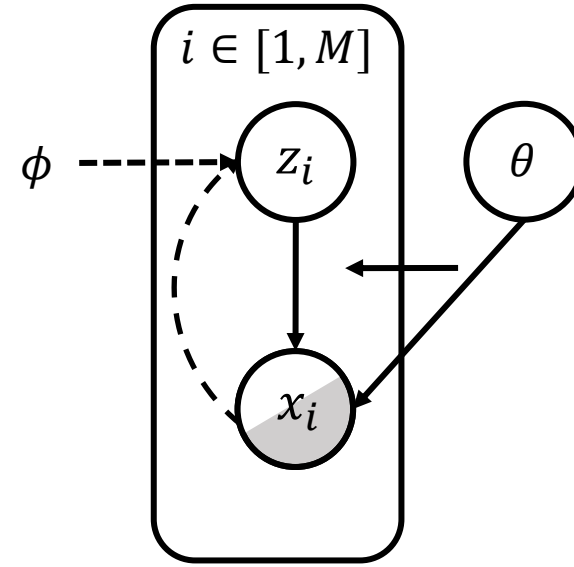
Does it work when there is few training data?

Partial Amortized Bayesian Deep Latent Gaussian Model (PA-BELGAM)



- Point estimate of global parameter θ

$$p(\mathbf{x}_o, \mathbf{z}) = \prod_{i=1}^N \prod_{d \in O_i} p(x_{i,d} | z_i) p(z_i)$$



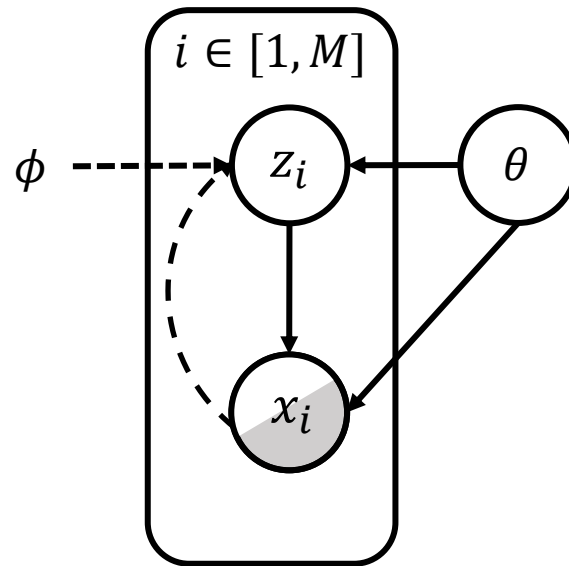
- Stochastic variable θ

$$p(\mathbf{x}_o, \theta, \mathbf{z}) = p(\theta) \prod_{i=1}^N \prod_{d \in O_i} p(x_{i,d} | z_i, \theta) p(z_i)$$

Partial Amortized Bayesian Deep Latent Gaussian Model (PA-BELGAM)

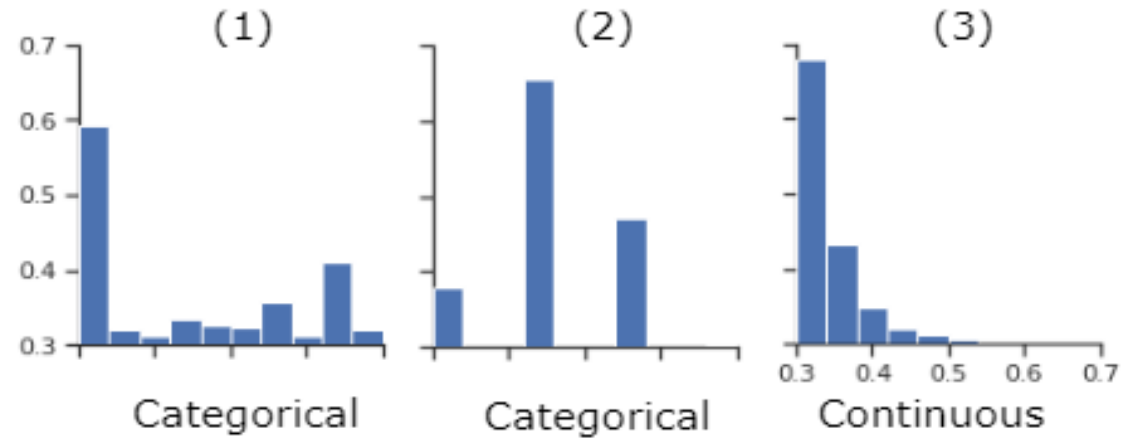
Amortized inference
for local latent
variables

SGHMC for global
latent variables

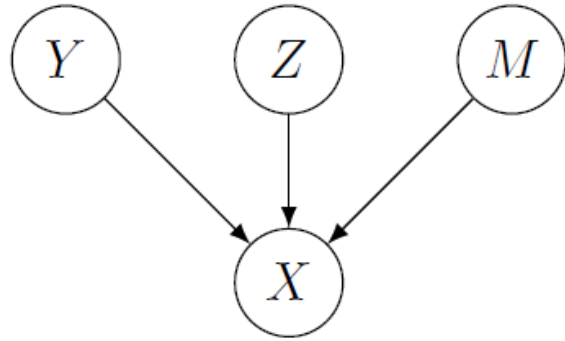


$$q(\theta, \mathbf{z} | \mathbf{x}_o) \approx q(\theta | \mathbf{x}_o) q_\phi(\mathbf{z} | \mathbf{x}_o)$$

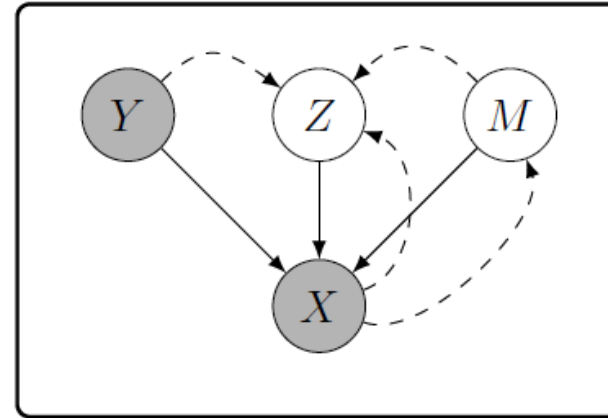
When Data are Heterogenous



When Robustness is Needed



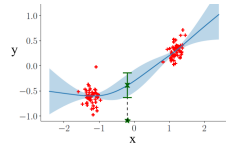
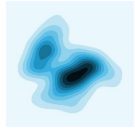
Causal Reasoning



Deep Causal
Manipulation
Augmented Model

Summary

The probability distribution $\pi(\theta)$ is intractable

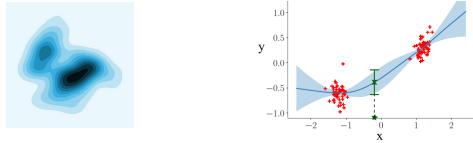


Approximate Inference

$$\int F(\theta) \pi(\theta) d\theta$$

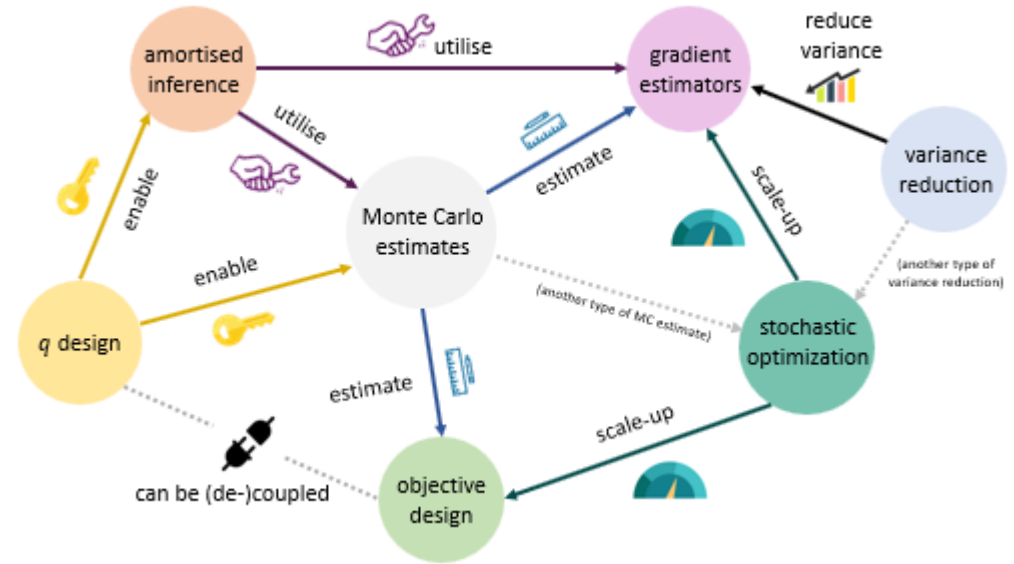
Summary

The probability distribution $\pi(\theta)$ is intractable



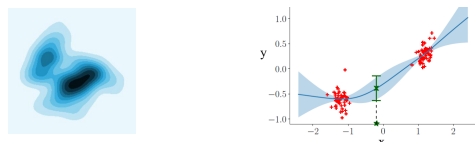
Approximate Inference

$$\int F(\theta) \pi(\theta) d\theta$$



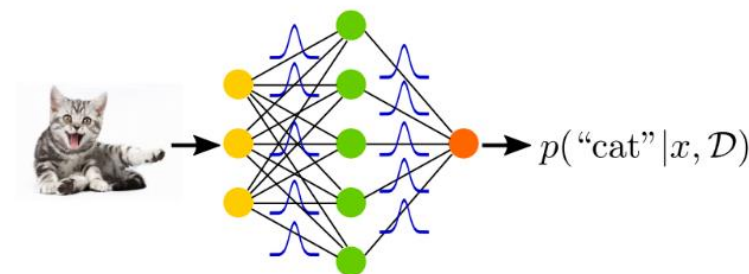
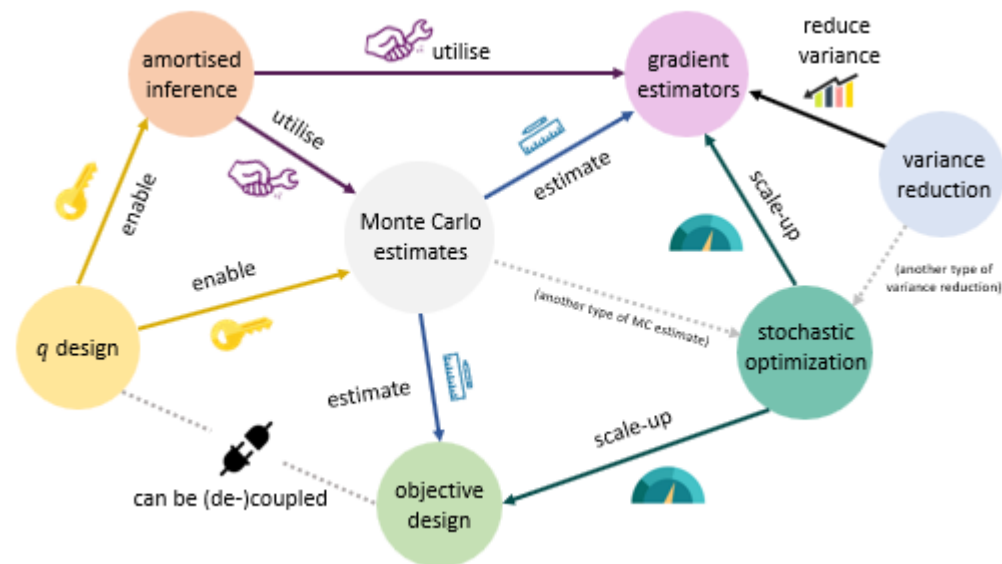
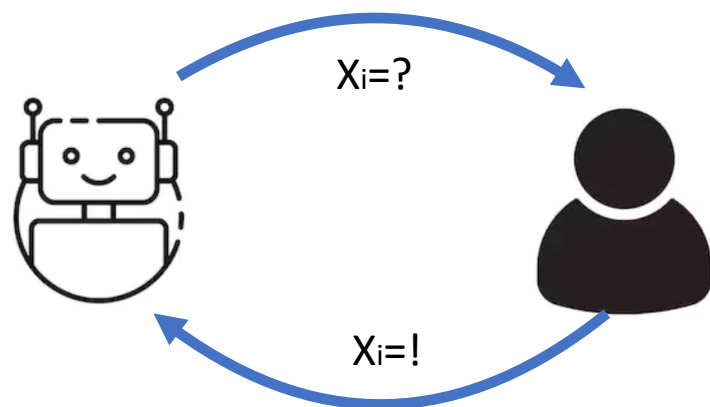
Summary

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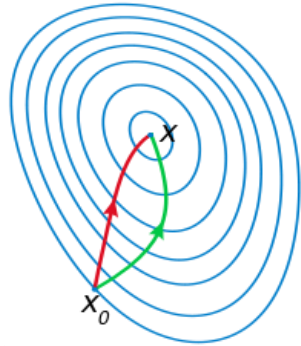
Approximate Inference

$$\int F(\theta) \pi(\theta) d\theta$$



Future Directions: Methodology

Better optimization



Zhang et al. Noisy natural gradient as variational inference. ICML 2018

Khan et al. Fast and Scalable Bayesian Deep Learning by Weight-Perturbation in Adam. ICML 2018

Future Directions: Methodology

Better optimization

q distribution design
objective design
amortised inference
scalable inference
...

Combined approaches

rejection sampling
importance sampling
SMC, MCMC, Quasi MC
...



q distribution design
objective design
amortised inference
scalable inference
...

Future Directions: Methodology

Better optimization

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scalable inference
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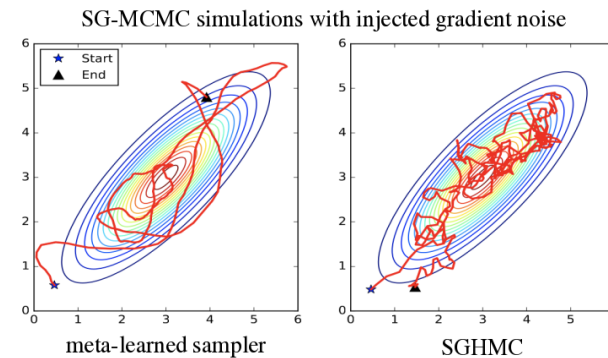
Combined approaches

rejection sampling
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q distribution design
objective design
amortised inference
scalable inference
...

Meta-learning inference algorithms



Gong et al. Meta-learning for Stochastic Gradient MCMC. ICLR 2019

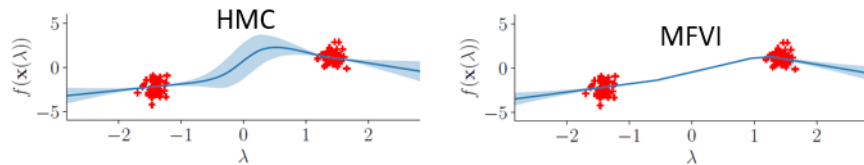
Zhang et al. Meta-Learning for Variational Inference. AABI 2019

Future Directions: Error Analyses

Errors in inference

$$D[q(\theta)||p(\theta|D)] = ?$$

$$D[\underline{q(y^*|x^*)}||p(y^*|x^*, D)] = ?$$
$$= \int p(y^*|x^*, \theta) q(\theta) d\theta$$



Analysis needed for **deep** probabilistic models!

- Optimization error
- Approximation gap

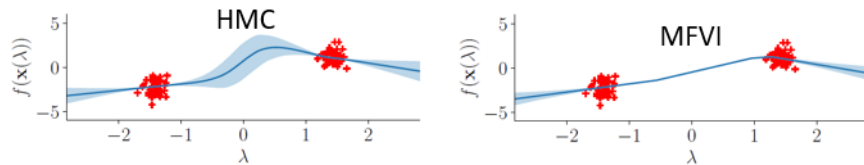
Future Directions: Error Analyses

Errors in inference

$$D[q(\theta) || p(\theta|D)] = ?$$

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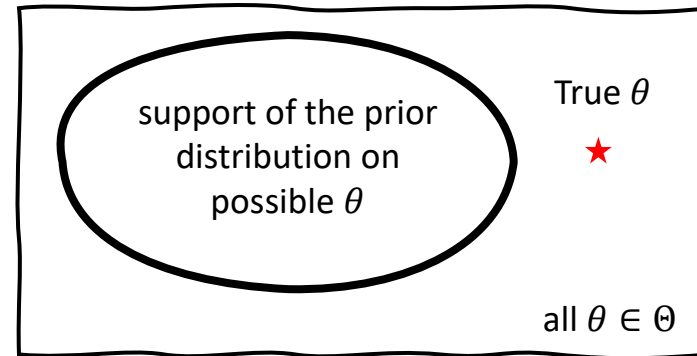
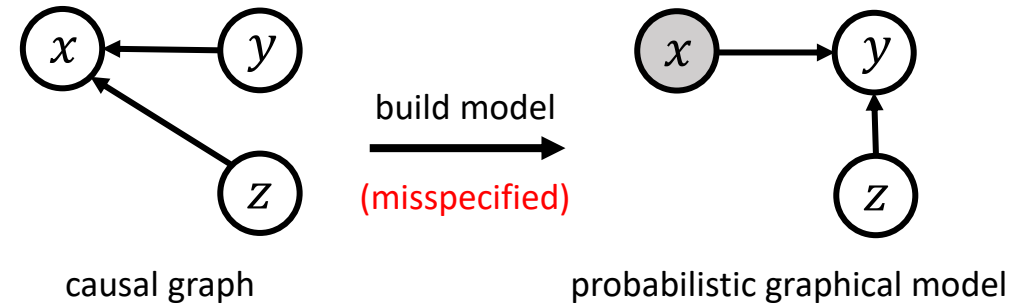
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Analysis needed for **deep** probabilistic models!

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- Approximation gap

Model misspecification

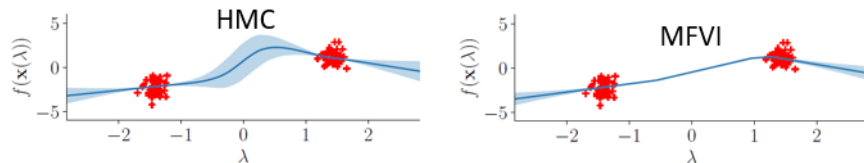


Future Directions: Error Analyses

Errors in inference

$$D[q(\theta)||p(\theta|D)] = ?$$

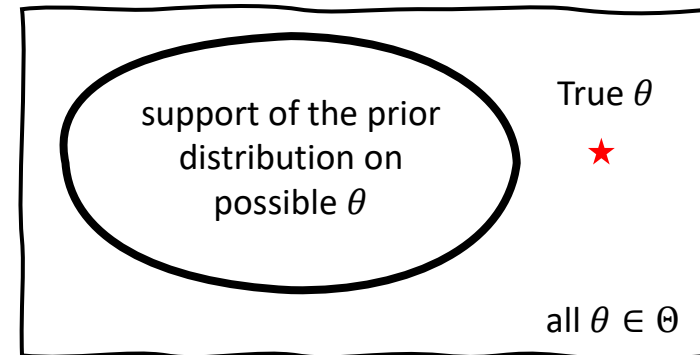
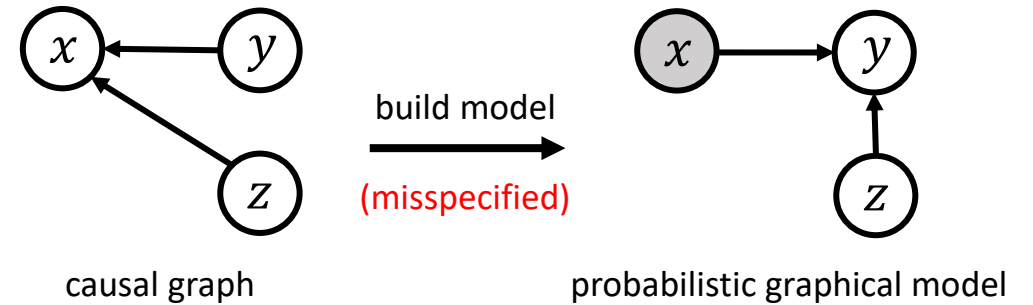
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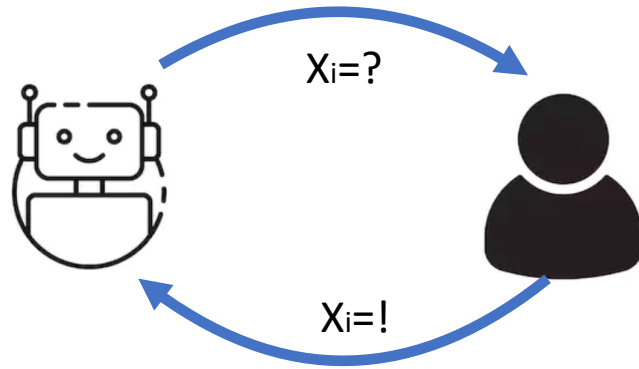
Model misspecification



Separation of inference & modelling?

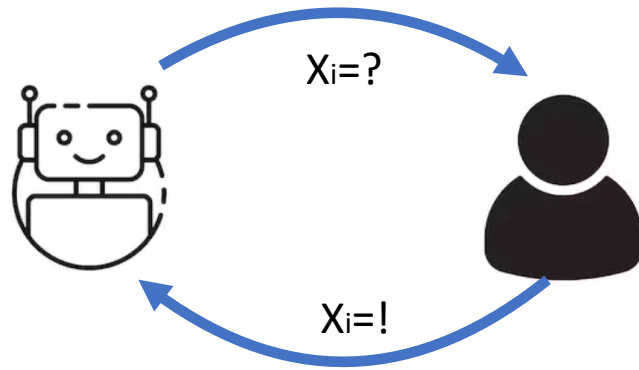
Future Directions: Applications

Uncertainty estimation

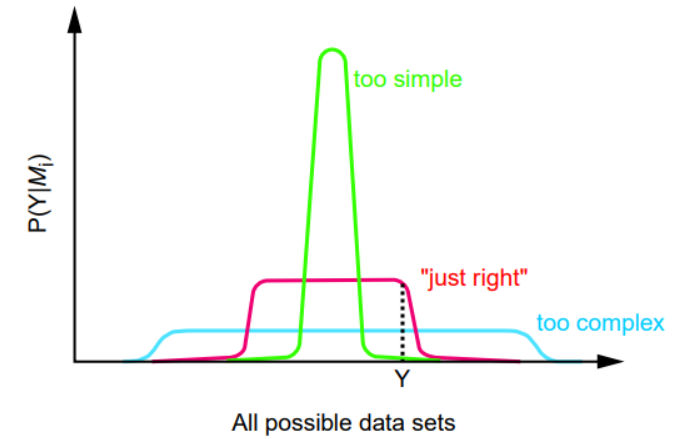


Future Directions: Applications

Uncertainty estimation

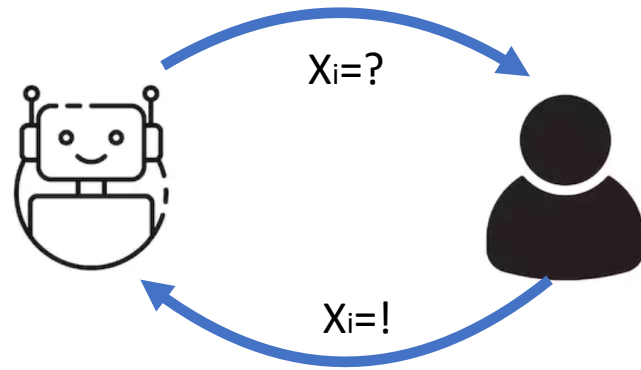


Model selection & averaging

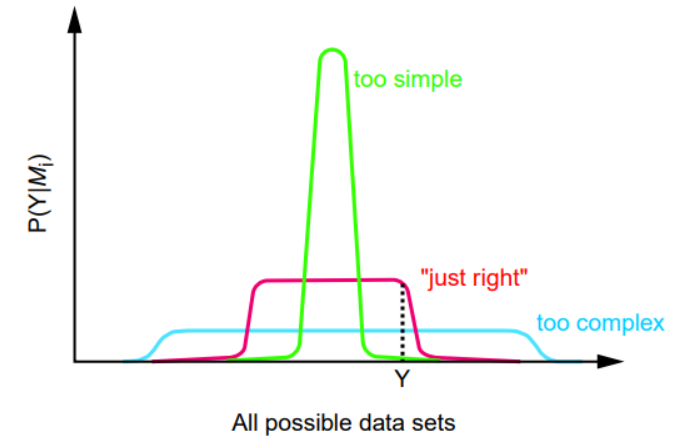


Future Directions: Applications

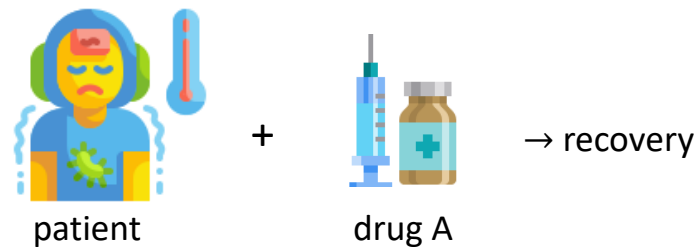
Uncertainty estimation



Model selection & averaging



Causal reasoning



"what if the patient was treated with drug B?"



Thank You!

Questions? Ask at:

liyzhen2@gmail.com (Yingzhen Li)

Cheng.Zhang@microsoft.com (Cheng Zhang)

